

# *Quantifying leaf trait covariation and its controls across climates and biomes*

Article

Accepted Version

Yang, Y., Wang, H., Harrison, S. P. ORCID:  
<https://orcid.org/0000-0001-5687-1903>, Prentice, I. C., Wright,  
I. J., Peng, C. and Lin, G. (2019) Quantifying leaf trait  
covariation and its controls across climates and biomes. *New  
Phytologist*, 221 (1). pp. 155-168. ISSN 1469-8137 doi:  
<https://doi.org/10.1111/nph.15422> Available at  
<https://centaur.reading.ac.uk/78149/>

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To link to this article DOI: <http://dx.doi.org/10.1111/nph.15422>

Publisher: Wiley

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1 **Quantifying leaf trait covariation and its controls across**  
2 **climates and biomes**

3

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25 Revised version for *New Phytologist*

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## 29 **Summary**

- 30 • Plant functional ecology requires the quantification of trait variation and its  
31 controls. Field measurements on 483 species at 48 sites across China were used to  
32 analyse variation in leaf traits, and assess their predictability.
- 33 • Principal components analysis (PCA) was used to characterize trait variation,  
34 redundancy analysis (RDA) to reveal climate effects, and RDA with variance  
35 partitioning to estimate separate and overlapping effects of site, climate, life-form  
36 and family membership.
- 37 • Four orthogonal dimensions of total trait variation were identified: leaf area (LA),  
38 internal-to-ambient CO<sub>2</sub> ratio ( $\chi$ ), leaf economics spectrum traits (specific leaf  
39 area (SLA) *versus* leaf dry matter content (LDMC) and nitrogen per area ( $N_{\text{area}}$ )),  
40 and photosynthetic capacities ( $V_{\text{cmax}}$ ,  $J_{\text{max}}$  at 25°C). LA and  $\chi$  covaried with  
41 moisture index. Site, climate, life form and family together explained 70% of trait  
42 variance. Families accounted for 17%, and climate and families together 29%  
43 LDMC and SLA showed the largest family effects. Independent life-form effects  
44 were small.
- 45 • Climate influences trait variation in part by selection for different life forms and  
46 families. Trait values derived from climate data via RDA showed substantial  
47 predictive power for trait values in the available global data sets. Systematic trait  
48 data collection across all climates and biomes is still necessary.

49

50 **Key words:** climate, leaf economics spectrum, multivariate analysis, photosynthetic  
51 capacity, phylogeny, plant functional traits.

52

## 53 **Introduction**

54 Functional traits generally do not vary independently, but show broadly predictable  
55 patterns of covariation (Armbruster *et al.*, 1996; Watson *et al.*, 2016). The covariation  
56 of traits may mean that traits share genetic controls, or that they have related roles in  
57 community assembly and function (Wright *et al.*, 2007; Fajardo *et al.*, 2011).  
58 Quantifying the covariation of vegetative traits and their controls is important for an  
59 understanding of how plants drive ecosystem processes and determine the responses  
60 of ecosystems to environmental change (Wright *et al.*, 2007; Shipley *et al.*, 2011;  
61 Swenson 2013; van Bodegom *et al.*, 2014; Kong *et al.*, 2014; Kraft *et al.*, 2015).  
62 Although a number of large-scale studies have quantified both trait covariation (e.g.  
63 Wright *et al.*, 2004; Armbruster *et al.*, 2014; Peiman & Robinson, 2017) and  
64 trait-environment relationships,(e.g. Wright *et al.*, 2005; Harrison *et al.*, 2010; Liu *et al.*,  
65 2012; Maire *et al.*, 2015; Meng *et al.*, 2015), a number of general issues await  
66 resolution. These include:

67 (1) The dimensionality of trait space – that is, the extent to which combinations of  
68 different traits are independent, *versus* belonging to a set of covarying traits as  
69 exemplified by the leaf economics spectrum (LES) (Wright *et al.*, 2004, 2005). The  
70 intrinsic dimensionality of traits is the minimum number of independent axes that  
71 adequately describe the functional variation among species, and is therefore an  
72 important quantity in comparative ecology (Laughlin, 2014).

73 (2) The extent to which trait variation is determined by climate, versus the  
74 co-existence of multiple trait values in the same climate (Adler *et al.*, 2013;  
75 Valladares *et al.*, 2015).

76 (3) The extent to which trait variation and trait-environment correlations are linked to  
77 ‘hard-wired’ physiognomic (life-form) and/or phylogenetic differences among species,  
78 and the role of environment in selecting among life forms and clades (Díaz *et al.*,  
79 2013; Ackerly, 2009; Donovan *et al.*, 2014).

80 The dimensionality question has received attention in plant functional ecology partly  
81 because of the universal nature of the LES, which is considered as the outcome of a  
82 tradeoff between resource acquisition and conservation – representing different  
83 general strategies for existence, rather than adaptations to environment (Wright *et al.*,  
84 2007; Kong *et al.*, 2014; Reich, 2014). An early synthesis led to a proposal for four  
85 trait dimensions indexed by leaf mass per area and lifespan (i.e. the LES), seed mass  
86 and seed output, leaf and twig size, and plant height (Westoby *et al.*, 2002). Wright *et al.*  
87 *al.* (2007) found three independent trait dimensions represented by specific leaf area  
88 (SLA), seed/fruit size and leaf size in seven neotropical forests. The most extensive  
89 study (in terms of the number of species considered) to date was by Díaz *et al.* (2016),  
90 who showed that variation among species in height, stem specific density, leaf mass  
91 per area, seed mass, and nitrogen per unit mass ( $N_{\text{mass}}$ ) could be reduced to two  
92 dimensions, the first indexing plant size, the second the LES. However, these various  
93 studies have considered only a limited set of traits or combined information from  
94 disparate sources, and did not attempt to quantify the climatic or phylogenetic controls  
95 on traits.

96 In this paper, we examine a suite of leaf traits, using co-located measurements to  
97 quantify the contributions of climate, site, life form and phylogeny to trait variation at a  
98 large geographic scale. Our analysis is based on an extensive data set (Wang *et al.*,  
99 2018), containing information on multiple leaf traits from different regions of China.  
100 We focused on seven leaf traits that together capture many functions of plants (Table  
101 S1). The traits considered include four commonly measured traits: leaf area (LA),  
102 specific leaf area (SLA), leaf dry matter content (LDMC) and leaf nitrogen per unit  
103 area ( $N_{\text{area}}$ ), and also three traits that determine photosynthetic rates: maximum  
104 carboxylation rate ( $V_{\text{cmax}}$ ) and maximum electron transport rate ( $J_{\text{max}}$ ), derived from  
105 gas exchange measurements in the field, and the ratio of intercellular to ambient  
106 carbon dioxide ( $\text{CO}_2$ ) concentration (often denoted as  $c_i:c_a$  but called  $\chi$  here following  
107 Prentice *et al.*, 2014) derived from leaf stable carbon isotope ( $\delta^{13}\text{C}$ ) measurements.

108 We used multivariate analysis to quantify the dimensionality of variation in this set of  
109 traits, and the nature and dimensionality of trait-climate relationships. We used  
110 variance partitioning to attribute trait variations (for all traits, and each trait separately)  
111 to differences among sites, climate variations across sites, and distinctions among life  
112 forms and plant families. We finally applied the trait-climate relationships derived  
113 from the data set to various global datasets for specific traits, in order to assess their  
114 generality and potential wider application.

## 115 **Materials and methods**

### 116 **Dataset description**

117 The data are derived from the China Plant Trait Database (Wang *et al.*, 2018), which  
118 contains information on morphological, physical, chemical and photosynthetic traits  
119 from 122 sites and provides information on more than 1215 species. The database was  
120 designed to provide comprehensive sampling of different vegetation types and  
121 climates. It employs a standardized taxonomy and includes information on life form,  
122 plant family, site location, elevation, and climate. LA, SLA,  $N_{\text{area}}$ , LDMC and leaf  
123  $\delta^{13}\text{C}$  data from multiple species were available at 48 sites, including 483 species  
124 altogether, distributed through the eastern half of China (Fig. 1a, Table S2). The sites  
125 from northeastern China are distributed along an aridity gradient (Prentice *et al.*,  
126 2011), including steppes, grasslands and temperate deciduous broadleaf forests. The  
127 sites from southwestern China represent tropical and subtropical evergreen broadleaf  
128 forests, and tropical dry woodlands. Temperate deciduous forests in central China and  
129 boreal forests in the far north of China were also included. Collectively these data  
130 cover the principal climatic and vegetation zones of the region (Fig. 1b). At each site,  
131 a stratified sampling strategy ensured that measurements were available for the main  
132 species in each canopy stratum, including up to 25 species of trees. Species were  
133 classified by life form as trees, small trees, lianas, shrubs, forbs and graminoids.  
134 Bamboos, herbaceous climbers, geophytes and pteridophytes were present only in

135 small numbers in the dataset and were not included in our analysis. Fig. S1 shows  
136 frequency distributions of each trait within each life form for forest and non-forest  
137 sites. Table S3 lists the total number of samples in each class.

138 Details of trait measurement methods can be found in Wang *et al.* (2018). LA, SLA,  
139  $N_{\text{area}}$  and LDMC were measured on samples collected in the field following standard  
140 protocols (Cornelissen *et al.*, 2003). LA was taken as the projected area of a leaf, or  
141 leaflet in the case of compound leaves.  $V_{\text{cmax}}$  was calculated from the light-saturated  
142 rate of net CO<sub>2</sub> fixation at ambient CO<sub>2</sub> ( $A_{\text{sat}}$ ) using the so-called one-point method,  
143 which provides a rapid and effective alternative to the measurement of a full  $A$ - $c_i$   
144 curve (De Kauwe *et al.*, 2016).  $J_{\text{max}}$  was calculated from the light-saturated rate of net  
145 CO<sub>2</sub> fixation at high CO<sub>2</sub> ( $A_{\text{max}}$ ). Both  $V_{\text{cmax}}$  and  $J_{\text{max}}$  were adjusted to a standard  
146 temperature of 25°C using the methods proposed by Niinemets *et al.* (2014). The  
147 adjusted values are called  $V_{\text{cmax}25}$  and  $J_{\text{max}25}$ . Leaf  $\delta^{13}\text{C}$  measurements were converted  
148 to <sup>13</sup>C discrimination and thence to  $\chi$ , eliminating the effects of latitude and sampling  
149 year as described in Cornwell *et al.* (2017):

$$150 \quad \delta^{13}\text{C}_{\text{air},1992} = a * \left( \sin \left( \varphi * \frac{\pi}{180} \right) \right)^2 + \sin \left( \varphi * \frac{\pi}{180} \right) - c \quad (1)$$

151 where  $\varphi$  is latitude and  $a$ ,  $b$  and  $c$  are parameters estimated by regression with values  $a$   
152  $= 0.0819$ ,  $b = 0.0983$  and  $c = 7.7521$  (Cornwell *et al.*, 2017), and

$$153 \quad \delta^{13}\text{C}_{\text{air}} = \delta^{13}\text{C}_{\text{air},1992} + g(y - 1992) \quad (2)$$

154 where  $y$  is the sampling year and  $g = -0.0467$ , and

$$155 \quad \chi = (\delta^{13}\text{C}_{\text{air}} - \delta^{13}\text{C}_{\text{plant}} - a') / (b' - a') \quad (3)$$

156 where  $a'$  is the discrimination against <sup>13</sup>CO<sub>2</sub> during diffusion through stomata (4.4‰)  
157 and  $b'$  is the discrimination against <sup>13</sup>CO<sub>2</sub> during carboxylation (27‰) (Farquhar *et al.*,  
158 1982). Cernusak *et al.* (2013) showed that about 80% of the variation in instantaneous

159 gas exchange measurements of  $\chi$  could be accounted for by a linear relationship to  $\delta^{13}\text{C}$ ,  
160 supporting the use of equation (3). Estimates of  $\chi$  based on  $\delta^{13}\text{C}$  measurements are used  
161 here, however, because they reflect longer-term growth conditions better.

162 Three bioclimate variables adequately represent the controls on vegetation structure  
163 and composition across China (Wang *et al.*, 2013). These are the accumulated  
164 photosynthetically active radiation during the thermal growing season ( $\text{PAR}_0$ ), defined  
165 as the period when daily temperature is above  $0^\circ\text{C}$ ; the daily mean temperature during  
166 the thermal growing season ( $\text{mGDD}_0$ ); and the ratio of mean annual precipitation to  
167 annual equilibrium evapotranspiration (moisture index, MI), calculated using SPLASH  
168 (Davis *et al.*, 2017). The primary data for the calculation of these bioclimatic variables  
169 were derived from 1814 meteorological stations (740 stations with data from 1971 to  
170 2000, the rest from 1981 to 1990), interpolated to 1 km resolution with elevation as a  
171 covariate using ANUSPLIN V4.37 (Hutchinson 2007).

## 172 **Gap filling**

173 Photosynthetic measurements were only available for 14 sites in the China Plant Trait  
174 Database; however, these sites comprise 53% of the species represented in the data set.  
175 Photosynthetic measurements were not available for the temperate forests of  
176 Changbai Mountain, and the Inner Mongolia grasslands. In order to allow multivariate  
177 analysis of a larger data set,  $V_{\text{cmax}}$  values for species at these sites were gap-filled  
178 using a back-propagation neural network using LMA,  $N_{\text{area}}$ , LA,  $\chi$  and moisture index  
179 (MI) as predictors (`newff` function in Matlab 2010a). The neural network is a  
180 machine learning technique that often provides better performance than conventional  
181 statistical methods for this type of application (Paruelo *et al.*, 1997; Papale *et al.*, 2003;  
182 Moffat *et al.*, 2010). The data were divided into two parts: a calibration data set used  
183 to determine the weights in the neural network (75% of data points), and a validation  
184 data set used to assess the network performance (25% of data points). The method  
185 achieved an acceptable accuracy with  $R^2 = 0.49$  between observed and predicted

186 values for the calibration data set and 0.50 for the validation data set.  $J_{\max}$  values were  
187 then estimated from  $V_{\max}$  values using a linear regression fitted to data from all sites  
188 where both  $A_{\text{sat}}$  and  $A_{\max}$  were measured. The regression equation used for gap-filling  
189 is  $\ln J_{\max,25} = -0.0221 \text{ mGDD}_0 + 0.7329 \ln V_{\max,25} + 2.0362$  ( $R^2 = 0.75$ ,  $P < 0.01$ ).

## 190 **Multivariate analysis and variance partitioning**

191 *Principal components analysis* (PCA) and *redundancy analysis* (RDA) are powerful  
192 multivariate analysis techniques with many ecological applications (White *et al.*, 2005;  
193 Maire *et al.*, 2015; Scheibe *et al.*, 2015). As a dimensionality reduction technique,  
194 PCA projects a set of data on correlated variables on to a series of composite,  
195 uncorrelated variables called principal components (James *et al.*, 1990). In RDA,  
196 these variables are chosen to maximize the extent of their correlation with a set of  
197 predictor variables (Borcard *et al.*, 1992) and are therefore described as “constrained”  
198 axes of variation. RDA also extracts further “unconstrained” axes, which are the  
199 principal components of the variation that remains after the fitted effects of the  
200 predictor variables have been removed. Here, PCA is used to analyse trait covariation;  
201 RDA is used to analyse the relationships of trait variation to climate variables; and the  
202 unconstrained axes of RDA are used to characterize the residual (within-site) variation  
203 in traits. These analyses were performed using the *vegan* package in R (Oksanen *et al.*,  
204 2017). LA was square-root transformed before analysis to yield a linear measure of  
205 leaf size.  $\chi$  was logit-transformed ( $\text{logit } \chi = \ln [\chi/(1 - \chi)]$ ). All other traits (including  
206  $\sqrt{\text{LA}}$ ) were natural log-transformed. All traits were thus converted to dimensionless  
207 quantities in the range  $(-\infty, \infty)$ , allowing PCA and RDA to be carried out using the  
208 covariance matrix among traits with no need for further standardization. Each trait  
209 thereby has its ‘natural’ weight in the analysis. For log-transformed variables, this  
210 treatment implies that a trait with, say, 10-fold variation has twice the weight of a trait  
211 with 5-fold variation. The weight can be quantified by the standard deviation of the  
212 transformed variables ( $\ln \sqrt{\text{LA}}$ : 1.17,  $\ln \text{SLA}$ : 0.50,  $\ln \text{LDMC}$ : 0.38,  $\ln N_{\text{area}}$ : 0.59,  $\ln$

213  $V_{\text{cmax}25}$ : 0.58,  $\ln J_{\text{max}25}$ : 0.48, logit  $\chi$ : 1.37; see also Table 3). PCA and RDA were  
214 repeated using only the species-site combinations for which actual (as opposed to  
215 gap-filled) photosynthetic trait data were available (Figs S2-S4, Tables S4-S5).

216 *Variation partitioning* quantifies the amount of variation in a predicted quantity (in  
217 multiple regression) or set of quantities (in RDA) that can be explained by different  
218 groups of predictors (Legendre & Legendre, 2012). We used the Legendre method  
219 (Legendre & Anderson, 1999; Peres-Neto *et al.*, 2006; Meng *et al.*, 2015), which  
220 explicitly accounts for correlations between groups by distinguishing unique and  
221 overlapping contributions from each group. The results are most conveniently  
222 displayed as Venn diagrams. The method was used here with RDA to assign trait  
223 variation to components linked to climate, sites, life forms, families, and the  
224 intersections of these controls.

## 225 **Trait prediction**

226 We evaluated the predictive power of the fitted trait-climate relationships in the RDA  
227 analysis, first on the data set as a whole and then using a cross-validation approach  
228 (Picard & Cook, 1984; Kohavi 1995). We performed five iterations, in which 80% of  
229 the data was used for training and 20% retained for validation. The average  
230 root-mean-squared error (RMSE) across all five trials provides the final measure of  
231 goodness-of-fit.

232 The general predictive power of the trait-climate relationships was then tested using  
233 four independent global trait data sets: leaf economics traits (SLA, LDMC,  $N_{\text{area}}$ ) from  
234 Wright *et al.* (2004);  $\sqrt{\text{LA}}$  from Wright *et al.* (2017); photosynthetic traits ( $V_{\text{cmax}25}$ ,  
235  $J_{\text{max}25}$ ) from De Kauwe *et al.* (2016), including data from Bahar *et al.* (2017); and  $\chi$   
236 from Cornwell *et al.* (2017) (Table S6). Each of these data sets provides geolocated  
237 site-based measurements across continents, vegetation types and climates (Figure S5).  
238 We derived climate variables for each site from the nearest 10-minute grid cell in the

239 CRU 2.0 dataset (New *et al.* 2002), which provides long-term monthly means of  
240 temperature, precipitation, and sunshine duration for the standard period 1961-1990.  
241 PAR<sub>0</sub>, mGDD<sub>0</sub>, and MI were calculated in the same way as for the sites in China, using  
242 SPLASH to calculate MI (Davis *et al.*, 2017).

243 We screened out measurements from sites in the global data sets where MI > 1.4 or  
244 mGDD<sub>0</sub> < 10 because these are beyond the limits of the climates sampled in China.  
245 Some of the δ<sup>13</sup>C measurements in Cornwell *et al.* (2017) are < -30%. We assume that  
246 these reflect incomplete mixing of CO<sub>2</sub> between the free atmosphere and the forest  
247 understorey. We excluded these measurements. The number of sites and individual  
248 measurements from each global data set used to test the climate-trait predictions is  
249 shown in Table S6. Trait values at each global site were directly predicted from climate  
250 inputs, using the RDA model previously derived from the data in China. Ordinary  
251 least-squares regression was used to compare observed (y) with predicted (x) trait  
252 values.

## 253 **Results**

### 254 **Four dimensions of trait variation**

255 PCA of traits from all species and sampling sites revealed four independent axes of trait  
256 variation (Fig. 2, Table 1). The first four principal components together account for 95%  
257 of total trait variation. The first two axes are dominated by LA and  $\chi$ , orthogonal to one  
258 another. These two axes together account for 79% of total trait variation: this large  
259 fraction draws attention to the large span of variability in these traits, especially leaf  
260 area. The third axis, accounting for 11% of total trait variation, primarily represents the  
261 LES, with SLA opposed to  $N_{\text{area}}$  and LDMC. The plot of axis 3 against axis 4, which  
262 accounts for 6% of total trait variation, shows that  $V_{\text{cmax}}$  and  $J_{\text{max}}$  vary closely together,  
263 but orthogonally to the LES.

264 Analysis based on sites with complete data only (Fig. S2, Table S4) shows that the first four  
265 principal components have similar explanatory power to the main analysis (93%) and,  
266 although the axes are rotated with respect to the axes derived from the larger data set, they  
267 show the same four dimensions of variation with LA, LES, photosynthetic capacity and  $\chi$   
268 varying independently of one another. The patterns of trait covariation can also be seen  
269 by examining the matrix of pairwise correlations between traits (Fig. S6). The  
270 differences between Fig. S6(a) based on the gap-filled data set, and Fig. S6(b) based  
271 on sites with complete data, show the (slight) effect of gap-filling.  $V_{\text{cmax}}$  and  $J_{\text{max}}$  are  
272 highly correlated (0.84) before gap filling. The largest difference is that the negative  
273 correlations of both  $V_{\text{cmax}}$  and  $J_{\text{max}}$  with leaf area *increase* due to the gap filling. This  
274 evidently does not contradict our inference from PCA on the gap-filled data set, i.e.  
275 that photosynthetic capacities are largely uncorrelated with the other traits.

#### 276 **Trait variation related to climate**

277 The three bioclimatic variables together account for 37% of trait variation (Table 2).  
278 Three successive RDA axes (Fig. 3, Table 2) describe the patterns of trait variation  
279 with climate, and show that the between-site patterns of trait covariation imposed by  
280 climatic gradients differ from those found in the data set as a whole. The first RDA  
281 axis is overwhelmingly dominant, and is related to the gradient of MI from  
282 desert-steppe to moist forests. LA and  $\chi$  vary together along this gradient, with both  
283 large leaves and large  $\chi$  characteristic of wetter environments. The second RDA axis  
284 accounts for 2% of trait variation, and is related to the covariation of mean  
285 growing-season temperature and total growing-season light availability along the  
286 latitudinal gradient from the boreal zone to the tropics. Trait variation on this axis  
287 resembles the LES: warmer, higher irradiance climates are characterized by plants  
288 with lower SLA, higher LDMC and higher  $N_{\text{area}}$ . The third RDA axis accounts for  
289 only 0.4% of trait variation. Analysis based on sites with complete data only (Fig. S3,  
290 Table S5) shows the same patterns.

## 291 **Residual trait variation, unrelated to climate**

292 The unconstrained axes (or residual principal components) calculated by RDA after  
293 climatic differences among sites have been accounted for (Fig. 4, Table 2) provide  
294 insight into trait variation that is expressed within sites and across all climates. The  
295 patterns of this residual variation, as shown by the first four unconstrained axes, are  
296 similar to the patterns shown by the principal components of the whole data set (Fig. 2,  
297 Table 1), with evidence for four independent dimensions of variation associated with  
298 successive components dominated by  $\chi$ , LA, LES traits and photosynthetic capacities,  
299 respectively. Analysis based on sites with complete data only (Fig. S4, Table S5)  
300 shows the same four dimensions.

301 The same general patterns of non-climate-related trait covariation are also clear on  
302 inspection of the partial correlations among transformed trait values, after the effects  
303 of climatic predictors have been removed (Fig. 5). Deeper colours in Fig. 5 indicate  
304 larger absolute magnitudes of correlation. The traits can be seen to fall into four  
305 blocks: one comprising  $V_{\text{cmax}}$  and  $J_{\text{max}}$  (positively correlated), one comprising the  
306 traits that contribute to the LES (SLA negatively correlated with LDMC and  $N_{\text{area}}$ ),  $\chi$ ,  
307 and LA. While  $\chi$  shows almost no correlation with any of the other traits, LA is  
308 weakly negatively correlated with  $V_{\text{cmax}}$  and  $J_{\text{max}}$  (Fig. 5), as is SLA.

## 309 **Multiple controls of trait variation**

310 Venn diagrams (Fig. 6) summarize the percentage contributions of climate, site, life  
311 form and family (including intersecting contributions) to total trait variation, and to  
312 variation in each separate trait. The intersection regions represent trait variation that  
313 cannot be unambiguously attributed to one control or another, because of correlations  
314 among the controls. For example, substantial intersections between climate and family  
315 occur because these controls are not independent: different families are selected for in  
316 different climates. Anomalously large values are highlighted in bold in Fig. 6 and one

317 anomalously small value indicated by italics. No values are shown for climate  
318 independently of site, because differences in climate are determined by site locations.  
319 Table 3 also shows the total percentage of variance associated with each control  
320 (including intersections with other controls).

321 Considering the variation among all traits together (Fig. 6), climate, site, family and  
322 life form jointly account for 70% of total trait variance. The most important features  
323 of the partitioning are (1) the joint effect of climate with family (23%), which is the  
324 dominant driver of trait variation in this dataset; (2) the substantial fraction of  
325 variance due to family alone (17%), independent of climate or life form; and (3) the  
326 fact that most of the total variance associated with life form (16%) is also linked to  
327 climate (8%). There is some additional effect of climate independent of family (8%);  
328 and some effect of site independent of climate (12%), which is presumably related to  
329 edaphic or microclimatic factors.

330 The partitioning of trait variance for individual traits (Fig. 6) generally resembles that  
331 for all traits. However, 48% of total trait variation in LDMC is linked to family, and  
332 41% linked to family independent of other controls. Only 4% of the variation in  
333 LDMC is linked to climate, and none to climate and family together. For SLA, 41% of  
334 total trait variation is linked to family (with 14% linked to family and life form  
335 together independent of other controls); 15% is linked to climate, but only 4% to  
336 climate and family together. These anomalies indicate a particularly strong  
337 phylogenetic component to variation in LDMC and, to a lesser extent, SLA. The  
338 unexplained variation is greater for  $V_{\text{cmax}25}$  (47%) and  $J_{\text{max}25}$  (41%) than for the other  
339 traits.

340 After climate, site and family effects have been accounted for, the remaining  
341 (independent) contribution of life form to trait variation is small. The total life-form  
342 contribution is < 10% for all traits except LA and  $\chi$ , and the unique contribution of life  
343 form independent of all other controls is very slight, < 2.5% for all traits. Forbs and

344 graminoids show different ranges of trait values in forest and non-forest vegetation  
345 (Fig. S1). Specifically, SLA and LDMC of forbs and graminoids decrease between  
346 forests and non-forests while  $N_{\text{area}}$ ,  $V_{\text{cmax}}$  and  $J_{\text{max}}$  increase. That is, for all these traits,  
347 life forms occupying the understorey in forest vegetation become more ‘tree-like’ in  
348 non-forest vegetation, suggesting that these traits are more determined by the light  
349 environment than by any intrinsic difference among life forms.

### 350 **Worldwide prediction of traits based on the observed climate-trait relationships**

351 The RDA analyses show that climate (including indirect effects mediated by selection  
352 for life forms and families) is the major determinant of trait variation for most of the  
353 traits examined, except for LDMC and SLA, which show a substantial independent  
354 phylogenetic component. This generalization is supported by predictions of the mean  
355 site values for each trait (Fig S7). At species level, the adjusted  $R^2$  between observed  
356 and predicted values for LDMC is only 0.08, and for SLA 0.16 (Table S7), while the  
357 relationship is better for other traits – from 0.24 for  $V_{\text{cmax}25}$  to 0.52 for  $\sqrt{\text{LA}}$ . The  
358 average adjusted  $R^2$  across traits is 0.28. Partitioning the data into woody and  
359 non-woody components has little impact on the quality of the prediction for most traits,  
360 but prediction of LDMC and SLA is better for non-woody than woody species (Table  
361 S7). Although predictability is imperfect, because of the (demonstrated) influence of  
362 non-climatic factors on all of the traits, these analyses nonetheless show that it is  
363 possible to predict all four dimensions of trait variation, to first order, from climate.

364 The prediction of trait values in global data sets provides a more stringent test of the  
365 universality of the derived climate-trait relationships (Fig. 7, Table 4). At site level,  
366 the lowest adjusted  $R^2$  value between observed and predicted trait values is again for  
367 LDMC (0.01), but for SLA it is 0.31. For other traits, adjusted  $R^2$  ranged from 0.25  
368 ( $J_{\text{max}}$ ) to 0.34 ( $\sqrt{\text{LA}}$ ). The average across traits is 0.31, excluding LDMC. The  
369 observed values for  $\ln V_{\text{cmax}25}$  tend to be higher than the predicted values, whereas the  
370 observed values of  $\ln \text{SLA}$  tend to be lower than the predicted values (Fig. 7).

371 However the regression slopes for these traits are not significantly different from  
372 unity (Table 4). The OLS regression slopes for  $\ln \sqrt{LA}$ ,  $J_{\max 25}$  and  $\ln \chi$  are in the range  
373 from 0.48 to 1. RMSE values (Table 4) are larger in the global comparison than in the  
374 calibration set for  $\ln \sqrt{LA}$  and SLA; but closely similar for  $N_{\text{area}}$ ,  $V_{\text{cmax}25}$  and  $J_{\max 25}$ , and  
375  $\chi$ . The average RMSE across traits excluding LDMC is slightly less in the global  
376 comparison (0.42) than in the calibration set (0.61).

## 377 **Discussion**

### 378 **The ecological significance of leaf-trait dimensions**

379 The four dimensions of total leaf-trait variation reported here indicate the existence of  
380 independent variation among species in LA,  $\chi$ , photosynthetic capacity, and the LES.  
381 The RDA based on climate shows a smaller dimensionality, with most of the variation  
382 concentrated on a single axis from wet to dry environments. LA is both expected and  
383 observed to increase with plant-available moisture, due to energy-balance constraints  
384 (Wright *et al.*, 2017).  $\chi$  is both expected and observed to increase with atmospheric  
385 moisture according to the least-cost hypothesis (Prentice *et al.*, 2014). These  
386 hydroclimatic controls on both LA and  $\chi$  are presumed to be the cause of (a) the  
387 dominance of a single dimension of trait-environment relationships across the region,  
388 related to moisture/aridity, and (b) the observed close covariation of LA and  $\chi$   
389 between sites along the aridity gradient – contrasting with their independence in the  
390 data as a whole. Analysis of the residual (non-climatic) component of trait variation  
391 however shows, once again, four independent dimensions, with a pattern closely  
392 similar to that shown in total leaf-trait variation, and orthogonal variation of LA and  $\chi$ .

393 Multivariate analysis confirms the universal nature of the LES, as indexed here by  
394 SLA, LDMC (which tends to be high when SLA is low), and  $N_{\text{area}}$ . Unlike  $N_{\text{mass}}$  (N  
395 concentration per unit mass),  $N_{\text{area}}$  increases with *decreasing* SLA because the  
396 structural component of leaf N increases in proportion to LMA (see e.g. Onoda *et al.*,

397 2004, 2017; Wright *et al.*, 2005; Osnas *et al.*, 2013; Dong *et al.*, 2017a). The LES is  
398 identified in the PCA, and in the residual trait variation after consideration of climate  
399 effects in RDA. However, it also appears in the climatically constrained RDA as a  
400 second-order pattern correlated with the latitudinal gradient. In other words, there is a  
401 shift in the average position of species along the LES (towards lower SLA) with  
402 increasing growing-season length and warmth, although this shift accounts only for a  
403 small proportion (2%) of total trait variance. The LES reflects the inescapable linkage  
404 between high construction costs and long payback times of leaves with low SLA  
405 (Kikuzawa, 1991; Reich *et al.*, 1997; McMurtrie & Dewar, 2011; Funk & Cornwell,  
406 2013). The shift towards lower-SLA leaves in warmer climates is primarily due to the  
407 shift of dominance from deciduous to evergreen woody plants. The increase in  
408 growing-season length (towards a year-round growing season in the tropics) favours  
409 longer-lived evergreen leaves with lower SLA in warmer climates, as shown here and  
410 in other studies.

411 Both the gap-filled data set and the non-gap-filled subset show that the two  
412 photosynthetic capacities ( $V_{\text{cmax}}$  and  $J_{\text{max}}$ ) covary closely (Fig. S6), as is expected  
413 from the co-ordination hypothesis – which predicts that leaves should not possess  
414 excess capacity in either carboxylation or electron transport, as photosynthesis  
415 depends on both (Chen *et al.*, 1993; Maire *et al.*, 2012). However both traits show  
416 substantial variation within sites. When  $V_{\text{cmax}}$  and  $J_{\text{max}}$  were entered into the analysis  
417 after adjustment to local growth temperature, as opposed to 25°C, the results were  
418 very similar (not shown). Opposite trends of variation in  $V_{\text{cmax}}$  and  $J_{\text{max}}$  are shown  
419 only in the (minor) third axis of the RDA, accounting for 0.4% of total trait variance  
420 and driven by differences among sites in summer temperature that are independent of  
421 the latitudinal gradient. This pattern is consistent with expectations, as a decline in the  
422  $J_{\text{max}}:V_{\text{cmax}}$  ratio with increasing temperature has been shown experimentally (Kattge &  
423 Knorr, 2007) and predicted theoretically (Wang *et al.*, 2017a). The decline is larger

424 when the two photosynthetic capacities are estimated at prevailing growth  
425 temperature, but persists when they are adjusted to 25°C.

#### 426 **Contributions to leaf trait variation**

427 The variance partitioning results presented here demonstrate that family and climate  
428 effects (except for LDMC and SLA) overlap considerably. In other words, a  
429 substantial part of trait variation with climate is due to families replacing one another  
430 along environmental gradients. After family, climate and site effects have been taken  
431 into account, independent life-form effects become unimportant. Thus, to first order,  
432 the principal controls on trait variation in this data set are family identity, climate, and  
433 climatic selection among families. Additional effects of site (independent of climate)  
434 could in principle be due to microclimatic and/or edaphic differences among sites,  
435 which have not been investigated. LDMC and to a lesser extent SLA show stronger  
436 family effects than other traits, while the effects of climate on these traits appear to be  
437 largely independent of family identity.

#### 438 **Implications for vegetation modelling**

439 Vegetation models based on continuous variation in trait space sample ‘plants’ from a  
440 continuum of trait values (e.g. Scheiter *et al.*, 2013; Fyllas *et al.*, 2014). This approach  
441 requires specifying which traits can vary; by how much; and the extent to which  
442 different traits covary, in other words, the effective dimensionality of trait space. Our  
443 analyses of leaf traits, including traits derived from stable isotope and gas exchange  
444 measurements, indicate that at least four independent dimensions of trait variation  
445 need to be considered; that realistic modelling of functional diversity must allow for  
446 within-site variation in each of these dimensions; and that environmental differences  
447 force patterns of trait covariation across sites that can be different from patterns  
448 observed within sites.

449 With the exception of LDMC, which shows a particularly strong phylogenetic  
450 component, the trait-environment relationships found here should be amenable to  
451 process-based modelling. The energy balance implications of leaf size (Michaletz *et*  
452 *al.*, 2016; Dong *et al.*, 2017b; Wright *et al.*, 2017) mean that this trait is crucial for  
453 survival, particularly in cold climates or in hot, dry climates. As the biophysical  
454 controls of leaf size are relatively well understood, it should be straightforward to  
455 build energy-balance constraints on leaf size into trait-based models. Shifts in the LES  
456 along environmental gradients could also be modelled, given the well-established  
457 relationship of leaf longevity and SLA (Wright *et al.*, 2004) and the experimentally  
458 determined variations of SLA with environmental factors (Poorter *et al.*, 2009). The  
459 distribution of SLA within communities could be represented by a pattern of  
460 covariation in leaf longevity, SLA, LDMC and the structural component of  $N_{\text{area}}$ , as  
461 shown here and in other studies.

462

463 The co-ordination hypothesis predicts both  $V_{\text{cmax}}$  and the ratio of  $J_{\text{max}}$  to  $V_{\text{cmax}}$ ,  
464 including the observed dependence of both quantities on growth temperature (Wang *et*  
465 *al.*, 2017b). Large-scale patterns in  $V_{\text{cmax}}$  and the metabolic component of  $N_{\text{area}}$  can be  
466 predicted theoretically (Dong *et al.*, 2017a). The co-ordination hypothesis also  
467 predicts the observed seasonal acclimation of  $V_{\text{cmax}}$  and  $J_{\text{max}}$  (Togashi *et al.*, 2018).  
468 Thus, at the level of community mean values, it seems likely that  $V_{\text{cmax}}$  can be  
469 successfully modelled as a function of environment (Ali *et al.*, 2016). A  
470 temperature-dependent ratio of  $J_{\text{max}}$  to  $V_{\text{cmax}}$  would then allow prediction of  $J_{\text{max}}$ .

471

472 The  $\text{CO}_2$  drawdown from air to leaf, indexed by  $\chi$ , is predicted by most vegetation  
473 models by simultaneous solution of the FvCB equations to predict assimilation rate as  
474 a function of leaf-internal  $\text{CO}_2$  ( $c_i$ ) and the diffusion equation to predict  $c_i$  as a  
475 function of ambient  $\text{CO}_2$  ( $c_a$ ), stomatal conductance and assimilation rate (Farquhar *et*  
476 *al.*, 1980). Theoretically and empirically well-founded relationships between  $\chi$  and

477 environmental variables (Wang *et al.*, 2017b) provide an alternative way to model  $\chi$   
478 directly as a function of environment, and thus to predict assimilation rates more  
479 straightforwardly than in many current models.

#### 480 **Challenges and future directions**

481 This analysis illustrates the power of large trait data sets spanning a large range of  
482 climates, and including measurements from multiple co-existing species at each field  
483 site, to reveal general patterns. It also shows the utility of multivariate analysis to  
484 summarize patterns, and variance partitioning to attribute trait variability to different  
485 (and sometimes intersecting) causes. But despite the availability of large plant-trait  
486 data compilations (e.g. Kattge *et al.*, 2011), the number of sites that include all of any  
487 specified set of plant traits is often disappointingly small – because different research  
488 groups typically collect data on different sets of traits. There remains a need for more  
489 extensive trait data collection including photosynthetic traits and isotopic  
490 measurements in addition to conventional leaf traits, and for such data collection to  
491 extend to the full range of the world’s climates. There has been a limited amount of  
492 comparative work, for example, on photosynthetic traits, which are essential for all  
493 process-based vegetation modelling. Moreover, compared to leaf traits, there is a  
494 paucity of data on other field-measurable traits (notably stem hydraulic properties)  
495 that may be equally important for plant functional ecology. As is well illustrated by  
496 the global data sets that we used to test the predictive capacity of trait-climate  
497 relationships, the site- and/or species-metadata available are often limited. There  
498 remains a need for extensive, targeted collection and analysis of plant trait data,  
499 including co-located morphological, gas-exchange and isotopic measurements, and  
500 spanning the world’s major environmental and floristic gradients.

501 **Acknowledgments**

502 This research has been by supported by High-end Foreign Expert Programmes of  
503 China (GDW20156100290, GDW20166100147) (ICP and SPH), the National Natural  
504 Science Foundation of China (41701051, 31600388) (YY and HW), the National Basic  
505 Research Program of China (2013CB956600) (GL and CP), the Fundamental Research  
506 Funds for the Central Universities (YY), the QianRen Program, and the Natural  
507 Sciences and Engineering Research Council of Canada (NSERC) Discover Grant (CP).  
508 SPH acknowledges support from the ERC-funded project GC2.0 (Global Change 2.0:  
509 Unlocking the past for a clearer future, grant number 694481). This research  
510 contributes to the AXA Chair Programme in Biosphere and Climate Impacts and the  
511 Imperial College initiative on Grand Challenges in Ecosystems and the Environment  
512 (ICP). We thank O. Atkin, K. Crous, T. Domingues, D. Ellsworth, H. Togashi, Ü.  
513 Niinemets and L. Weerasinghe for providing the photosynthesis data ( $V_{cmax25}$ ,  $J_{max25}$ )  
514 used in the validation.

515 **Author contributions**

516 YY, HW, SPH and ICP collectively devised the analysis strategy and interpreted the  
517 results. YY carried out all of the statistical analyses and wrote the first draft of the  
518 manuscript. IJW provided additional advice on the analysis and interpretation of trait  
519 variation patterns. All authors provided input to the final draft.

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748

## 749 **Figure legends**

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751 shown as red dots superimposed on a simplified vegetation map of China in (a); these  
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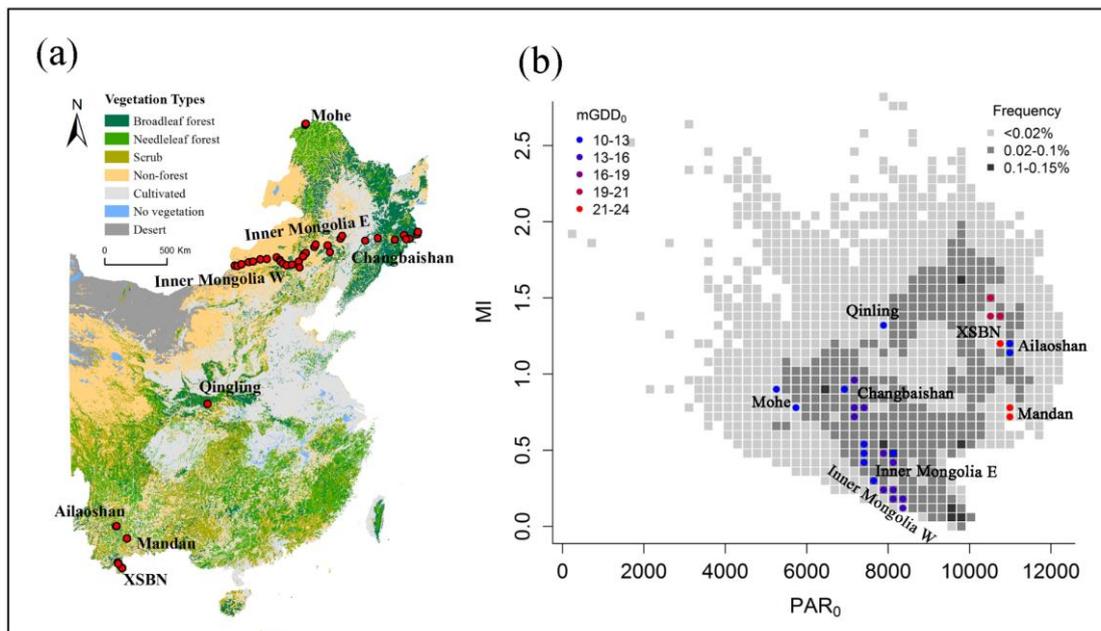
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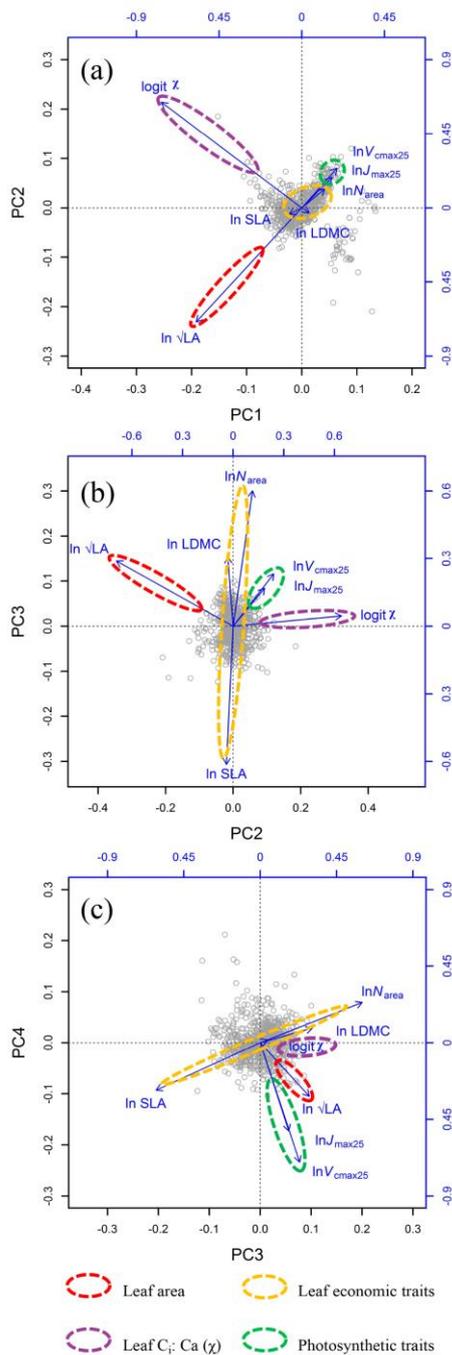
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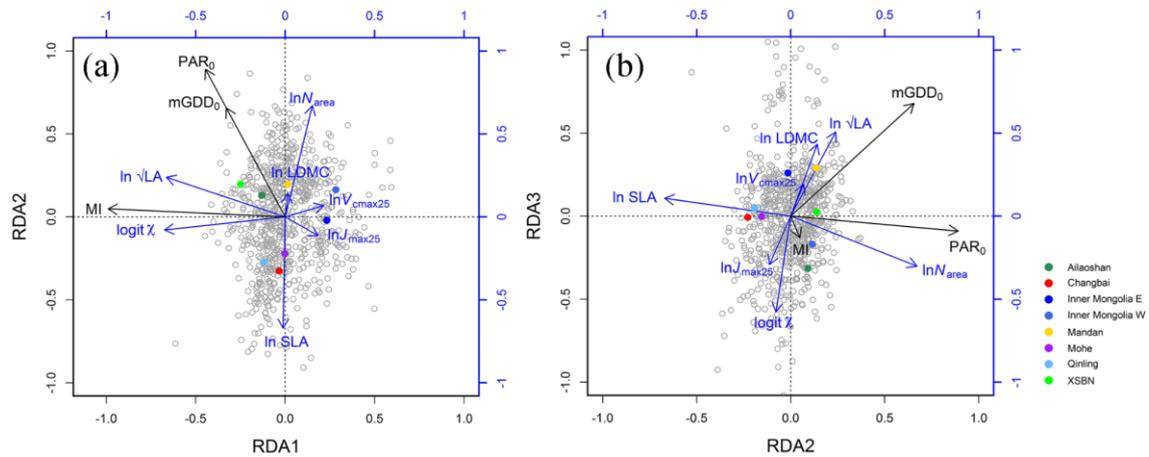


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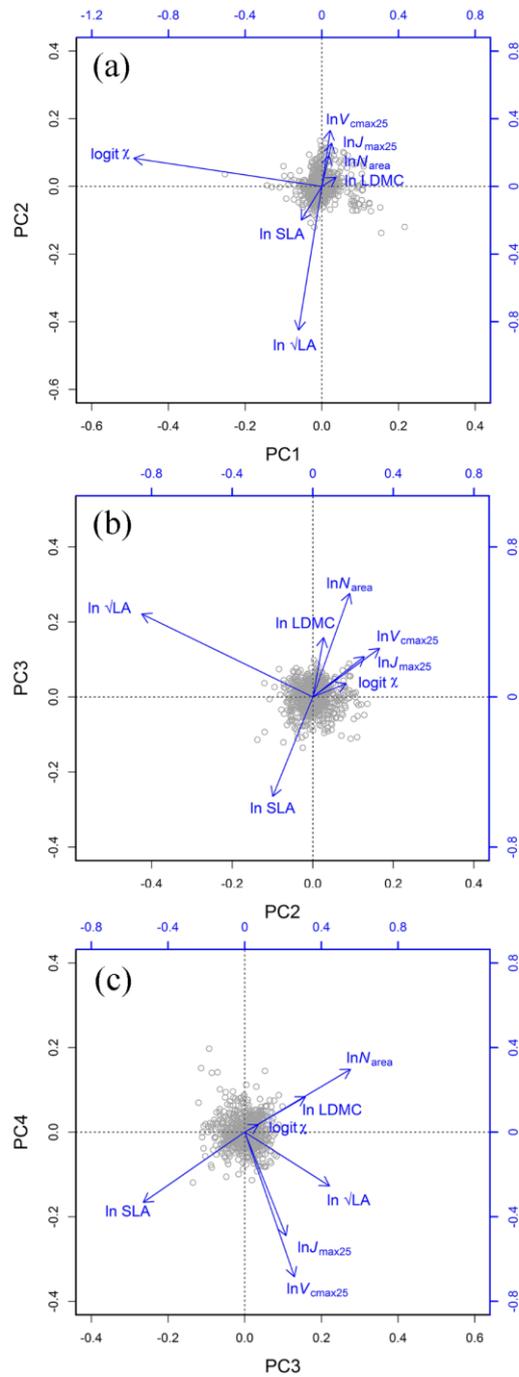
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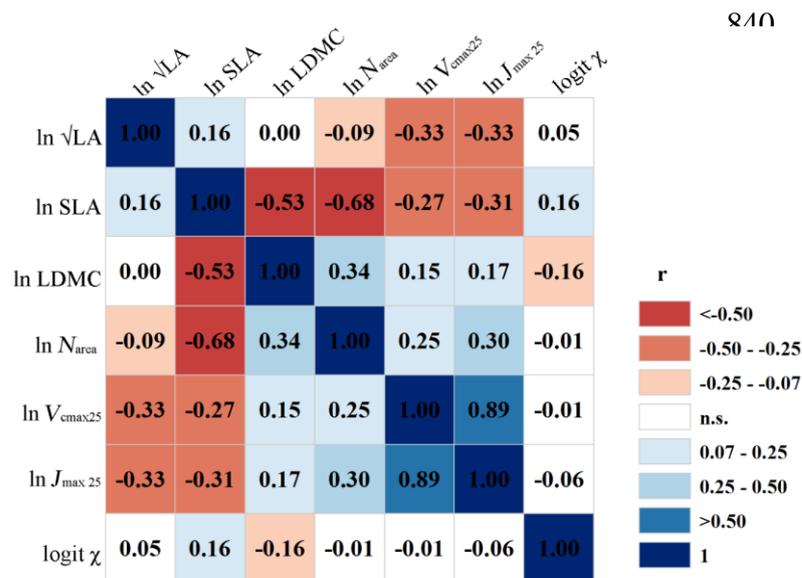
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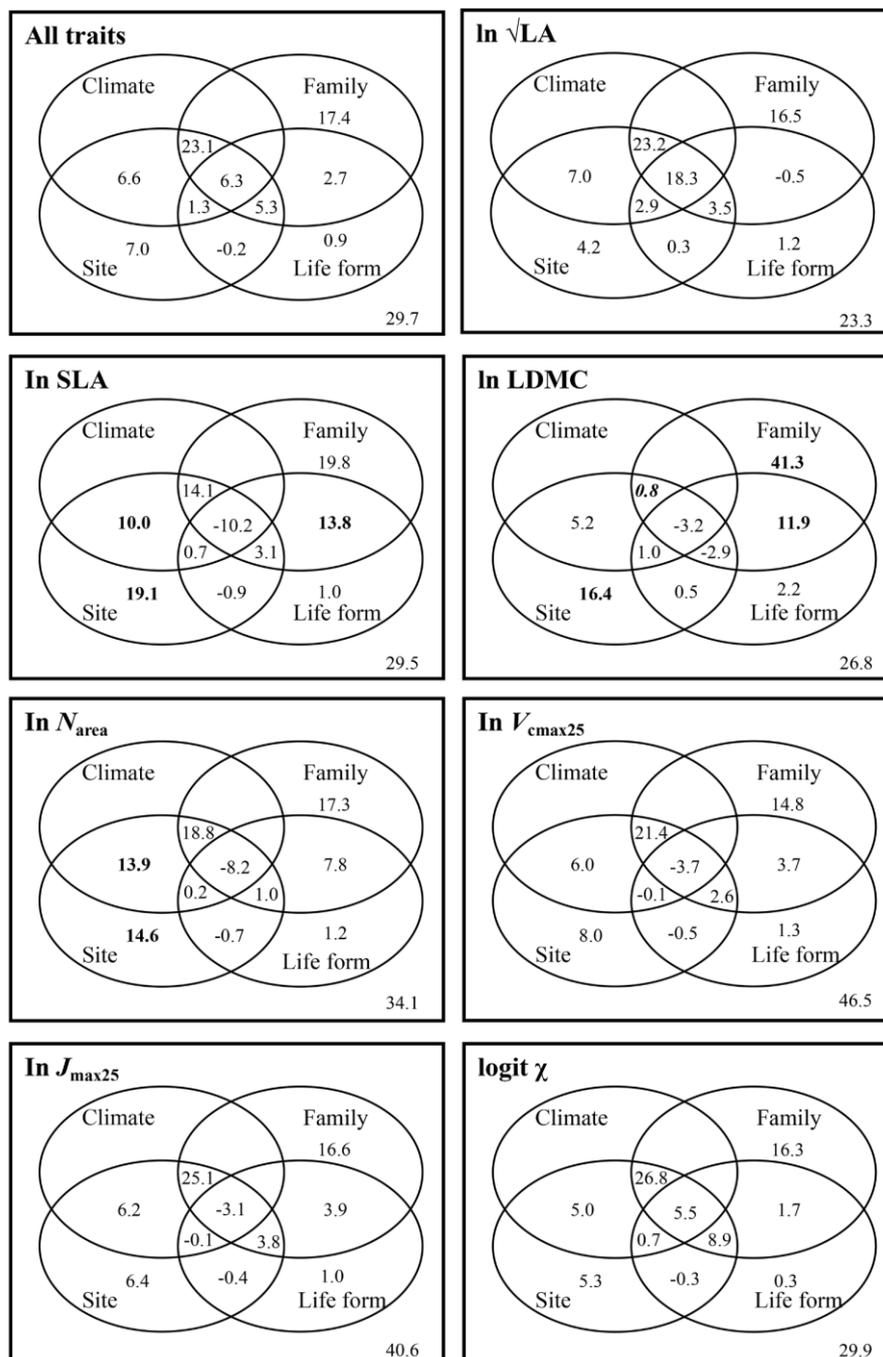
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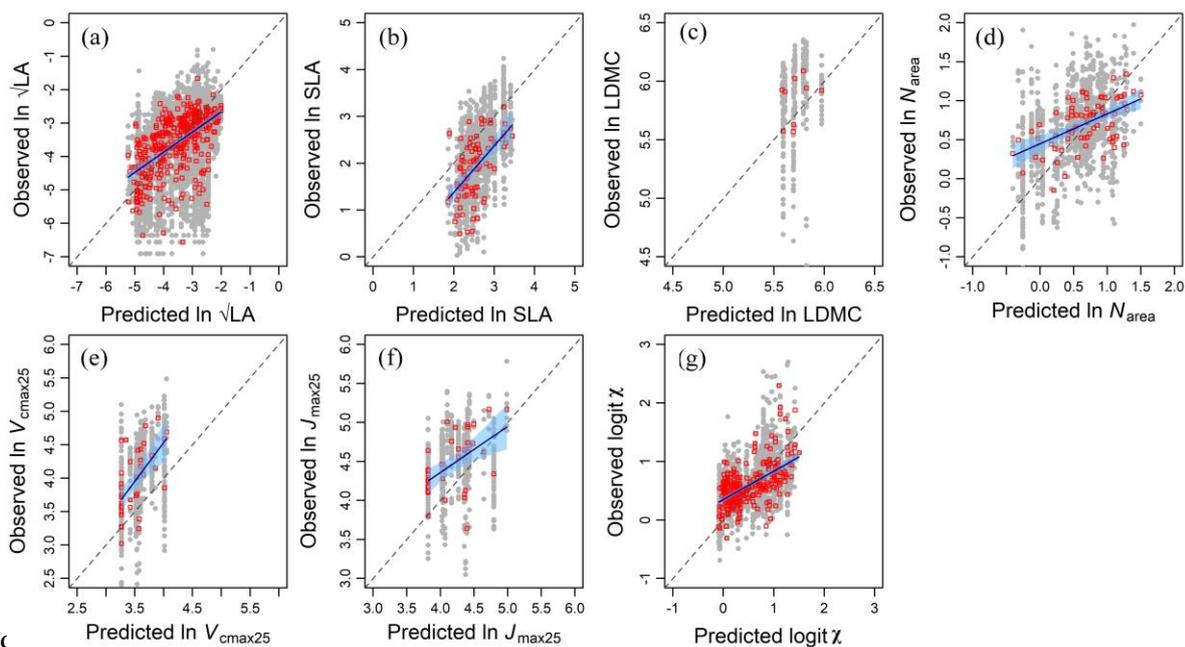
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860

861 Table 1 Trait loadings, eigenvalues, and the percentage of trait variation explained by  
 862 successive principal components in the trait PCA. Loadings > 0.3 in magnitude are  
 863 shown in **bold**.

	PC1	PC2	PC3	PC4
$\ln \sqrt{LA}$	<b>-0.57</b>	<b>-0.69</b>	0.29	<b>-0.31</b>
$\ln SLA$	-0.07	-0.04	<b>-0.61</b>	-0.28
$\ln LDMC$	0.04	-0.03	<b>0.31</b>	0.09
$\ln N_{area}$	0.12	0.11	<b>0.60</b>	0.24
$\ln V_{cmax,25}$	0.19	0.24	0.23	<b>-0.70</b>
$\ln J_{max,25}$	0.16	0.19	0.17	<b>-0.52</b>
$\text{logit } \chi$	<b>-0.76</b>	<b>0.64</b>	0.05	0.02
<b>Eigenvalue</b>	2.57	0.90	0.50	0.25
<b>Explained (%)</b>	58.0	20.4	11.3	5.6
<b>Cumulative (%)</b>	58.0	78.5	89.8	95.4

864

865 Table 2 Trait loadings, eigenvalues, and the percentage of trait variation explained by  
866 successive RDA axes (constrained by climate) and residual principal components, with  
867 axes 1 and 2 mirrored to facilitate comparison with the PCA. Loadings > 0.3 in  
868 magnitude are shown in **bold**.

	RDA1	RDA2	RDA3	PC1	PC2	PC3	PC4
$\ln \sqrt{LA}$	<b>-0.66</b>	0.24	<b>0.51</b>	0.12	<b>-0.85</b>	<b>-0.44</b>	0.25
$\ln SLA$	-0.01	<b>-0.67</b>	0.11	0.11	-0.20	<b>0.53</b>	<b>0.33</b>
$\ln LDMC$	0.02	0.14	<b>0.43</b>	-0.08	0.05	<b>-0.32</b>	-0.17
$\ln N_{area}$	0.15	<b>0.67</b>	<b>-0.30</b>	-0.04	0.18	<b>-0.55</b>	<b>-0.30</b>
$\ln V_{cmax,25}$	0.22	0.07	0.19	-0.04	<b>0.33</b>	-0.26	<b>0.68</b>
$\ln J_{max,25}$	0.18	-0.11	-0.29	-0.05	0.26	-0.22	<b>0.49</b>
$\text{logit } \chi$	<b>-0.67</b>	-0.08	<b>-0.58</b>	<b>0.98</b>	0.17	-0.07	-0.04
<b>Eigenvalue</b>	1.55	0.08	0.02	1.19	0.75	0.42	0.24
<b>Explained (%)</b>	34.9	1.8	0.4	26.8	17.0	9.6	5.3
<b>Cumulative (%)</b>	34.9	36.7	37.1	63.9	80.9	90.5	95.9

869

870 Table 3 Total contributions (%) of climate, family, site and life form to trait variation.

871 Standard deviations (weights) of the transformed variables are also given.

	All traits	$\ln \sqrt{LA}$	$\ln SLA$	$\ln LDMC$	$\ln N_{area}$	$\ln V_{cmax25}$	$\ln J_{max25}$	$\text{logit } \chi$
Weights		1.17	0.50	0.38	0.59	0.58	0.48	1.37
Climate	37.3	51.4	14.6	3.7	24.7	23.6	28.1	38.0
Family	54.8	61.0	40.5	48.0	36.7	38.8	46.3	59.0
Site	49.4	59.4	35.9	17.8	39.6	33.7	37.9	51.8
Life form	16.3	<b>25.8</b>	7.5	9.4	1.3	3.4	5.1	<b>16.7</b>

872

873 Table 4 Prediction accuracy of the trait-climate RDA model for independent global data  
874 sets at site level. \* indicates that the slope is significantly different from 1 ( $P < 0.01$ ), #  
875 indicates that the intercept is significantly different from 0 ( $P < 0.01$ ). \*\* indicates that  
876 the regression is significant ( $P < 0.01$ ).

Traits	Slope	Intercept	$R^2_{adj}$	$n$	RMSE	Source of data
ln $\sqrt{LA}$	0.60* (0.52, 0.70)	-1.45# (-1.72, -1.10)	0.34**	388	0.70	Wright et al. (2017)
ln SLA	0.99 (0.68, 1.31)	-0.61 (-1.41, 0.19)	0.31**	87	0.53	Wright et al. (2004)
ln LDMC	n.s.	n.s.	0.01	9	0.20	Wright et al. (2004)
ln $N_{area}$	0.38* (0.24, 0.52)	0.45# (0.34, 0.56)	0.28**	77	0.26	Wright et al. (2004)
ln $V_{cmax25}$	1.16 (0.62, 1.69)	-0.11 (-1.97, 1.76)	0.33**	38	0.40	De Kauwe et al. (2016)
ln $J_{max25}$	0.59* (0.27, 0.92)	1.99# (0.62, 3.36)	0.25**	38	0.33	De Kauwe et al. (2016)
logit $\chi$	0.48* (0.40, 0.57)	0.35# (0.30, 0.40)	0.33**	281	0.29	Cornwell et al. (2017)