

Empirical likelihood tests for nonparametric detection of differential expression from RNA-seq data

Article

Supplemental Material

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Supplementary material

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1 Empirical likelihood

For n i.i.d one dimensional observations x_1, \dots, x_n the empirical likelihood (Owen, 1988) can be defined as

$$f(x) = \prod_{i=1}^n p_i, \quad (1)$$

where we assign each observation a weight p_i , and constrain these such that $\sum_{i=1}^n p_i = 1$, $\forall i, 0 \leq p_i \leq 1$. Focussing on the empirical likelihood for the mean μ of our observations x_i , we simply require that

$$\sum_{i=1}^n p_i x_i = \mu. \quad (2)$$

Then we have three constraints, and aim to find the p_i that maximise the empirical likelihood $f(x)$ under these constraints. Fortunately by using Lagrange multipliers we can find the optimal p_i by solving a one dimensional root finding problem. Defining

$$G = \sum_{i=1}^n \log(np_i) - n\lambda \sum_{i=1}^n p_i(x_i - \mu) + \gamma \left(\sum_{i=1}^n p_i - 1 \right), \quad (3)$$

and taking the partial derivative with respect to p_i , applying the method of Lagrange multipliers (Owen, 2001) we have

$$\frac{\partial G}{\partial p_i} = \frac{1}{p_i} - n\lambda(x_i - \mu) + \gamma = 0, \quad (4)$$

and we can solve for γ by considering

$$\sum_{i=1}^n p_i \frac{\partial G}{\partial p_i} = 0 \quad (5)$$

$$\sum_{i=1}^n (1 - n\lambda p_i(x_i - \mu) + p_i \gamma) = 0 \quad (6)$$

$$n + \gamma = 0, \quad (7)$$

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since we know $\sum_{i=1}^n p_i(x_i - \mu) = 0$. Then substituting $\gamma = -n$ into equation 4 we have

$$\frac{1}{p_i} - n\lambda(x_i - \mu) - n = 0 \quad (8)$$

$$p_i = \frac{1}{n\lambda(x_i - \mu) + n}, \quad (9)$$

and so p_i depends only on solving equation 4 for λ . We know that

$$\sum_{i=1}^n p_i(x_i - \mu) = 0 \quad (10)$$

$$\sum_{i=1}^n \frac{(x_i - \mu)}{n\lambda(x_i - \mu) + n} = 0, \quad (11)$$

and so we can solve for λ for a given value of μ using a univariate root finding algorithm. Then using equation 9 we can find the p_i and calculate the empirical likelihood in equation 1.

1.1 Euclidean likelihood

The Euclidean likelihood (Baggerly, 1998) defines the log likelihood as

$$\log f(x|\mu) = -\frac{1}{2} \sum_{i=1}^n (np_i - 1)^2, \quad (12)$$

with the constraints $\sum_{i=1}^n p_i = 1$ and $\sum_{i=1}^n p_i x_i - \mu = 0$. Again we apply the method of Lagrange multipliers (Owen, 2001)

$$G = -\frac{1}{2} \sum_{i=1}^n (np_i - 1)^2 - n\lambda \sum_{i=1}^n p_i(x_i - \mu) + \gamma \left(\sum_{i=1}^n p_i - 1 \right), \quad (13)$$

and setting the partial derivative of G with respect to p_i to zero we have

$$\frac{\partial G}{\partial p_i} = n(1 - np_i) - n\lambda(x_i - \mu) + \gamma = 0 \quad (14)$$

$$\frac{1}{n} \sum_{i=1}^n (n(1 - np_i) - n\lambda(x_i - \mu) + \gamma) = 0 \quad (15)$$

$$-n\lambda(\bar{x} - \mu) + \gamma = 0. \quad (16)$$

Substituting $\gamma = n\lambda(\bar{x} - \mu)$ back into equation 14

$$n(1 - np_i) - n\lambda(x_i - \mu) + n\lambda(\bar{x} - \mu) = 0 \quad (17)$$

$$p_i = \frac{1}{n} (1 - \lambda(x_i - \bar{x})). \quad (18)$$

Given that $\sum_{i=1}^n p_i(x_i - \mu) = 0$, we can substitute equation 18 to give

$$\sum_{i=1}^n \frac{(x_i - \mu)}{n} (1 - \lambda(x_i - \bar{x})) = 0 \quad (19)$$

$$\bar{x} - \mu - \sum_{i=1}^n \frac{\lambda}{n} (x_i - \mu)(x_i - \bar{x}) = 0 \quad (20)$$

$$\bar{x} - \mu - \sum_{i=1}^n \frac{\lambda}{n} (x_i - \bar{x})(x_i - \bar{x}) = 0 \quad (21)$$

$$\bar{x} - \mu - \lambda s = 0, \quad (22)$$

where s is defined as $s = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})(x_i - \bar{x})$. Substituting λ into equation 18 we have

$$p_i = \frac{1}{n} \left(1 - \frac{1}{s} (\bar{x} - \mu)(x_i - \bar{x}) \right), \quad (23)$$

and substituting p_i into equation 12 we arrive at

$$\log f(x|\mu) = - \sum_{i=1}^n \left(\frac{1}{s} (\bar{x} - \mu)(x_i - \bar{x}) \right)^2 \quad (24)$$

$$= - \frac{1}{s^2} (\bar{x} - \mu)^2 \left(\sum_{i=1}^n (x_i - \bar{x})^2 \right) \quad (25)$$

$$= - \frac{1}{s} n (\bar{x} - \mu)^2, \quad (26)$$

References

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