

*Randomized controlled trial of the effect of
a home visiting intervention on infant
cognitive development in peri-urban South
Africa*

Article

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**Randomized controlled trial of the effect of a home visiting intervention on infant
cognitive development in peri-urban South Africa**

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Summary

Aim To determine whether, in an impoverished South African community, an intervention that benefitted infant attachment also benefitted cognitive development.

Method Pregnant women were randomized to intervention (220) and no treatment control groups (229). The intervention was home-based parenting support for attachment, delivered until six months postpartum. At 18 months, infants were assessed on attachment⁶, and cognitive development (Bayley MDI) (127 intervention, 136 control). Infant MDI was examined in relation to intervention, socio-economic risk, antenatal depression, and infant sex and attachment.

Results Overall, there was little effect of the intervention on MDI ($p=.094$, $d=0.20$), but there was an interaction between intervention and risk ($p=.03$, $\eta_p^2=.02$): MDI scores of infants of lower risk intervention group mothers were, on average, 4.84 points higher than those of other infants ($p=.002$, $d=.41$). Antenatal depression was not significant once intervention and risk were controlled ($p=.08$); there was no association between infant MDI and either sex ($p=.41$) or attachment ($p=.56$).

Conclusion Parenting interventions for infant cognitive development may benefit from inclusion of specific components to support infant cognition, beyond those that support attachment, and may be most effective for infants over six months. They may need augmentation with other input where adversity is extreme.

Keywords: Infants, cognitive development, intervention, parenting, Low and Middle Income Country

Abbreviations: MDI Mental Development Index. SCID Structured Clinical Interview for DSM-IV diagnoses

What this paper adds

An intervention that benefitted infant attachment, only benefitted cognitive development where families did not experience severe socio-economic adversity

In the context of such adversity, psychological interventions for infant cognitive development may need augmentation with other input

Intervention for cognitive development may be more effective if children are over six months

Raised rates of parenting difficulties occur in the context of the poverty and mental health problems that commonly obtain in low and middle income countries (LMIC)^{1,2}. These are, in turn, associated with problems in infant psychological development, such as insecure attachment and poor cognitive functioning³. These early developmental difficulties are important, since they tend to endure and they predict a range of problems that affect children's life course trajectories (e.g., conduct disorder, educational failure and employment prospects⁴). Accordingly, interventions are needed that target parenting in infancy. Moreover, in LMIC contexts, it is important that interventions are low cost and use readily available resources. In our earlier epidemiological work in a disadvantaged peri-urban settlement in South Africa, Khayelitsha, we found high rates of maternal depression, parenting difficulties and insecure infant attachment^{3,5}. Subsequently, in a randomized controlled trial (RCT), we showed benefits to these outcomes of a home-visiting programme, delivered by lay community workers from late pregnancy through the first six months postpartum⁶.

An important question is whether the benefit of our intervention to infant attachment security extended to other infant psychological outcomes, and in particular cognitive development. Establishing the limits of interventions' effectiveness, as well as their benefits, is important in informing policy and practice. Child cognitive performance in South Africa is of considerable concern: in 2011, fewer than half Grade 3 children achieved the basic educational level considered acceptable⁷, and in an international review of 9-10 year olds' literacy, including several LMIC's, South Africa was at the bottom of the performance table⁸.

Child cognitive performance from late infancy is a good predictor of later cognitive functioning⁴. If our intervention did indeed benefit infant cognitive functioning as well as attachment security, this would argue for its being implemented without substantial modification, in order to be of relatively general benefit to infant psychological

development. Notably, however, our intervention was targeted at caretaking of particular relevance to *attachment* (e.g., responsiveness to infant distress), and its benefit to domains of infant functioning such as cognitive performance may have been more limited. Indeed, there is increasing recognition of the *specificity* of associations between different parenting qualities and particular child outcomes^{9,10}. Of special relevance to the question addressed here are findings that the parenting qualities most relevant to attachment (protection and comforting when infants are distressed or vulnerable) do not necessarily predict child cognitive outcome and, *vice versa*, that parental support for child cognitive achievements (e.g., scaffolding and guided learning) does not necessarily promote socioemotional developments such as attachment security^{9,10}. Notably, effective interventions for infant and child cognitive outcome in LMIC's^{11,12,13} have often involved parents being coached in active play and stimulation of their child, and far less evidence is available concerning the cognitive benefits of non-cognitively focussed parenting curricula, particularly for children under two years. The current paper addresses this question. We report the effects of our intervention on infant performance on a standard measure of infant cognitive development, the Mental Development Index (MDI) of the Bayley II Scales¹⁴. This was assessed concurrently with attachment security at 18 months.

Studies of interventions for child development, including cognitive outcomes, have found background risk to be relevant. In high income countries (HICs), benefits of home-visiting are particularly clear in the context of greater socio-economic risk¹⁵. However, although such associations have been found in LMIC's or generally disadvantaged populations¹¹, there is some evidence that those at greatest risk (e.g., through low education) may benefit *less* from parenting programmes^{1,11,16}. Such differential effects are important to determine so that interventions can be appropriately targeted. In Khayelitsha, despite poverty being ubiquitous, living conditions vary (e.g., in provision of water, electricity), and

therefore we examined whether intervention effects differed according to the level of socio-economic risk.

In addition to postnatal influences on child cognitive development, antenatal depression has been found to pose a direct risk¹⁷, and so we investigated its effects too. Further, since some studies report male infants' cognitive development to be more vulnerable to effects of adversity than females'¹⁸, we investigated direct and moderating effects of infant sex. Finally, given the specificity of parenting effects on different aspects of child psychological development, we examined associations between infant cognitive outcome and attachment security.

Method

Design

The study design is described in detail in our previous report on the effects of our intervention on infant attachment⁵ (Trial registration number: ISRCTN25664149). The CONSORT is shown in Figure 1. It was conducted in two adjoining areas ('SST' and 'Town II') in Khayelitsha, a disadvantaged peri-urban settlement near Cape Town, South Africa. SST is an informal settlement characterized by high levels of unemployment (two-thirds of the population) and poverty (shacks without electricity or running water); in Town II the standard of living is somewhat better. It was an RCT in which pregnant women were randomly assigned to the intervention group or a no treatment control group. The intervention was delivered in mothers' homes by trained community workers from the third trimester of pregnancy until six months postpartum. Infant cognitive development was independently assessed at 18 months. The study was approved by the research ethics committees of the University of Reading and the Health Sciences faculty of the Medical School of the University of Cape Town, and participants gave written informed consent.

Figure 1 CONSORT diagram for the RCT about here

Participants

House to house visits were made at three-weekly intervals in the study area over a 22 month period. Four hundred and fifty-two pregnant women were identified, and invited to participate. All but three agreed. After women gave consent, demographic variables and socio-economic risk factors were recorded, and mothers were randomly assigned to either the intervention (N= 220) or control group (N=229) by minimisation, balancing for antenatal depression, planned pregnancy and housing area (see CONSORT). We anticipated substantial participant loss because many mothers travel from rural areas to deliver their infants, and then return; indeed, approximately one fifth of the sample could not be followed up because mothers moved away, or their infants died (see CONSORT). Of those originally enrolled, 342 (76%) were assessed at 18 months (165 in the intervention group, 177 in the control group). Of these, attachment to the mother was assessed in 318 (93%) infants (156 intervention, 162 control)⁶, and cognitive development was assessed in 263 (77%) (127 intervention, 136 control).

The intervention

Four trained home visitors visited the mother twice in pregnancy, and then on 14 occasions up to six months postpartum (75% mothers received all 16 visits, and over 90% received at least eight). Visits lasted one hour. The home visitors all lived in Khayelitsha and were mothers themselves. Two had completed schooling; none had education or training beyond school. The intervention was manualized⁵. It included key principles of the WHO's 'Improving the Psychosocial Development of Children', and of 'The Social Baby'²⁰. It provided psychological support to the mother, using counselling, and items from the Neonatal

Behavioural Assessment Scale²¹ to enhance maternal awareness of infant social engagement, and provide strategies for managing infant distress. The home visitors received three weeks of training in the intervention over a four-month period, and weekly group supervision from a community clinical psychologist.

All mothers (control and intervention) received fortnightly visits by a community health worker from a local NGO that monitored maternal and infant health.

Measures

Socio-economic risk. We used socio-economic risk indices employed in previous research, namely, teenage parenthood, unplanned pregnancy, less than seven years education¹⁷, and indices more specific to our sample: poor partner support, lack of electricity at home, and additional children.

Antenatal depression. Antenatal depression (and two and six months depression) was assessed by the Structured Clinical Interview for DSM-IV diagnoses (the SCID)²². This interview has good reliability and validity, and has been widely used in South Africa. Interviews were administered by a trained researcher and audiotaped. Diagnoses were independently confirmed with a senior clinician.

Infant cognitive development. The Bayley Scales, version II¹⁴, was administered at 18 months and the Mental Development Index (MDI) used as the measure of cognitive outcome. The assessment was conducted in research premises in Khayelitsha, by a trained researcher who was blind to group. Since the principal outcome was attachment security, this was assessed first, followed by a break. Subsequently, if the infant's state permitted, the MDI was administered.

Statistical analysis

We examined equivalence in socio-demographic characteristics between intervention and control groups using t test or chi-square, as appropriate. Next, we examined relationships between individual risk factors. Given their associations, a composite was created. Then, using Analysis of Variance (ANOVAs) and Analysis of Covariance (ANCOVAs), we examined effects of the intervention on child cognitive development, including socio-economic risk as a potential moderator, as well as relevant covariates (antenatal and postnatal depression). We also explored effects of individual risk factors through further ANOVAs, and bivariate correlations, using bootstrapping with bias correction to adjust for multiple comparisons. Adjusted standard errors and adjusted confidence intervals are provided. Equality of variances was tested through Levene's test, and in all cases data met assumption criteria. Finally, in secondary analyses, we examined effects (main and potential moderating) of infant sex, and whether infant cognitive development and attachment security showed the same pattern of relationship to treatment and risk and were associated with each other. Effect sizes were computed, with d being calculated for main effects and partial eta-squared (η^2_p) for ANOVA interactions.

Results

Participant characteristics

Demographic and risk factors are shown in Table 1. The current sample did not differ from those originally recruited, nor from all those with attachment assessments. (Infant attachment for the slightly smaller sample with cognitive assessments showed the same benefit of the intervention as the full sample (Wald = 7.8, df = 1, OR=2.2 (95% Confidence Interval (95% CI) = 1.3-3.8), $p = .005$). There were no differences at base-line in demographic and risk

factors between intervention and control groups but, as predicted, at two months postpartum, compared to control group mothers, fewer intervention group mothers were depressed.

Table 1 about here

Intervention effects.

There was a trend for intervention group infants overall to have higher MDI scores than those of control group infants ($F(1, 261) = 2.8, p = .09, d=.20$). Means (and Confidence Intervals) were, respectively, 85.2 (95% CI = 83.4-87.1) and 83.1 (95% CI = 81.3-84.8).

Level of socio-economic risk as a moderator of intervention effects.

The six socio-economic risk factors showed a number of significant associations with one another (e.g., less education was associated with and having other children ($\phi = .19, p=.002$), unplanned pregnancy was associated with both being a teenager ($\phi = .27, p=.0001$), and lack of partner support ($\phi = .22, p = .0001$)). The six factors were therefore aggregated into a composite risk measure for each mother, which was then averaged; these continuous risk scores were then converted into a binary variable, using a median split (median = 0.7, IQR = 0.3) to specify groups at either higher, or else relatively lower, risk. For the current sample with cognitive assessments, 106 mothers (40%) were in the higher risk group. There were no differences in rates of high risk between intervention and control groups (42% and 39%, respectively $\chi^2(1) = 0.2, p = .71$).

The effect of risk status on the Bayley MDI scores was not significant ($p = .17, d=.17$). Nevertheless, there was a significant interaction between risk and intervention ($p=.03, \eta_p^2 = .02$): infants whose mothers had relatively lower socio-economic risk had significantly better MDI scores if their mothers received the intervention than the other groups of infants,

the difference, on average being 4.8 scale points (see Table 2a and Figure 2). When the scores for the lower risk intervention group were compared with those for the three other groups combined, the effect size was $d = .41$, and when compared to scores of only lower risk controls, $d = .42$ (both small effects). When both level of risk and the above interaction effect were entered in the model, the effect of intervention was no longer marginally significant ($F(1, 259) = 1.6, p = .20, \eta_p^2 = .006$), indicating that the benefit of intervention was carried by the group where mothers had lower risk.

Table 2 and Figure 2 The interaction effect of risk status and the intervention on child cognition. [about here](#)

To investigate whether the moderating effect of overall risk status was carried by specific factors, the risk composite was disaggregated, and main, and interaction (with intervention group) effects on the MDI were examined for each one. One significant main effect was observed, with infants whose mothers had electricity at home having significantly higher MDI scores than those of mothers without electricity ((electricity $M = 86.3$ (95% CI = 84.5-88.1), no electricity $M = 82.2$ (95% CI = 80.5-83.9)). Further, examination of the interaction showed that the benefit to infant cognitive outcome of the intervention applied *only* to those having electricity in the home (see Table 2b).

Antenatal depression

Compared to infants of non-depressed women, those whose mothers were antenatally depressed had significantly lower MDI scores ($F(1, 262) = 4.4, p = .04, d = .33$: antenatally depressed $M = 81.4$ (95% CI = 78.6-84.3), not antenatally depressed $M = 84.8$ (95% CI =

83.4-86.2)). (This effect still held when depression at two and six months was included as a covariate, $p = .03$ and $p = .05$, respectively). Accordingly, we examined whether the effect on infant MDI of the intervention, in interaction with maternal risk status, was still significant having taken antenatal depression into account. As shown in Table 3, the effect of the interaction between group and risk was unchanged, while that of antenatal depression was somewhat reduced, although a trend effect remained. Antenatal depression was also examined as potential moderator of the intervention effects; results did not show a significant interaction antenatal depression*intervention group ($F(1, 259) = 1.0, p = .32, \eta_p^2 = .004$).

Table 3 about here

Infant sex

Infant sex had no effect on Bayley MDI scores, either alone ($F(1, 259) = 0.7, p = .41, d=.13$), or in interaction with maternal risk ($F(1, 259) = 0.9, p = .34, \eta_p^2 = .004$). Similarly, infant sex did not significantly moderate the effects of intervention ($F(1, 249)=0.04, p = .84, \eta_p^2 = .0001$).

Infant cognitive outcome and attachment

Our principal finding that the beneficial effect of the intervention on infant cognitive performance was confined to families with a lower level of risk raised the question of whether a similar relationship between intervention and risk also applied to infant attachment. This was not the case, with the interaction between intervention and risk on infant attachment being non-significant (Wald = 0.03, df = 1, OR=0.9(95%CI 0.3-2.8), $p > .8$). Indeed, there was no association between infant attachment security and cognitive outcome, mean MDI scores for secure vs. insecure infants being 83.9 and 84.7 respectively ($p = .56$).

Discussion

In a socio-economically deprived peri-urban settlement in South Africa, a home visiting intervention, delivered by community workers to mothers in pregnancy and the first six postpartum months had no overall effect on infant cognition at 18 months in contrast to its benefit to attachment. Nevertheless, for those not living in conditions of particularly high socio-economic risk (principally those in dwellings with electricity), the intervention was of benefit to infant cognitive development. In addition to intervention and socio-economic risk effects, infants whose mothers were antenatally depressed tended to have lower cognitive scores.

A number of aspects of our intervention require comment. First, with regard to the overall lack of benefit of our intervention on infant cognitive development, it is important to bear in mind that the intervention was focused on providing the mother with psychological support and help in her attachment relationship with her infant (i.e., supporting the management of infant distress, and sensitizing mothers to infant social cues and attachment needs). Although some of the parenting qualities that help promote secure attachment are also relevant to child cognitive development (e.g., general responsiveness), other parenting practices were absent from our intervention which are known to be of specific benefit to child cognitive functioning. These are guided learning, verbal stimulation and the ‘scaffolding’ of infant attention and engagement with the environment. Thus, a more cognitively focused intervention than ours, with clearer didactic elements and encouragement to parents to practise, may have produced greater cognitive gains for the infants.

A second possible barrier to the effectiveness of our intervention, aside from its attachment *vs.* cognitive focus, was infant age: it is possible that parenting support for infant cognitive development is more effective when infants are older than those in our intervention- i.e., over six months, when infant attention and motor skills have developed

sufficiently to enable active engagement with the wider environment, and provide more opportunities for parents to facilitate infant cognitive skills. Indeed, successful interventions for infant cognitive development in LMICs have generally provided support for children up to one-to-three years of age^{12,13,23}.

A further aspect of our findings requiring comment is that the failure of the intervention to benefit infant cognitive development applied principally to infants of mothers experiencing particularly high levels of socio-economic risk. This may have been because of factors on either the mothers' or the infants' part. Thus, the extremes of adversity facing these mothers may have prevented them from engaging effectively with the intervention. Similarly, there may have been unmeasured effects of higher risk on the infants of these mothers, such as nutritional deficiencies or recurrent infections (e.g., respiratory and gastro-intestinal), that may have meant they could not benefit cognitively from the intervention. In either case, and in line with other studies^{11,13}, our results suggest that those in particularly adverse circumstances may need to receive broader support (e.g., economic, or nutritional), if child cognitive development is to improve in the context of home-visiting.

Finally, consistent with reports from HICs¹⁷, our findings suggest that it may be useful to tackle antenatal, as well as postnatal, maternal depression. Encouragingly, recent programmes show that screening for depression in pregnancy can be successfully integrated into primary level maternity services in LMIC contexts²⁴, and lay community workers can be trained to deliver support that is effective in reducing depression²⁵. Future research would benefit, therefore, from assessing whether interventions for antenatal depression can improve infant cognitive functioning.

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References

1. Walker, S. P., Wachs, T. D., Grantham-McGregor, S., Black, M. M., Nelson, C. A., Huffman, S. L., & Richter, L. Inequality in early childhood: risk and protective factors for early child development. *The Lancet* 2011; 378, 1325-1338.
2. Grantham-McGregor, S., Cheung, Y. B., Cueto, S., Glewwe, P., Richter, L., & Strupp, B. Developmental potential in the first 5 years for children in developing countries. *The Lancet* 2007; 369, 60-70.
3. Tomlinson, M., Cooper, P., & Murray, L. The Mother–Infant Relationship and Infant Attachment in a South African Peri-Urban Settlement. *Child Development* 2005; 76, 1044-1054.
4. Bornstein, M. Human infancy...and the rest of the lifespan. *Annual Review of Psychology* 2014; 65, 121-58

5. Cooper, P. J., Tomlinson, M., Swartz, L., Woolgar, M., Murray, L., & Molteno, C. Post-partum depression and the mother-infant relationship in a South African peri-urban settlement. *The British Journal of Psychiatry* 1999; 175, 554-558.
6. Cooper, P. J., Tomlinson, M., Swartz, L., Landman, M., Molteno, C., Stein, A., ... & Murray, L. Improving quality of mother-infant relationship and infant attachment in socioeconomically deprived community in South Africa: randomised controlled trial. *BMJ: British Medical Journal* 2009; 338, b974.
7. Department of Basic Education. *Report on the Annual National Assessments of 2011*. Pretoria Department of Basic Education, South Africa 2011
8. Mullis, I. V. *PIRLS 2006 international report: IEA's Progress in International Reading Literacy Study in primary school in 40 countries*. TIMSS & PIRLS International Study Center, Boston College 2007.
9. Grusec, J. E., & Davidov, M. Integrating different perspectives on socialization theory and research: A domain-specific approach. *Child Development* 2010; 81, 687-709.
10. Bornstein, M., & Plutchik, D. Cognitive and socio-emotional caregiving in developing countries. *Child Development* 2012; 83, 46-61
11. Engle, P. L., Fernald, L. C., Alderman, H., Behrman, J., O'Gara, C., Yousafzai, A., ... & Iltis, S. Strategies for reducing inequalities and improving developmental outcomes for young children in low-income and middle-income countries. *The Lancet* 2011; 378, 1339-1353.
12. Baker-Henningman, H. The role of early childhood education programmes in the promotion of child and adolescent mental health in low and middle income countries. *International Journal of Epidemiology* 2014; 43, 407-433

13. Grantham-McGregor, S., Fernald, L., Kagawa, R., & Walker, S. Effects of integrated child development and nutrition interventions on child development and nutrition status. *Annals of the New York Academy of Sciences* 2014; 1308,11-32
14. Bayley, N. *Bayley Scales of Infant Development: Manual*. New York: Psych. Corporation 1993.
15. Olds, D. L., Kitzman, H., Cole, R., Robinson, J., Sidora, K., Luckey, D. W., ... & Holmberg, J. Effects of nurse home-visiting on maternal life course and child development: Age 6 follow-up results of a randomized trial. *Pediatrics* 2004; 114, 1550-1559.
16. Aboud, F. Evaluation of early childhood parenting programme in rural Bangladesh. *Journal of Health Population Nutrition* 2007; 25 (1) 3-13
17. Evans, J., Melotti, R., Heron, J., Ramchandani, P., Wiles, N., Murray, L., & Stein, A. The timing of maternal depressive symptoms and child cognitive development: a longitudinal study. *Journal of Child Psychology and Psychiatry* 2012; 53, 632-640.
18. Murray, L., Halligan, S., & Cooper, P. Effects of postnatal depression on mother–infant interactions and child development. *Wiley-Blackwell Handbook of Infant Development* 2010; *The*, Vol 2, Second edn., 192-220.
19. World Health Organization. *Improving the Psychosocial Development of Children*. Geneva 1995.
20. Murray, L., & Andrews, L. *The Social Baby: Understanding babies' communication from birth*. Surrey: CP Publishing 2000.
21. Brazelton, T. B., & Nugent, J. K. (Eds.). *Neonatal behavioral assessment scale* (No. 137). Cambridge University Press 1995.

22. First, M. B., Spitzer, R. L., Gibbon M., & Williams, J. B.W. *Structured Clinical Interview for DSM-IV Axis I Disorders, Clinician Version (SCID-CV)*. Washington, D.C.: American Psychiatric Press Inc. 1996.
23. Wallander, J., Bann, C., Biasini, F., Goudar, S., Pasha, O., Chomba, E., McClure, E., & Carlo, W. Development of children at risk for adverse outcomes participating in early intervention in developing countries: a randomized controlled trial. *Journal of Child Psychology and Psychiatry* 2014; 55, 1251-1259
24. Honikman, S., van Heyningen, T., Field, S., Baron, E., & Tomlinson, M. Stepped care for maternal mental health: a case study of the perinatal mental health project in South Africa. *PLoS Medicine* 2012; 9, e100122
25. Rahman, A., Malik, A., Sikander, S., Roberts, C., & Creed, F. Cognitive behaviour therapy-based intervention by community health workers for mothers with depression and their infants in rural Pakistan: a cluster-randomised controlled trial. *The Lancet* 2008; 372, 902-909.

Table 1 Demographic and risk factors.

| | Intervention N=127 | Control N=136 |
|------------------------------------------------------|-----------------------|------------------|
| Mother & Family Socio-demographics | | |
| Maternal Age | | |
| Mean | 25.6 | 26.6 |
| (95% confidence interval) | (24.6-26.6) | (25.5-27.6) |
| Range | 15-39 | 16-43 |
| *Teenage mother | 24(19%) | 27(20%) |
| Maternal Education | | |
| *Educated for 6 years or less | 35 (28%) | 41 (30%) |
| Marital status and support | | |
| Married | 56 (44%) | 60 (44%) |
| *No partner support | 31 (25%) | 29 (21%) |
| Maternal depression | | |
| Antenatal | 24 (19%) | 27 (20%) |
| 2 months postnatal | 23 (19%) | 39 (30%) |
| 6 months postnatal | 13 (11%) | 22 (18%) |
| Housing | | |
| *No electricity | 69 (54%) | 70 (51%) |
| Living Area | | |
| SST | 71 (56%) | 72 (53%) |
| Town II | 56 (44%) | 64 (47%) |
| Pregnancy History & Child Characteristics | | |
| Planned pregnancy | | |
| *Unplanned pregnancy | 52 (41%) | 50 (37%) |
| Other children | | |
| *Not primiparous | 68 (55%) | 77 (58%) |
| Child Sex | | |
| Male | 70 (55%) | 68 (50%) |

Child birthweight

| | | |
|-------------------------|-----------|-----------|
| Birthweight in kg. Mean | 3.1 | 3.1 |
| (95% CI) | (3.0-3.2) | (3.0-3.2) |

*variables contributing to the risk measure

Table 2. Risk and interaction effects of risk composite and individual risk factors on the Bayley MDI

| | Intervention | | Control | | Risk Effect | Interaction Risk with intervention group |
|------------------------------------|--------------|-------------------|------------|-------------------|---------------------|------------------------------------------|
| | M(sd) | SEadj [95% adjCI] | M(sd) | SEadj [95% adjCI] | | |
| 2a. Risk Composite | | | | | | |
| high risk | 82.4(9.9) | 1.4[79.5-85.3] | 83.7(9.6) | 1.3[80.9-86.4] | F(1, 261) = 1.9 | F(1, 259) = 5.0 |
| low risk | 87.3(11.2) | 1.3[84.7-89.6] | 82.7(10.2) | 1.1[80.5-85.3] | $p = .17, d = .17$ | $p = .03, \eta_p^2 = .02$ |
| 2b. Individual risk factors | | | | | | |
| Teenage mother | | | | | F(1, 258) = 1.0 | F(1, 256) = 1.5 |
| Teenager | 88.3(10.9) | 2.2[83.6-92.9] | 82.9(8.6) | 1.6[79.6-86.0] | $p = .31, d = .15$ | $p = .22, \eta_p^2 = .006$ |
| Not teenager | 84.5(10.8) | 1.1[82.5-86.4] | 83.1(10.3) | 1.0[81.1-85.0] | | |
| Educated 6yrs or less | | | | | F(1, 261) = 0.3 | F(1, 259) = 0.04 |
| 6 yrs or less | 84.5(11.2) | 1.9[80.9-88.2] | 82.8(10.9) | 1.7[79.7-86.0] | $p = .56, d = .07$ | $p = .84, \eta_p^2 = .0001$ |
| 7 yrs or more | 85.5(10.8) | 1.1[83.6-87.7] | 83.2(9.6) | 1.0[81.5-85.2] | | |
| No partner support | | | | | F(1, 260) = 0.03 | F(1, 258) = 2.1 |
| No help | 83.4(12.0) | 2.2[79.0-87.6] | 84.6(9.2) | 1.8[80.9-88.2] | $p = .86, d = .02$ | $p = .14, \eta_p^2 = .008$ |
| Some help | 85.9(10.5) | 1.1[84.0-88.0] | 82.7(10.2) | 1.0[80.8-84.6] | | |
| No electricity | | | | | F(1, 261) = 10.2 | F(1, 259) = 11.8 |
| No electricity | 81.4(9.2) | 1.1[79.2-83.6] | 83.1(9.7) | 1.2[80.9-85.2] | $p = .002, d = .39$ | $p = .001, \eta_p^2 = .04$ |
| Electricity | 89.9(10.9) | 1.5[87.0-92.7] | 83.1(10.3) | 1.3[80.5-85.8] | | |
| Unplanned pregnancy | | | | | F(1, 261) = 0.02 | F(1, 259) = 3.4 |
| Unplanned | 83.9(9.4) | 1.3[81.3-86.2] | 84.7(8.7) | 1.2[82.3-87.4] | $p = .86, d = .02$ | $p = .06, \eta_p^2 = .01$ |
| Planned | 86.2(11.8) | 1.3[83.5-88.6] | 82.2(10.6) | 1.1[80.2-84.2] | | |
| Not primiparous | | | | | F(1, 254) = 2.2 | F(1, 252) = 2.3 |
| No primiparous | 83.7(10.2) | 1.2[81.3-86.2] | 82.9(9.6) | 1.1[80.8-85.1] | $p = .14, d = .18$ | $p = .13, \eta_p^2 = .009$ |
| Primiparous | 87.5(11.1) | 1.5[84.3-90.6] | 82.8(10.4) | 1.4[79.9-85.9] | | |

Note: Bootstrapping with bias correction was applied to correct for multiple comparisons; adjusted standard error (SEadj) and 95% Confidence Intervals (95%CIadj).

Table 3. Effects of risk status on Bayley MDI controlling for antenatal depression

| | F | P | η^2 |
|--------------------------------|-----|-----|----------|
| Antenatal depression | 3.1 | .08 | .01 |
| Intervention Group | 1.5 | .21 | .01 |
| Risk status | 1.0 | .30 | .004 |
| Intervention Group*risk status | 5.1 | .02 | .02 |