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Accepted Version

Dunn, G., Fowler, D., Rollinson, R., Freeman, D., Kuipers, E., Smith, B., Steel, C., Onwumere, J., Jolley, S., Garety, P. and Bebbington, P. (2012) Effective elements of cognitive behaviour therapy for psychosis: results of a novel type of subgroup analysis based on principal stratification. *Psychological Medicine*, 42 (5). pp. 1057-1068. ISSN 1469-8978 doi: 10.1017/S0033291711001954 Available at <https://centaur.reading.ac.uk/24692/>

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To link to this article DOI: <http://dx.doi.org/10.1017/S0033291711001954>

Publisher: Cambridge University Press

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# **Effective elements of CBT for psychosis: Results of a novel type of subgroup analysis based on principal stratification**

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**Word count: Abstract 237      Text 4652**

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## **ABSTRACT**

**Background:** Meta-analyses show that cognitive behaviour therapy for psychosis (CBT-P) improves distressing positive symptoms. However, it is a complex intervention involving a range of techniques. No previous study has assessed the delivery of the different elements of treatment and their effect on outcome.

**Aims:** To assess the differential effect of type of treatment delivered on the effectiveness of CBT-P, using novel statistical methodology.

**Method:** The Psychological prevention of Relapse in Psychosis (PRP) trial was a multi-centre randomised controlled trial (RCT) which compared CBT-P with treatment as usual. Therapy was manualised, and detailed evaluations of therapy delivery and client engagement were made. Follow-up assessments were made at 12 and 24 months. In a planned analysis, we applied principal stratification (involving structural equation modelling with finite mixtures) to estimate intention-to-treat effects for sub-groups of participants, defined by qualitative and quantitative differences in receipt of therapy, while maintaining the constraints of randomization.

**Results:** Consistent delivery of *full therapy*, including specific cognitive and behavioural techniques, was associated with clinically and statistically significant increases in months in remission, and decreases in psychotic and affective symptoms. Delivery of *partial therapy* involving engagement and assessment was not effective.

**Conclusions:** Our analyses suggest that CBT-P is of significant benefit on multiple outcomes to patients able to engage in the full range of therapy procedures. The novel statistical methods illustrated in this report have general application to the evaluation of heterogeneity in the effects of treatment.

**Trial Registration:** isrctn.org Identifier ISRCTN 83557988

**Declaration of interest:** None.

## INTRODUCTION

The pharmacological treatment of positive symptoms of psychosis is only moderately successful (Leucht *et al.* 2009). Cognitive behaviour therapy for psychosis (CBT-P) consistently reduces psychotic symptoms in people with distressing medication-resistant symptoms, although average effect sizes are quite small (Pilling *et al.* 2002; Jones *et al.* 2004; Zimmerman *et al.* 2005; Wykes *et al.*, 2008). Recent guidelines recommend it (NICE, 2009).

The Prevention of Relapse in Psychosis trial was designed to evaluate the effectiveness of CBT-P in reducing relapse and improving symptoms. This trial compared CBT-P and Family Intervention (FI) with treatment as usual (TAU), and is fully described elsewhere (Garety *et al.* 2008). Intention-to-treat (ITT) analysis found no benefits for the primary outcomes of relapse and days in hospital, or for reduction in psychotic symptoms, but did find a significant improvement in depression (Garety *et al.* 2008). The trial manual describes cognitive and behavioural techniques targeting the various symptoms and problems presented by individuals. Therapists shape the techniques to the particular problems that emerge during therapy (Fowler *et al.* 1995). Therapy therefore varies widely, being tailored to the individual needs of a heterogeneous group of clients with differing levels of capacity and willingness to engage.

However, CBT-P has so far only been evaluated as an overall package. We lack information about the delivery of different therapeutic techniques, given that clients may not be ready or able to countenance particular interactions (Durham *et al.* 2003). Single case studies have suggested that it is the more active techniques that lead to

symptomatic changes, rather than the necessary but preliminary stage of relationship-building and assessment (Fowler & Morley, 1989; Chadwick *et al.* 2003).

The present study investigates how far competent CBT therapists were able to deliver different types of therapy technique, and the impact this has on efficacy. Reliable methods for monitoring therapeutic delivery in CBT-P allow us to examine its relationship with outcome (Durham *et al.* 2003; Startup *et al.* 2002; Rollinson *et al.* 2007; Rollinson *et al.* 2008). Based on our cognitive model (Garety *et al.* 2001), we hypothesised that outcomes would be improved when therapists were able to deliver the more specific cognitive and behavioural techniques.

A key feature of this study is our novel statistical approach for analysing differential efficacy in randomised trials, informed by increasing recognition of the biases and confounding inherent in past attempts at *post hoc* estimation of outcomes in relation to aspects of therapeutic quality (Dunn & Bentall, 2007). The new approach estimates *ITT effects for subgroups* by comparing the effects of intervention with putative effects in the control arms ignored in traditional analyses. The evaluation of differential efficacy of CBT-P techniques formed part of the original protocol.

## **METHODS**

### **Settings**

The trial took place in five mental health services: two in inner city London, one in suburban outer London, one in a provincial city (Norwich), and one in a rural area (Norfolk).

### **Study design**

The PRP trial comprised two pathways with separate randomisation, stratified within the five participating centres, and within in-patient or out-patient status at induction. The first ('individual pathway') included participants without carers randomly allocated to two groups: both received good standard care (treatment as usual, TAU), while the experimental group also received CBT-P. In the second pathway ('carer pathway'), those with carers were allocated to three groups (CBT-P plus TAU, FI plus TAU; or TAU alone). The current analysis is restricted to hypotheses concerning CBT-P only, so the FI participants were excluded.

### **Participants**

We approached consecutive patients with recent relapses, whether or not they had been admitted. They invited to take part once they could give informed consent. The inclusion criteria were: current clinical diagnosis of non-affective psychosis (F2: ICD-10; WHO, 1992; DSM-IV; APA, 1994); age 18-65 years; a second or subsequent psychotic episode starting not more than three months before induction; and a rating of at least 4 (moderate severity) for at least one positive symptom on the Positive and Negative Syndrome Scale (Kay, 1991). Exclusion criteria comprised: a primary diagnosis of alcohol or substance dependency, organic syndrome or learning disability; spoken English inadequate for engaging in psychological therapy; and unstable residential arrangements.

Participants provided informed consent under protocols approved by the appropriate ethics committees. Full details of trial protocol are provided elsewhere (Garety *et al.* 2008). Participants were assessed at baseline before randomisation, and at 3, 6, 12 and 24 months. The CBT-P was completed within 12 months, while TAU continued throughout.

### **Treatment**



CBT-P was delivered for nine months, with a planned minimum of 12, and maximum of 20 sessions. The therapy in our generic CBT-P manual (Fowler *et al.* 1995) was augmented with specific relapse prevention techniques

### **Therapy provision**

One hundred and thirty three people were allocated to CBT-P. They received a mean of 14.3 sessions (s.d. 7.8), each lasting on average one hour. The number of sessions was very similar in the individual and carer arms.

### **Trial therapists: training and monitoring of adherence and competence**

Five lead trial therapists ('lead therapists', all doctorate level or equivalent clinical psychologists employed full time on the trial, provided therapy to 96 CBT-P participants. A further 37 CBT-P participants were seen by therapists employed by the NHS Trusts running the local mental health services ('trust therapists'). The trust therapists were doctoral clinical psychologists and nurses with specialist training in CBT-P. All were fully trained, and closely managed and supervised: details of recruitment, training and quality control are provided elsewhere (Garety *et al.* 2008). The Revised Cognitive Therapy for Psychosis Adherence Scale (R-CTPAS) (Rollinson *et al.* 2008) is a measure of *fidelity*, designed to provide precise definitions of the minimum therapeutic delivery of CBT-P activity. It covers 21 different types of CBT-P techniques. Therapist *competence* was measured by the Cognitive Therapy Scales (CTS; Young & Beck, 1980). All raters were trained to criterion on this scale, and met regularly to check rating reliability. 185 tapes from 66 therapy participants (62% of the total treated) were sent for formal monitoring by the lead therapists from other centres. In 90% of the sample, the CBT-P delivered in taped interviews was both adherent and competent. In 8 cases (8.3%), the therapy was regarded as supportive work rather than CBT-P. A randomly selected sub-sample of 36 tapes was

sent to external expert raters: their ratings showed excellent agreement with the internal raters (Garety *et al.* 2008).

All therapists also used the R-CTPAS to provide self-report assessments of their therapy sessions. Agreement between tape-rated and self-reported ratings of R-CTPAS across multiple raters was satisfactory, with intraclass correlation coefficients for composite scores ranging from 0.5 to 0.8.

### **Summary scores of therapy delivery**

In the present paper, the taped and self reported adherence ratings from the R-CTPAS (Rollinson *et al.* 2008) were used to create a single summary score for the therapy actually received by each person treated in the trial. These were derived from the factor analysis of the R-CTPAS, described above. The first factor, which we here term ‘partial therapy’, comprised *engagement and assessment techniques*: i.e. active attempts to engage in therapeutic strategies; the ‘Columbo style’ (which assesses the degree to which therapists promote guided discovery); and the collaborative assessment of psychotic experience and delusional beliefs. The second (‘full therapy’) comprised *active therapy techniques*: i.e. relapse prevention interventions; enhancing self regulatory strategies; developing a personal model of relapse; developing a model of psychosis; work on reinterpreting the meaning of delusional beliefs and hallucinations; and schema work. Both factors described components of active therapy, and the first factor should not be confused with befriending (Sensky *et al.*, 2000).

Our intention was to ensure that this overall assessment of therapy techniques was reliable. We used data from 1019 sessions from 102 participants where there were

sufficient R-CTPAS data to be fully representative of level of therapy. 31 participants had significant levels of missing individual session data, and were coded in the present analysis as 'not known'. There were no significant differences on the baseline variables of interest between the sub-sample analysed and those with missing data. There was a good spread of sessions across the whole duration of therapy: 30% of the sessions evaluated came from block 1 (sessions 1-4), 26% from block 2 (sessions 5-9), 22% from block 3 (sessions 10-14), and 21% from block 4 (sessions 15 and above). Average item scores from all available sessions were calculated for each participant. The R-CTPAS manual uses a score of 1 (within a range from -7 to +7) to indicate the minimum threshold for highly competent delivery of individual techniques. In order for an individual technique to be considered present across the course of therapy, the averaged item score needed to be one or above where self-report data were available, or to be judged above the competence threshold in at least three sessions, supported by a tape. This aimed to reflect the definite presence of a therapy technique occurring across the course of therapy, and was deliberately chosen to signify the unequivocal delivery of high quality interventions.

For *full* or *partial* therapy to be considered present, at least one of the composite active intervention or engagement and assessment items listed above needed to be present above this threshold across the course of therapy. Clients who received fewer than five therapy sessions formed a third, *no therapy*, group, as this number of sessions was regarded as too small for the delivery of effective CBT-P.

### **Control Condition**

TAU consisted of good standard care, delivered according to national and local service protocols and guidelines, including the provision of anti-psychotic medication. The frequency and nature of service contacts was monitored, as were medication

regimes. TAU did not preclude the provision of psychological interventions by locality teams, although this was unusual.

### **Primary outcome measures**

The primary outcome variable, relapse, was assessed by a blind panel evaluation procedure (Craig *et al.* 2004; Bebbington *et al.* 2006). Consensus remission and relapse ratings were applied to detailed extracts of the clinical case notes by paired members of the research team, using manualised *a priori* operational definitions (Bebbington *et al.* 2006). The original trial report gives full details (Garety *et al.* 2008). Here we present the data as the total number of months in full remission separately over the first and second years of the trial. Data on hospital admissions were collected through the hospital administration systems.

### **Secondary outcome measures**

Secondary outcomes were rated by research assessors at interview, and again considerable efforts were made to achieve blind ratings (Garety *et al.* 2008). They comprised:

*Positive and Negative Syndrome Scale* (PANSS) (APA, 1994). This is a 30-item, seven point (1-7) rating instrument assessing psychotic symptoms over the last week.

We present results for the PANSS Total (30 items) and PANSS Positive scores (7 items).

*Beck Depression Inventory – II* (Beck *et al.* 1996). The BDI-II is a self-report 21-item, four point scale (0-3) for the assessment of depression in relation to the previous fortnight.

### **Statistical Analysis**

All analyses reported in the main trial paper (Garety *et al.* 2008) were based on the ITT principle, allowing for potential biases arising from loss to follow-up (under the

assumption that missing outcomes were missing Missing at Random (MAR) using the terminology of Little & Rubin (2002).

All analyses presented in the present paper involve estimating ITT effects within three classes of participant. These three classes (*principal strata* (Angrist *et al.* 1996; Frangakis & Rubin, 2002)) are defined by the potential outcome of participants' treatment allocation. Stratum 1 (no therapy) comprises participants who would receive little or no therapy (CBT-P) regardless of their randomised allocation. Stratum 2 (partial therapy) comprises those who received partial therapy in the CBT-P group, together with those controls *who would have received partial therapy if they had been allocated to the CBT-P condition*. Finally, Stratum 3 (full therapy) comprises those participants who received full therapy in the CBT-P group, together with those controls *who would have done so, had they been allocated to CBT-P*. These three strata are only partly identified – class membership is known for most participants allocated to CBT-P, but not for the controls, and the model is correspondingly said not to be identified (i.e. unique stratum-specific treatment (ITT) effects cannot be derived). However, it is possible to identify empirically baseline covariates that predict the type of treatment delivered in the randomised-in participants. Because of randomisation, these can also be used to predict potential treatment compliance in the randomised group. In the present study, the best predictors were treatment centre (location), presence of a carer, inpatient status and sex of the patient. There was no association between baseline symptomatology and the type of therapy received.

We used the predictors as covariates in a latent class model to predict principal stratum membership. The same covariates were used in the simultaneously fitted

analysis of covariance model used to estimate the stratum-specific treatment (ITT) effects. Model identification was further improved by assuming that the ITT effect in the ‘no therapy’ stratum was zero (i.e. allocation to CBT-P has no average effect when the participant fails to take up the offered therapy). This is a so-called *exclusion restriction*.

A further refinement is based on the realization that the probability of participants having missing outcome data (i.e. loss to follow-up) is likely to be dependent on stratum membership: the no-therapy group, for example, would seem less likely to provide outcome data than those from the other two strata. The missing data mechanism might still be MAR, but it might equally be Latently-Ignorable (LI) (Frangakis & Rubin, 1999). In the LI model, the probability of loss to follow-up is jointly dependent on stratum membership and the outcome of random allocation, and also on baseline covariates (the structure of the missing data model then being analogous to that for the outcomes). The exclusion restrictions for the missing data indicator were exactly the same as for the final outcomes. Technical details, and illustrations using data from psychological treatment trials, are provided elsewhere (Dunn *et al.* 2005; Emsley *et al.* 2010).

All analyses reported in the present paper have been carried out using *Mplus* version 5.2 (Muthen & Muthen, 1998-2009). In order to avoid local maxima (invalid estimates), 1000 randomly perturbed sets of starting values were used. All standard errors were estimated using bootstrapping (250 replications) (Efron & Tibshirani, 1993). Note that in none of the analyses have we allowed for individual therapist

effects (clustering by therapist) – for technical reasons it would have not been feasible –assuming rather that these are subsumed by the centre effects.

## **Results**

In all participants receiving full therapy, the techniques associated with partial therapy were also delivered. In every case, at each level of therapy, multiple techniques were present. 42 participants had full therapy, and 39 partial therapy. A further 21 participants had fewer than five sessions of therapy, thus falling into our no therapy group. It should be emphasised that partial therapy met the definition of highly competent cognitive therapy, and was observed to be accompanied by attempts by the therapists to deliver the techniques of full therapy as well. However, in partial therapy such attempts, by definition, fell below the pre-defined threshold for the identification of full therapy.

Table 1 provides information on the demographic characteristics of the trial participants. Table 2 illustrates the distribution of those in the CBT-P group receiving no therapy, partial therapy and full therapy, cross-classified by various baseline factors. Treatment centre (location) appears to be the best predictor of therapy received (note, in particular that 19 of the 42 patients receiving full therapy were from Centre 3).

### **Tables 1 & 2 near here**

Table 3 provides information on the two main outcomes (time in remission and PANSS total scores) by treatment arm, separately for the no therapy, partial therapy and full therapy subgroups. There are no obvious patterns, and the full therapy subgroup did no better than the others. However, using the mean outcomes for these

sub-groups in this way *cannot distinguish between effects arising from the treatment of interest, and those deriving from treatment-independent prognosis* (confounding or selection effects). Hence the need for more refined analysis. The requirement that is missing for the comparison of subgroup treatment effects is the average outcome in the respective principal strata in the control (TAU) condition.

**Table 3 near here.**

We now summarise the analyses based on the use of principal stratification. In Table 4, we provide estimates of stratum-specific ITT effects for our four chosen outcomes, displayed separately for follow-up at 12 and 24 months. There were virtually no missing data for the number of months in remission, and we assumed that such missing data as existed were missing at random. Data from research interviews were more likely to be missing, and we therefore used two separate methods for dealing with missing data. In the first, we assumed data were missing at random (MAR). The second analysis assumes missing outcomes were latently ignorable (LI).

The table shows differences between treatment and control groups for each of the two principal strata corresponding to partial therapy and full therapy, respectively. Full treatment brings about nearly six additional months in remission between induction and 12-month follow-up (indicated by an ITT effect with a positive sign), and an additional two months between the 12 and 24-month assessments. The 12-month effect is statistically significant (at the 5% level). There is a suggestion that the ITT effect in the partial therapy group may be negative (detrimental), but the effects are not statistically significant.



The results for months in remission are mirrored in the findings for PANSS and BDI scores. At 12 months, the full therapy group had a statistically significant 16-point advantage on PANSS total score over the control group (an ITT effect with a negative sign). Under the assumption that missing data were latently ignorable, the PANSS advantage fell to 12 points, and was no longer significant. At 24 months, the advantage was still 11 points (12 under latently ignorable assumptions), albeit no longer statistically significant. Again, there is a suggestion that partial therapy might be detrimental rather than beneficial. Stratum-specific ITT effects for the PANSS Positive and BDI scores were consistent with the above findings, though none of the effects was statistically significant.

**Table 4 near here.**

The results in Table 4 indicate that stratum-specific ITT effects for the 24 month outcomes were very similar to those at 12 months. We therefore decided to refine our analyses by estimating stratum-specific ITT effects that were assumed to be common (i.e. the same) for the first and second 12-month periods of follow up. The refined ANCOVA model for the outcomes was now bivariate. This allows for period-specific effects of the baseline covariates and correlations between the residuals of the outcomes at the two periods, and is an example of a Seemingly Unrelated Regression (SUR) (Cox & Wermuth, 1996). The rationale was to improve both precision and statistical power, justified in the light of the consistency of effects across measures and periods.

Table 5 records the estimates of the stratum-specific ITT effects common to the two periods covered in follow-up. Initially, each result is presented three times (three rows of ITT estimates). The first carries no exclusion restrictions (a relaxation of the

assumptions in the models fitted above). In the second, the no-therapy group is set to zero (a single pair of exclusion restrictions – one for the 12 months outcome and another for 24 months – corresponding to our initial models). The last analysis includes similar constraints imposed on both the no-therapy and partial therapy groups (two pairs of exclusion restrictions). The introduction of these additional exclusion restrictions prevents the effect of partial therapy from being detrimental, and is therefore a stringent test of the effect in the full therapy stratum. This change (and the relaxation of all restrictions as in the top row) provides a check on the sensitivity of the estimate of the effect of full therapy to a different set of model assumptions. As in Table 4, the results are calculated under different assumptions about the distribution of missing data.

#### **Table 5 near here**

We made two further sensitivity checks. The first involved merging the first two principal strata into one (the ITT estimates in rows four and five). The second involved allocating those with a missing treatment indicator in the CBT-P arm either to the worst option (no therapy) or to the best (full therapy). Overall, the refined analyses based on bivariate outcomes confirmed the findings in Table 4. Whatever the measure used, there was a considerable and significant advantage in the full therapy group, and the suspicion of detriment in those receiving only partial therapy. Finally we return to Table 2(a). There was a centre effect in the delivery of treatment: the delivery of full therapy was more frequently achieved in the rural county of Norfolk (Centre 3). If our conclusions concerning the effects of receiving full therapy are valid, then the direct implication is that the ITT effect of CBT in Centre 3 (rural

Norfolk) would be very different to that in the other centres. We therefore carried out a more conventional analysis of centre effects (i.e. testing the Centre 3 by CBT interaction). This is available from the first author. It demonstrated that the treatment (Intention-To-Treat) effects in Centre 3 were commensurably better in terms of PANSS total and BDI scores, but not the PANSS Positive score or months in remission.

## **Discussion**

This study used a novel approach to estimate the treatment effects of subgroups of the arm receiving CBT-P in a large randomised controlled treatment trial. The approach provides less biased estimates of the effect size of such subgroups by taking account of the potential outcomes had such cases been randomised to the control group. The study compared three categories of treatment. Our hypothesis was that the subgroup which engaged with and received full CBT-P would have better outcomes than those who received partial therapy or who dropped out. Our results are wholly consistent with this hypothesis. Treatment was effective if, and only if, clients received full therapy. Gains were large, and both clinically and statistically significant. They were also consistent, applying both to the number of months recovered and relapse free (the primary outcome), and to psychotic and depressive symptom outcomes. Participants who received therapy consisting only of engagement and assessment work did not benefit, and neither did those who dropped out. There is a suggestion that therapy had a somewhat deleterious effect on the former group.

This is a novel analysis, based on a development of the methods of Complier-Average Causal Effect (CACE) estimation (Angrist *et al.* 1996; Frangakis & Rubin, 2002).

The latter has already been applied to randomised controlled trials in psychiatry (Dunn *et al.* 2003; Horvitz-Lennon *et al.* 2005; Bellamy *et al.* 2007; Serfaty *et al.* 2009). The analysis is dependent on modelling that aims *a priori* to circumvent the biased estimates of treatment effects obtained by traditional per protocol approaches to analysis. In applying this technique, we have arrived at an estimate strikingly different from the overall ITT result we reported previously, which showed no effect of CBT-P beyond reducing depression at 24 months. Moreover, the modelled effect is not apparent from simple observation of the mean effects of subgroups within the treated arm alone (Table 3). Without a proper understanding of the assumptions underpinning estimations of treatment effects in randomised trials, this may seem counterintuitive. We must however take account of the fact that simple descriptions of mean effects in subgroups of a single arm of a trial (the treatment arm) are actually highly biased estimates of treatment effects. Such descriptive statistics do not take account of biases due to dropout, and to the putative effects if those randomised to treatment had instead been randomised to control. The modelling used here has been developed specifically to overcome such biases, and is described elsewhere in specialist publications (Frangakis & Rubin, 2002; Frangakis & Rubin, 1999; Dunn *et al.* 2005; Emsley *et al.* 2010). The approach has application to any situation where heterogeneity in treatment response is analysed in terms of subgroups defined by post-randomisation explanatory variables. Examples include medication adherence, therapeutic alliance, and intermediate biomarkers such as immune response.

The trial was designed *a priori* to study the effects of differing levels of therapy delivery. We used detailed observations of adherence and competence to identify those who received full CBT-P. Only 40% of participants did so. This raises the

question of why so few actually received full therapy. CBT-P is complex, and its effective delivery depends on the interaction between therapist and patient, and hence on two types of factor: those relating to the patient (readiness and willingness to engage, the nature of symptoms, awareness, levels of distress) and those relating to the therapist (ability, training, supervision, adherence and competence). We ensured that the therapists in the trial were trained to the highest standards, and this was supported by our detailed monitoring of therapy sessions. Despite this, they were able to deliver full therapy only to a minority. This might therefore be the result of patient attributes in this sample, although we must emphasise that there were no baseline differences in symptoms. While people with psychosis have well-known problems with engagement in therapeutic relationships, the techniques of CBT-P have been specifically designed to minimise them. Nevertheless, in a substantial minority in the present study, therapists were not able to move much beyond maintaining engagement and working collaboratively with clients to make sense of their problems. It would have been interesting to relate the characteristics of the CBT-p received with the strength of the therapeutic alliance and to look at their joint relationship with the effects of therapy. However, the statistical methods required to undertake this work are in their infancy (Dunn & Bentall, 2007; Emsley *et al.*, 2010).

The superior delivery of therapy and better treatment effects in Norfolk are interesting. The difficulties of delivering complex interventions in inner city areas are well-known to clinicians, and might be attributed variously to low levels of social support, high levels of deprivation, and relative residential instability. Such contextual disadvantages remain a therapeutic challenge.

Consistent attempts were made to deliver more active cognitive and behavioural techniques to all clients, but with many it was impossible to achieve the level necessary for the *a priori* definition of full therapy. There are many possible reasons for this, some of which may be especially characteristic of un-selected, recently relapsed groups. Despite initial willingness, after a few weeks some patients no longer wanted to receive therapy. Some had symptoms, but were not distressed by them, some had responded to the reinstatement of medication (this was not a medication resistant sample, indeed many relapses seemed to follow discontinuation of medication) and no longer saw the point of a psychological treatment. Some simply lacked interest in working with a therapist, and others had limited awareness of their problems. Despite such difficulties, our therapists still managed to keep these clients engaged in therapy. We had clear observational evidence of therapists establishing a basic working cognitive behavioural relationship, systematically carrying out assessment, and promoting collaborative guided discovery in a highly skilled manner. It must be emphasised however that, in this trial, persistence was sometimes associated with a worsening of symptoms. This is an important observation, with implications for clinical practice. We conclude that if therapists have not managed to move into the active phase of therapy within a circumscribed period, it may not be worth persisting, although clinical experience backs the option of a later return to therapy. Only clients with whom therapists can deliver a substantial amount of active therapy appear to benefit: future work should aim to identify them.

In summary, this analysis shows clearly that CBT-P has widespread and beneficial effects when delivered as intended in a group of relapse-prone patients. These effects apply to our original primary and secondary outcomes, of relapse prevention and

symptomatic improvement. CBT-P is clearly therefore a useful and effective intervention. However, our results also indicate that those clients whom therapists cannot engage in substantial active therapy may not benefit: at best it is not cost-effective to continue therapy under such circumstances.

**Funding/Support:** The study was supported by a Wellcome Trust Programme Grant (062452). The developments in statistical methodology (GD) were supported by MRC Methodology Research Programme Grants G0600555 and G0900678.

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**Table 1****Demographic characteristics of participants**

	<b>TAU (n=140)</b>	<b>CBT(n=133)</b>
<b>Gender</b>		
Male	97	95
Female	43	38
<b>Ethnicity</b>		
White	105	88
Black Caribbean	8	14
Black African	14	13
Black other	4	3
Indian	0	5
Other	9	10
<b>Employment</b>		
Employed	8	5
Employed part-time	4	2
Voluntary employment	1	7
Unemployed	117	112
Student	8	3
Retired	2	1
Housewife/husband	0	2
Unknown	0	1
<b>Marital status</b>		
Single	107	91
Married	8	16
Divorced/separated	21	24
Widowed	1	2
Unknown	3	0
<b>Diagnosis</b>		
Schizophrenia	117	117
Schizo-affective disorder	23	12
Delusional disorder	0	4
<b>In hospital?</b>		
Inpatient	96	92
Outpatient	44	41
<b>History of violence</b>		
No	102	86
Yes	35	42
<b>History of suicide or self-harm</b>		
No	80	81
Yes	55	46
Mean Age (s.d.)	36.8 (10.9)	39.0 (10.7)
Mean length of illness in years(s.d.)	10.1 (8.6, n=135)	10.9 (8.4, n=131)
Mean no. of admissions (s.d.)	4.4 (4.6, n=130)	4.7 (5.2, n=129)
PANSS total mean (s.d.)	65.8 (15.8)	63.2 (13.7)
PANSS Positive mean (s.d.)	18.5 (5.3)	17.8 (5.5)
BDI II mean (s.d.)	20.7 (13.5, n=136)	22.2 (12.2, n=130)

**Table 2****Number of participants receiving each level of therapy****(a) By Location**

	Centre 1	Centre 2	Centre 3	Centre 4	Centre 5	Total
Control group	26	19	35	35	25	140
CBT group						
No therapy	2	4	6	2	7	21
Partial therapy	8	6	11	2	39	
Full therapy	9	1	19	8	5	42
Not known	3	2	3	13	10	31

**(b) By Carer**

	Has carer	Does not have carer
Control group	41	99
CBT group		
No therapy	10	11
Partial therapy	11	28
Full therapy	13	29
Not known	12	19

**Table 2 (contd.)**

**(c) By Inpatient status**

	Inpatient	Outpatient
Control group	96	44
CBT group		
No therapy	12	9
Partial therapy	23	16
Full therapy	32	10
Not known	25	6

**(d) By Sex**

	Male	Female
Control group	96	44
CBT group		
No therapy	16	5
Partial therapy	25	14
Full therapy	32	10
Not known	23	8

**Table 3****Outcomes by level of therapy (mean, sd, N)**

	Months remission (0-12 months)	Months remission (12-24 months)	PANSS total (baseline)	PANSS total (12 months)	PANSS total (24 months)
Control group	7.60, 4.47, 137	8.77, 4.56, 132	65.83, 15.76, 140	59.39, 16.61, 113	57.04, 15.96, 109
CBT group					
No therapy	5.10, 5.25, 20	8.45, 4.57, 20	63.00, 16.16, 21	56.40, 14.68, 10	52.27, 15.37, 11
Partial therapy	8.39, 3.78, 38	9.68, 3.91, 38	66.15, 12.68, 39	60.13, 14.41, 38	58.94, 14.57, 36
Full therapy	7.57, 4.52, 42	8.60, 4.43, 42	63.50, 13.09, 42	55.97, 13.98, 37	56.27, 11.59, 37
Not known	7.58, 4.39, 31	8.57, 3.77, 30	59.42, 13.79, 31	53.96, 16.26, 26	53.96, 14.79, 25



**Table 4**  
**ITT estimates within principal strata, separately for 12- and 24-months outcomes**  
**(bootstrapped standard errors in brackets)**

**(a) Negligible amounts of missing data (assumed MAR)**

	No therapy (Non-compliers)	Partial therapy	Full therapy
<b>Months in Remission</b>			
0-12 months	0*	-1.7 (1.1)	+5.6 (2.7)**
12-24 months	0*	-2.1 (1.4)	+2.1 (2.1)

**(b) With missing outcomes assumed to be MAR**

	No therapy (Non-compliers)	Partial therapy	Full therapy
<b>PANSS total score</b>			
12 months	0*	+6.4 (3.7)	-16.4 (6.8)**
24 months	0*	+7.5 (4.6)	-11.3 (5.9)
<b>PANSS positive score</b>			
12 months	0*	+1.8 (1.6)	-5.5 (2.8)
24 months	0*	+3.4 (2.4)	-2.8 (3.3)
<b>BDI score</b>			
12 months	0*	+2.2 (4.5)	-2.8 (4.5)
24 months	0*	+2.7 (3.5)	-7.4 (4.6)

**(c) With missing outcomes assumed to be LI**

	No therapy (Non-compliers)	Partial therapy	Full therapy
<b>PANSS total score</b>			
12 months	0*	+7.7 (5.0)	-11.8 (8.2)
24 months	0*	+6.5 (5.1)	-12.3 (7.0)
<b>PANSS positive score</b>			
12 months	0*	+1.7 (2.2)	-3.0 (2.8)
24 months	0*	+3.7 (2.4)	-2.6 (3.2)
<b>BDI score</b>			
12 months	0*	+5.7 (5.1)	-4.9 (4.3)
24 months	0*	+2.5 (3.8)	-8.3 (6.3)

\* Model constraint (exclusion restriction)

\*\* Statistically significant (P<0.05): estimate two or more standard errors from zero

**Table 5****Estimated ITT effects within principal strata common to 12- and 24-months follow-up (bootstrapped standard errors in brackets)****(a) Negligible amounts of missing data (assumed MAR)**

	No therapy (Non-compliers)	Partial therapy	Full therapy
<b>Months in remission</b>	-2.8 (1.8)	-1.9 (0.5)**	+4.9 (1.5)**
	0*	-1.9 (0.5)**	+4.7 (1.7)**
	0*	0*	+4.4 (2.3)**
	No & Partial therapy (combined)		Full therapy
	0*		+3.5 (2.6)
	-2.3 (0.6)**		+4.4 (1.4)**
(Not known→No or partial therapy)	-2.3 (0.4)		+5.2 (1.2)
(Not known→Full therapy)	-2,5 (1.2)**		+3.7 (1.6)**

**Table 5 (continued)**

**(b) Missing data assumed to be MAR**

	No therapy (Non-compliers)	Partial therapy	Full therapy
<b>PANSS total score</b>	+9.0 (5.6)	+3.6 (3.1)	-13.0 (3.6)**
	0*	+5.2 (3.2)	-12.7 (4.1)**
	0*	0*	-12.2 (4.8)**
<b>PANSS positive score</b>	+3.2 (2.2)	+0.6 (1.4)	-3.9 (1.7)**
	0*	+1.2 (1.6)	-4.0 (1.9)**
	0*	0*	-3.6 (2.0)
<b>BDI score</b>	-0.5 (4.6)	+5.3 (3.8)	-8.5 (3.5)**
	0*	+6.3 (3.4)	-8.5 (3.1)**
	0*	0*	-7.0 (3.7)
		No & Partial therapy (combined)	Full therapy
<b>PANSS total score</b>		0*	-10.8 (5.4)**
		+6.0 (2.1)**	-11.7 (3.5)**
(Not Known→No or partial therapy)		+5.4 (1.9)**	-14.0 (3.2)**
(Not Known→Full therapy)		+8.2 (4.3)	-9.4 (4.5)**
<b>PANNS positive score</b>		0*	-2.0 (2.8)
		+2.0 (1.3)	-3.4 (2.2)
(Not Known→No or partial therapy)		+1.7 (0.9)	-4.2 (2.1)**
(Not Known→Full therapy)		-3.1 (2.5)	+2.5 (2.3)
<b>BDI score</b>		0*	-6.7 (3.3)**
		+3.6 (2.1)	-8.1 (3.1)**
(Not Known→No or partial therapy)		+2.5 (1.8)	-9.2 (3.2)**
(Not Known→Full therapy)		+4.0 (2.4)	-6.3 (2.7)**

\* Exclusion restriction

\*\* Statistically significant (P<0.05): estimate two or more standard errors from zero

## APPENDIX: Example of Mplus Input file (modelling 12 & 24 month PANSS scores)

```
TITLE: PRP Therapy - Principal Stratification
DATA:
  FILE IS "TherapyPredictors.dat";
VARIABLE:
  NAMES ARE r1 r2 r3 r4 carer sex output treat          ! r1 to r4 are binary dummies indicating treatment centre
  c1 c2 c3 pan0 pan12 resp12 pan24 resp24;           ! resp12 and resp24 are missing value indicators
  CLASSES c(3);
  TRAINING c1 c2 c3;                                ! binary dummies indicating compliance status
  CATEGORICAL resp12 resp24;
  USEVARIABLES treat sex output r1 r2 r3 r4 r5
  pan0 pan12 resp12 pan24 resp24
  c1 c2 c3 carer resp12 resp24;
  MISSING
  pan12 (999) pan24 (999)
ANALYSIS: TYPE=MIXTURE;
  ESTIMATOR=ML;
  STARTS = 1000 20;
  BOOTSTRAP=250;
MODEL:
  %OVERALL%
  PAN12 ON treat carer r1 r2 r3 r4 pan0 sex output;
  PAN24 ON treat carer r1 r2 r3 r4 pan0 sex output;
  pan12 WITH pan24;
  resp12 ON treat carer r1 r2 r3 r4 pan0 sex output;
  resp24 ON treat carer r1 r2 r3 r4 pan0 sex output;

  C#1 ON carer r1 r2 r3 r4 PAN0 sex output;
  C#2 ON carer r1 r2 r3 r4 PAN0 sex output;

  %C#1% ! No therapy
  [pan12 pan24 resp12$1 resp24$1];
  pan12 pan24 (3);
  pan12 ON treat@0 ;
```

pan24 ON treat@0 ;  
resp12 ON treat@0;  
resp24 ON treat@0;

%C#2% ! Partial therapy  
[pan12 pan24 resp12\$1 resp24\$1];  
pan12 pan24 (4);  
pan12 ON treat\*0 (1);  
pan24 ON treat\*0 (1);  
resp12 ON treat\*0 (6);  
resp24 ON treat\*0 (6);

%C#3% ! Full therapy  
[pan12 pan24 resp12\$1 resp24\$1];  
pan12 pan24 (5);  
pan12 ON treat\*0 (2);  
pan24 ON treat\*0 (2);  
resp12 ON treat\*0 (7);  
resp24 ON treat\*0 (7);