

Quantifying leaf trait covariation and its controls across climates and biomes

Article

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2 climates and biomes

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- 4 Yanzheng Yang^{1,2,3,*}, Han Wang^{1,3}, Sandy P. Harrison^{3,4}, I. Colin
- 5 Prentice^{1,3,5,6}, Ian J. Wright⁶, Changhui Peng^{3,7,*} and Guanghui Lin^{1,8,*}

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- ¹Ministry of Education Key Laboratory for Earth System Modeling, Department of
- 8 Earth System Science, Tsinghua University, Beijing 100084, China.
- 9 ²Joint Center for Global Change Studies (JCGCS), Beijing 100875, China
- 10 ³Center for Ecological Forecasting and Global Change, College of Forestry,
- 11 Northwest A&F University, Yangling, Shaanxi 712100, China
- 12 ⁴School of Archaeology, Geography and Environmental Sciences (SAGES),
- 13 University of Reading, Reading, UK
- 14 ⁵AXA Chair of Biosphere and Climate Impacts, Imperial College London,
- 15 Department of Life Sciences, Silwood Park Campus, Buckhurst Road, Ascot SL5 7PY,
- 16 UK
- ⁶Department of Biological Sciences, Macquarie University, North Ryde, NSW 2109,
- 18 Australia
- 19 ⁷Department of Biological Sciences, Institute of Environmental Sciences, University
- 20 of Quebec at Montreal, C.P. 8888, Succ. Centre-Ville, Montréal H3C 3P8, QC,
- 21 Canada
- 22 ⁸Key Laboratory of Stable Isotope and Gulf Ecology, Graduate School at Shenzhen,
- Tsinghua University, Shenzhen, Guangdong 518055, China

- 25 Revised version for New Phytologist
- 26 (*Authors for correspondence: tel +86(10)62797230; email <u>yanzheng148@163.com</u>
- 27 (Y.Y.); tel +86(10)62797230; email <u>lingh@tsinghua.edu.cn</u> (G.L.); tel
- 28 +86(29)87080608; email <u>cpeng86@yahoo.com</u> (C.P.))

Summary

- Plant functional ecology requires the quantification of trait variation and its controls. Field measurements on 483 species at 48 sites across China were used to
- analyse variation in leaf traits, and assess their predictability.
- Principal components analysis (PCA) was used to characterize trait variation,
- redundancy analysis (RDA) to reveal climate effects, and RDA with variance
- partitioning to estimate separate and overlapping effects of site, climate, life-form
- and family membership.
- Four orthogonal dimensions of total trait variation were identified: leaf area (LA),
- internal-to-ambient CO_2 ratio (χ), leaf economics spectrum traits (specific leaf
- area (SLA) versus leaf dry matter content (LDMC) and nitrogen per area (N_{area}) ,
- and photosynthetic capacities ($V_{\rm cmax}$, $J_{\rm max}$ at 25°C). LA and χ covaried with
- 41 moisture index. Site, climate, life form and family together explained 70% of trait
- 42 variance. Families accounted for 17%, and climate and families together 29%
- LDMC and SLA showed the largest family effects. Independent life-form effects
- 44 were small.
- Climate influences trait variation in part by selection for different life forms and
- 46 families. Trait values derived from climate data via RDA showed substantial
- 47 predictive power for trait values in the available global data sets. Systematic trait
- data collection across all climates and biomes is still necessary.
- 49
- 50 **Key words:** climate, leaf economics spectrum, multivariate analysis, photosynthetic
- 51 capacity, phylogeny, plant functional traits.

Introduction

- 54 Functional traits generally do not vary independently, but show broadly predictable
- patterns of covariation (Armbruster et al., 1996; Watson et al., 2016). The covariation
- of traits may mean that traits share genetic controls, or that they have related roles in
- 57 community assembly and function (Wright et al., 2007; Fajardo et al., 2011).
- Quantifying the covariation of vegetative traits and their controls is important for an
- 59 understanding of how plants drive ecosystem processes and determine the responses
- of ecosystems to environmental change (Wright et al., 2007; Shipley et al., 2011;
- 61 Swenson 2013; van Bodegom et al., 2014; Kong et al., 2014; Kraft et al., 2015).
- 62 Although a number of large-scale studies have quantified both trait covariation (e.g.
- Wright et al., 2004; Armbruster et al., 2014; Peiman & Robinson, 2017) and
- trait-environment relationships, (e.g. Wright et al., 2005; Harrison et al., 2010; Liu et
- al., 2012; Maire et al., 2015; Meng et al., 2015), a number of general issues await
- 66 resolution. These include:
- 67 (1) The dimensionality of trait space that is, the extent to which combinations of
- 68 different traits are independent, versus belonging to a set of covarying traits as
- 69 exemplified by the leaf economics spectrum (LES) (Wright et al., 2004, 2005). The
- 70 intrinsic dimensionality of traits is the minimum number of independent axes that
- 71 adequately describe the functional variation among species, and is therefore an
- 72 important quantity in comparative ecology (Laughlin, 2014).
- 73 (2) The extent to which trait variation is determined by climate, versus the
- 74 co-existence of multiple trait values in the same climate (Adler et al., 2013;
- 75 Valladares *et al.*, 2015).
- 76 (3) The extent to which trait variation and trait-environment correlations are linked to
- 77 'hard-wired' physiognomic (life-form) and/or phylogenetic differences among species,
- and the role of environment in selecting among life forms and clades (Díaz et al.,
- 79 2013; Ackerly, 2009; Donovan et al., 2014).

The dimensionality question has received attention in plant functional ecology partly because of the universal nature of the LES, which is considered as the outcome of a tradeoff between resource acquisition and conservation - representing different general strategies for existence, rather than adaptations to environment (Wright et al., 2007; Kong et al., 2014; Reich, 2014). An early synthesis led to a proposal for four trait dimensions indexed by leaf mass per area and lifespan (i.e. the LES), seed mass and seed output, leaf and twig size, and plant height (Westoby et al., 2002). Wright et al. (2007) found three independent trait dimensions represented by specific leaf area (SLA), seed/fruit size and leaf size in seven neotropical forests. The most extensive study (in terms of the number of species considered) to date was by Díaz et al. (2016), who showed that variation among species in height, stem specific density, leaf mass per area, seed mass, and nitrogen per unit mass (N_{mass}) could be reduced to two dimensions, the first indexing plant size, the second the LES. However, these various studies have considered only a limited set of traits or combined information from disparate sources, and did not attempt to quantify the climatic or phylogenetic controls on traits. In this paper, we examine a suite of leaf traits, using co-located measurements to quantify the contributions of climate, site, life form and phylogeny to trait variation at a large geographic scale. Our analysis is based on an extensive data set (Wang et al., 2018), containing information on multiple leaf traits from different regions of China. We focused on seven leaf traits that together capture many functions of plants (Table S1). The traits considered include four commonly measured traits: leaf area (LA), specific leaf area (SLA), leaf dry matter content (LDMC) and leaf nitrogen per unit area (N_{area}) , and also three traits that determine photosynthetic rates: maximum carboxylation rate (V_{cmax}) and maximum electron transport rate (J_{max}) , derived from

gas exchange measurements in the field, and the ratio of intercellular to ambient

carbon dioxide (CO₂) concentration (often denoted as c_i : c_a but called χ here following

Prentice et al., 2014) derived from leaf stable carbon isotope (δ^{13} C) measurements.

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We used multivariate analysis to quantify the dimensionality of variation in this set of traits, and the nature and dimensionality of trait-climate relationships. We used variance partitioning to attribute trait variations (for all traits, and each trait separately) to differences among sites, climate variations across sites, and distinctions among life forms and plant families. We finally applied the trait-climate relationships derived from the data set to various global datasets for specific traits, in order to assess their generality and potential wider application.

Materials and methods

Dataset description

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The data are derived from the China Plant Trait Database (Wang et al., 2018), which contains information on morphological, physical, chemical and photosynthetic traits from 122 sites and provides information on more than 1215 species. The database was designed to provide comprehensive sampling of different vegetation types and climates. It employs a standardized taxonomy and includes information on life form, plant family, site location, elevation, and climate. LA, SLA, N_{area}, LDMC and leaf δ¹³C data from multiple species were available at 48 sites, including 483 species altogether, distributed through the eastern half of China (Fig. 1a, Table S2). The sites from northeastern China are distributed along an aridity gradient (Prentice et al., 2011), including steppes, grasslands and temperate deciduous broadleaf forests. The sites from southwestern China represent tropical and subtropical evergreen broadleaf forests, and tropical dry woodlands. Temperate deciduous forests in central China and boreal forests in the far north of China were also included. Collectively these data cover the principal climatic and vegetation zones of the region (Fig. 1b). At each site, a stratified sampling strategy ensured that measurements were available for the main species in each canopy stratum, including up to 25 species of trees. Species were classified by life form as trees, small trees, lianas, shrubs, forbs and graminoids. Bamboos, herbaceous climbers, geophytes and pteridophytes were present only in

- small numbers in the dataset and were not included in our analysis. Fig. S1 shows frequency distributions of each trait within each life form for forest and non-forest
- sites. Table S3 lists the total number of samples in each class.
- Details of trait measurement methods can be found in Wang et al. (2018). LA, SLA,
- N_{area} and LDMC were measured on samples collected in the field following standard
- 140 protocols (Cornelissen et al., 2003). LA was taken as the projected area of a leaf, or
- leaflet in the case of compound leaves. $V_{\rm cmax}$ was calculated from the light-saturated
- rate of net CO₂ fixation at ambient CO₂ (A_{sat}) using the so-called one-point method,
- which provides a rapid and effective alternative to the measurement of a full $A-c_i$
- 144 curve (De Kauwe et al., 2016). J_{max} was calculated from the light-saturated rate of net
- 145 CO₂ fixation at high CO₂ (A_{max}). Both V_{cmax} and J_{max} were adjusted to a standard
- temperature of 25°C using the methods proposed by Niinemets et al. (2014). The
- adjusted values are called $V_{\text{cmax}25}$ and $J_{\text{max}25}$. Leaf δ^{13} C measurements were converted
- to 13 C discrimination and thence to χ , eliminating the effects of latitude and sampling
- 149 year as described in Cornwell *et al.* (2017):

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$$\delta^{13}C_{air,1992} = a * \left(\sin\left(\varphi * \frac{\pi}{180}\right)\right)^2 + \sin\left(\varphi * \frac{\pi}{180}\right) - c$$
 (1)

- where φ is latitude and a, b and c are parameters estimated by regression with values a
- 152 = 0.0819, b = 0.0983 and c = 7.7521 (Cornwell *et al.*, 2017), and

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$$\delta^{13}C_{air} = \delta^{13}C_{air,1992} + g(y - 1992)$$
 (2)

where y is the sampling year and g = -0.0467, and

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$$\chi = (\delta^{13}C_{air} - \delta^{13}C_{plant} - a')/(b' - a')$$
 (3)

- where a' is the discrimination against ${}^{13}\text{CO}_2$ during diffusion through stomata (4.4%)
- and b' is the discrimination against ${}^{13}\text{CO}_2$ during carboxylation (27‰) (Farquhar et al.,
- 158 1982). Cernusak et al. (2013) showed that about 80% of the variation in instantaneous

gas exchange measurements of χ could be accounted for by a linear relationship to δ^{13} C, supporting the use of equation (3). Estimates of χ based on δ^{13} C measurements are used here, however, because they reflect longer-term growth conditions better.

Three bioclimate variables adequately represent the controls on vegetation structure and composition across China (Wang *et al.*, 2013). These are the accumulated photosynthetically active radiation during the thermal growing season (PAR₀), defined as the period when daily temperature is above 0°C; the daily mean temperature during the thermal growing season (mGDD₀); and the ratio of mean annual precipitation to annual equilibrium evapotranspiration (moisture index, MI), calculated using SPLASH (Davis *et al.*, 2017). The primary data for the calculation of these bioclimatic variables were derived from 1814 meteorological stations (740 stations with data from 1971 to 2000, the rest from 1981 to 1990), interpolated to 1 km resolution with elevation as a covariate using ANUSPLIN V4.37 (Hutchinson 2007).

Gap filling

Photosynthetic measurements were only available for 14 sites in the China Plant Trait Database; however, these sites comprise 53% of the species represented in the data set. Photosynthetic measurements were not available for the temperate forests of Changbai Mountain, and the Inner Mongolia grasslands. In order to allow multivariate analysis of a larger data set, $V_{\rm cmax}$ values for species at these sites were gap-filled using a back-propagation neural network using LMA, $N_{\rm area}$, LA, χ and moisture index (MI) as predictors (newff function in Matlab 2010a). The neural network is a machine learning technique that often provides better performance than conventional statistical methods for this type of application (Paruelo *et al.*, 1997; Papale *et al.*, 2003; Moffat *et al.*, 2010). The data were divided into two parts: a calibration data set used to determine the weights in the neural network (75% of data points), and a validation data set used to assess the network performance (25% of data points). The method achieved an acceptable accuracy with $R^2 = 0.49$ between observed and predicted

values for the calibration data set and 0.50 for the validation data set. J_{max} values were then estimated from V_{cmax} values using a linear regression fitted to data from all sites where both A_{sat} and A_{max} were measured. The regression equation used for gap-filling is $\ln J_{max,25} = -0.0221 \text{ mGDD}_0 + 0.7329 \ln V_{cmax,25} + 2.0362 (R^2 = 0.75, P < 0.01)$.

Multivariate analysis and variance partitioning

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Principal components analysis (PCA) and redundancy analysis (RDA) are powerful multivariate analysis techniques with many ecological applications (White et al., 2005; Maire et al., 2015; Scheibe et al., 2015). As a dimensionality reduction technique, PCA projects a set of data on correlated variables on to a series of composite, uncorrelated variables called principal components (James et al., 1990). In RDA, these variables are chosen to maximize the extent of their correlation with a set of predictor variables (Borcard et al., 1992) and are therefore described as "constrained" axes of variation. RDA also extracts further "unconstrained" axes, which are the principal components of the variation that remains after the fitted effects of the predictor variables have been removed. Here, PCA is used to analyse trait covariation; RDA is used to analyse the relationships of trait variation to climate variables; and the unconstrained axes of RDA are used to characterize the residual (within-site) variation in traits. These analyses were performed using the vegan package in R (Oksanen et al., 2017). LA was square-root transformed before analysis to yield a linear measure of leaf size. χ was logit-transformed (logit $\chi = \ln \left[\chi/(1-\chi) \right]$). All other traits (including \sqrt{LA}) were natural log-transformed. All traits were thus converted to dimensionless quantities in the range $(-\infty, \infty)$, allowing PCA and RDA to be carried out using the covariance matrix among traits with no need for further standardization. Each trait thereby has its 'natural' weight in the analysis. For log-transformed variables, this treatment implies that a trait with, say, 10-fold variation has twice the weight of a trait with 5-fold variation. The weight can be quantified by the standard deviation of the transformed variables (ln \sqrt{LA} : 1.17, ln SLA: 0.50, ln LDMC: 0.38, ln N_{area} : 0.59, ln

213 $V_{\text{cmax}25}$: 0.58, In $J_{\text{max}25}$: 0.48, logit χ : 1.37; see also Table 3). PCA and RDA were 214

repeated using only the species-site combinations for which actual (as opposed to

215 gap-filled) photosynthetic trait data were available (Figs S2-S4, Tables S4-S5).

Variation partitioning quantifies the amount of variation in a predicted quantity (in multiple regression) or set of quantities (in RDA) that can be explained by different groups of predictors (Legendre & Legendre, 2012). We used the Legendre method (Legendre & Anderson, 1999; Peres-Neto et al., 2006; Meng et al., 2015), which explicitly accounts for correlations between groups by distinguishing unique and overlapping contributions from each group. The results are most conveniently displayed as Venn diagrams. The method was used here with RDA to assign trait variation to components linked to climate, sites, life forms, families, and the

Trait prediction

intersections of these controls.

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226 We evaluated the predictive power of the fitted trait-climate relationships in the RDA

analysis, first on the data set as a whole and then using a cross-validation approach

(Picard & Cook, 1984; Kohavi 1995). We performed five iterations, in which 80% of

the data was used for training and 20% retained for validation. The average

root-mean-squared error (RMSE) across all five trials provides the final measure of

goodness-of-fit.

232 The general predictive power of the trait-climate relationships was then tested using

four independent global trait data sets: leaf economics traits (SLA, LDMC, N_{area}) from

234 Wright et al. (2004); $\sqrt{\text{LA}}$ from Wright et al. (2017); photosynthetic traits ($V_{\text{cmax}25}$.

 $J_{\text{max}25}$) from De Kauwe et al. (2016), including data from Bahar et al. (2017); and χ

from Cornwell et al. (2017) (Table S6). Each of these data sets provides geolocated

site-based measurements across continents, vegetation types and climates (Figure S5).

We derived climate variables for each site from the nearest 10-minute grid cell in the

239 CRU 2.0 dataset (New et al. 2002), which provides long-term monthly means of

temperature, precipitation, and sunshine duration for the standard period 1961-1990.

PAR₀, mGDD₀, and MI were calculated in the same way as for the sites in China, using

242 SPLASH to calculate MI (Davis *et al.*, 2017).

We screened out measurements from sites in the global data sets where MI > 1.4 or

 $mGDD_0 < 10$ because these are beyond the limits of the climates sampled in China.

Some of the δ^{13} C measurements in Cornwell *et al.* (2017) are < -30%. We assume that

these reflect incomplete mixing of CO₂ between the free atmosphere and the forest

understorey. We excluded these measurements. The number of sites and individual

measurements from each global data set used to test the climate-trait predictions is

shown in Table S6. Trait values at each global site were directly predicted from climate

inputs, using the RDA model previously derived from the data in China. Ordinary

least-squares regression was used to compare observed (y) with predicted (x) trait

values.

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Results

Four dimensions of trait variation

255 PCA of traits from all species and sampling sites revealed four independent axes of trait

variation (Fig. 2, Table 1). The first four principal components together account for 95%

of total trait variation. The first two axes are dominated by LA and χ , orthogonal to one

another. These two axes together account for 79% of total trait variation: this large

fraction draws attention to the large span of variability in these traits, especially leaf

area. The third axis, accounting for 11% of total trait variation, primarily represents the

LES, with SLA opposed to N_{area} and LDMC. The plot of axis 3 against axis 4, which

accounts for 6% of total trait variation, shows that $V_{\rm cmax}$ and $J_{\rm max}$ vary closely together,

but orthogonally to the LES.

Analysis based on sites with complete data only (Fig. S2, Table S4) shows that the first four principal components have similar explanatory power to the main analysis (93%) and, although the axes are rotated with respect to the axes derived from the larger data set, they show the same four dimensions of variation with LA, LES, photosynthetic capacity and χ varying independently of one another. The patterns of trait covariation can also be seen by examining the matrix of pairwise correlations between traits (Fig. S6). The differences between Fig. S6(a) based on the gap-filled data set, and Fig. S6(b) based on sites with complete data, show the (slight) effect of gap-filling. $V_{\rm cmax}$ and $J_{\rm max}$ are highly correlated (0.84) before gap filling. The largest difference is that the negative correlations of both $V_{\rm cmax}$ and $J_{\rm max}$ with leaf area *increase* due to the gap filling. This evidently does not contradict our inference from PCA on the gap-filled data set, i.e. that photosynthetic capacities are largely uncorrelated with the other traits.

Trait variation related to climate

The three bioclimatic variables together account for 37% of trait variation (Table 2). Three successive RDA axes (Fig. 3, Table 2) describe the patterns of trait variation with climate, and show that the between-site patterns of trait covariation imposed by climatic gradients differ from those found in the data set as a whole. The first RDA axis is overwhelmingly dominant, and is related to the gradient of MI from desert-steppe to moist forests. LA and χ vary together along this gradient, with both large leaves and large χ characteristic of wetter environments. The second RDA axis accounts for 2% of trait variation, and is related to the covariation of mean growing-season temperature and total growing-season light availability along the latitudinal gradient from the boreal zone to the tropics. Trait variation on this axis resembles the LES: warmer, higher irradiance climates are characterized by plants with lower SLA, higher LDMC and higher $N_{\rm area}$. The third RDA axis accounts for only 0.4% of trait variation. Analysis based on sites with complete data only (Fig. S3, Table S5) shows the same patterns.

Residual trait variation, unrelated to climate

The unconstrained axes (or residual principal components) calculated by RDA after climatic differences among sites have been accounted for (Fig. 4, Table 2) provide insight into trait variation that is expressed within sites and across all climates. The patterns of this residual variation, as shown by the first four unconstrained axes, are similar to the patterns shown by the principal components of the whole data set (Fig. 2, Table 1), with evidence for four independent dimensions of variation associated with successive components dominated by χ , LA, LES traits and photosynthetic capacities, respectively. Analysis based on sites with complete data only (Fig. S4, Table S5) shows the same four dimensions.

The same general patterns of non-climate-related trait covariation are also clear on inspection of the partial correlations among transformed trait values, after the effects of climatic predictors have been removed (Fig. 5). Deeper colours in Fig. 5 indicate larger absolute magnitudes of correlation. The traits can be seen to fall into four blocks: one comprising $V_{\rm cmax}$ and $J_{\rm max}$ (positively correlated), one comprising the traits that contribute to the LES (SLA negatively correlated with LDMC and $N_{\rm area}$), χ , and LA. While χ shows almost no correlation with any of the other traits, LA is weakly negatively correlated with $V_{\rm cmax}$ and $J_{\rm max}$ (Fig. 5), as is SLA.

Multiple controls of trait variation

Venn diagrams (Fig. 6) summarize the percentage contributions of climate, site, life form and family (including intersecting contributions) to total trait variation, and to variation in each separate trait. The intersection regions represent trait variation that cannot be unambiguously attributed to one control or another, because of correlations among the controls. For example, substantial intersections between climate and family occur because these controls are not independent: different families are selected for in different climates. Anomalously large values are highlighted in bold in Fig. 6 and one

317 anomalously small value indicated by italics. No values are shown for climate 318 independently of site, because differences in climate are determined by site locations. 319 Table 3 also shows the total percentage of variance associated with each control 320 (including intersections with other controls). 321 Considering the variation among all traits together (Fig. 6), climate, site, family and 322 life form jointly account for 70% of total trait variance. The most important features 323 of the partitioning are (1) the joint effect of climate with family (23%), which is the 324 dominant driver of trait variation in this dataset; (2) the substantial fraction of 325 variance due to family alone (17%), independent of climate or life form; and (3) the 326 fact that most of the total variance associated with life form (16%) is also linked to 327 climate (8%). There is some additional effect of climate independent of family (8%); 328 and some effect of site independent of climate (12%), which is presumably related to 329 edaphic or microclimatic factors. 330 The partitioning of trait variance for individual traits (Fig. 6) generally resembles that for all traits. However, 48% of total trait variation in LDMC is linked to family, and 331 332 41% linked to family independent of other controls. Only 4% of the variation in 333 LDMC is linked to climate, and none to climate and family together. For SLA, 41% of 334 total trait variation is linked to family (with 14% linked to family and life form together independent of other controls); 15% is linked to climate, but only 4% to 335 336 climate and family together. These anomalies indicate a particularly strong 337 phylogenetic component to variation in LDMC and, to a lesser extent, SLA. The 338 unexplained variation is greater for $V_{\text{cmax}25}$ (47%) and $J_{\text{max}25}$ (41%) than for the other 339 traits. 340 After climate, site and family effects have been accounted for, the remaining 341 (independent) contribution of life form to trait variation is small. The total life-form 342 contribution is < 10% for all traits except LA and χ , and the unique contribution of life

form independent of all other controls is very slight, < 2.5% for all traits. Forbs and

graminoids show different ranges of trait values in forest and non-forest vegetation (Fig. S1). Specifically, SLA and LDMC of forbs and graminoids decrease between forests and non-forests while N_{area} , V_{cmax} and J_{max} increase. That is, for all these traits, life forms occupying the understorey in forest vegetation become more 'tree-like' in non-forest vegetation, suggesting that these traits are more determined by the light environment than by any intrinsic difference among life forms.

Worldwide prediction of traits based on the observed climate-trait relationships

The RDA analyses show that climate (including indirect effects mediated by selection for life forms and families) is the major determinant of trait variation for most of the traits examined, except for LDMC and SLA, which show a substantial independent phylogenetic component. This generalization is supported by predictions of the mean site values for each trait (Fig S7). At species level, the adjusted R^2 between observed and predicted values for LDMC is only 0.08, and for SLA 0.16 (Table S7), while the relationship is better for other traits – from 0.24 for $V_{\rm cmax25}$ to 0.52 for $\sqrt{\rm LA}$. The average adjusted R^2 across traits is 0.28. Partitioning the data into woody and non-woody components has little impact on the quality of the prediction for most traits, but prediction of LDMC and SLA is better for non-woody than woody species (Table S7). Although predictability is imperfect, because of the (demonstrated) influence of non-climatic factors on all of the traits, these analyses nonetheless show that it is possible to predict all four dimensions of trait variation, to first order, from climate.

The prediction of trait values in global data sets provides a more stringent test of the universality of the derived climate-trait relationships (Fig. 7, Table 4). At site level, the lowest adjusted R^2 value between observed and predicted trait values is again for LDMC (0.01), but for SLA it is 0.31. For other traits, adjusted R^2 ranged from 0.25 (J_{max}) to 0.34 ($\sqrt{\text{LA}}$). The average across traits is 0.31, excluding LDMC. The observed values for $\ln V_{\text{cmax}25}$ tend to be higher than the predicted values, whereas the observed values of $\ln SLA$ tend to be lower than the predicted values (Fig. 7).

However the regression slopes for these traits are not significantly different from unity (Table 4). The OLS regression slopes for $\ln \sqrt{\text{LA}}$, $J_{\text{max}25}$ and $\ln \chi$ are in the range from 0.48 to 1. RMSE values (Table 4) are larger in the global comparison than in the calibration set for $\ln \sqrt{\text{LA}}$ and SLA; but closely similar for N_{area} , $V_{\text{cmax}25}$ and $J_{\text{max}25}$, and χ . The average RMSE across traits excluding LDMC is slightly less in the global comparison (0.42) than in the calibration set (0.61).

Discussion

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The ecological significance of leaf-trait dimensions

The four dimensions of total leaf-trait variation reported here indicate the existence of independent variation among species in LA, χ , photosynthetic capacity, and the LES. The RDA based on climate shows a smaller dimensionality, with most of the variation concentrated on a single axis from wet to dry environments. LA is both expected and observed to increase with plant-available moisture, due to energy-balance constraints (Wright et al., 2017). χ is both expected and observed to increase with atmospheric moisture according to the least-cost hypothesis (Prentice et al., 2014). These hydroclimatic controls on both LA and χ are presumed to be the cause of (a) the dominance of a single dimension of trait-environment relationships across the region, related to moisture/aridity, and (b) the observed close covariation of LA and χ between sites along the aridity gradient – contrasting with their independence in the data as a whole. Analysis of the residual (non-climatic) component of trait variation however shows, once again, four independent dimensions, with a pattern closely similar to that shown in total leaf-trait variation, and orthogonal variation of LA and χ . Multivariate analysis confirms the universal nature of the LES, as indexed here by SLA, LDMC (which tends to be high when SLA is low), and N_{area} . Unlike N_{mass} (N concentration per unit mass), Narea increases with decreasing SLA because the structural component of leaf N increases in proportion to LMA (see e.g. Onoda et al., 2004, 2017; Wright *et al.*, 2005; Osnas *et al.*, 2013; Dong *et al.*, 2017a). The LES is identified in the PCA, and in the residual trait variation after consideration of climate effects in RDA. However, it also appears in the climatically constrained RDA as a second-order pattern correlated with the latitudinal gradient. In other words, there is a shift in the average position of species along the LES (towards lower SLA) with increasing growing-season length and warmth, although this shift accounts only for a small proportion (2%) of total trait variance. The LES reflects the inescapable linkage between high construction costs and long payback times of leaves with low SLA (Kikuzawa, 1991; Reich *et al.*, 1997; McMurtrie & Dewar, 2011; Funk & Cornwell, 2013). The shift towards lower-SLA leaves in warmer climates is primarily due to the shift of dominance from deciduous to evergreen woody plants. The increase in growing-season length (towards a year-round growing season in the tropics) favours longer-lived evergreen leaves with lower SLA in warmer climates, as shown here and in other studies.

Both the gap-filled data set and the non-gap-filled subset show that the two photosynthetic capacities ($V_{\rm cmax}$ and $J_{\rm max}$) covary closely (Fig. S6), as is expected from the co-ordination hypothesis – which predicts that leaves should not possess excess capacity in either carboxylation or electron transport, as photosynthesis depends on both (Chen *et al.*, 1993; Maire *et al.*, 2012). However both traits show substantial variation within sites. When $V_{\rm cmax}$ and $J_{\rm max}$ were entered into the analysis after adjustment to local growth temperature, as opposed to 25°C, the results were very similar (not shown). Opposite trends of variation in $V_{\rm cmax}$ and $J_{\rm max}$ are shown only in the (minor) third axis of the RDA, accounting for 0.4% of total trait variance and driven by differences among sites in summer temperature that are independent of the latitudinal gradient. This pattern is consistent with expectations, as a decline in the $J_{\rm max}$: $V_{\rm cmax}$ ratio with increasing temperature has been shown experimentally (Kattge & Knorr, 2007) and predicted theoretically (Wang *et al.*, 2017a). The decline is larger

when the two photosynthetic capacities are estimated at prevailing growth temperature, but persists when they are adjusted to 25°C.

Contributions to leaf trait variation

The variance partitioning results presented here demonstrate that family and climate effects (except for LDMC and SLA) overlap considerably. In other words, a substantial part of trait variation with climate is due to families replacing one another along environmental gradients. After family, climate and site effects have been taken into account, independent life-form effects become unimportant. Thus, to first order, the principal controls on trait variation in this data set are family identity, climate, and climatic selection among families. Additional effects of site (independent of climate) could in principle be due to microclimatic and/or edaphic differences among sites, which have not been investigated. LDMC and to a lesser extent SLA show stronger family effects than other traits, while the effects of climate on these traits appear to be largely independent of family identity.

Implications for vegetation modelling

Vegetation models based on continuous variation in trait space sample 'plants' from a continuum of trait values (e.g. Scheiter *et al.*, 2013; Fyllas *et al.*, 2014). This approach requires specifying which traits can vary; by how much; and the extent to which different traits covary, in other words, the effective dimensionality of trait space. Our analyses of leaf traits, including traits derived from stable isotope and gas exchange measurements, indicate that at least four independent dimensions of trait variation need to be considered; that realistic modelling of functional diversity must allow for within-site variation in each of these dimensions; and that environmental differences force patterns of trait covariation across sites that can be different from patterns observed within sites.

With the exception of LDMC, which shows a particularly strong phylogenetic component, the trait-environment relationships found here should be amenable to process-based modelling. The energy balance implications of leaf size (Michaletz *et al.*, 2016; Dong *et al.*, 2017b; Wright *et al.*, 2017) mean that this trait is crucial for survival, particularly in cold climates or in hot, dry climates. As the biophysical controls of leaf size are relatively well understood, it should be straightforward to build energy-balance constraints on leaf size into trait-based models. Shifts in the LES along environmental gradients could also be modelled, given the well-established relationship of leaf longevity and SLA (Wright *et al.*, 2004) and the experimentally determined variations of SLA with environmental factors (Poorter *et al.*, 2009). The distribution of SLA within communities could be represented by a pattern of covariation in leaf longevity, SLA, LDMC and the structural component of N_{area} , as shown here and in other studies.

The co-ordination hypothesis predicts both $V_{\rm cmax}$ and the ratio of $J_{\rm max}$ to $V_{\rm cmax}$, including the observed dependence of both quantities on growth temperature (Wang et al., 2017b). Large-scale patterns in $V_{\rm cmax}$ and the metabolic component of $N_{\rm area}$ can be predicted theoretically (Dong et al., 2017a). The co-ordination hypothesis also predicts the observed seasonal acclimation of $V_{\rm cmax}$ and $J_{\rm max}$ (Togashi et al., 2018). Thus, at the level of community mean values, it seems likely that $V_{\rm cmax}$ can be successfully modelled as a function of environment (Ali et al., 2016). A temperature-dependent ratio of $J_{\rm max}$ to $V_{\rm cmax}$ would then allow prediction of $J_{\rm max}$.

The CO₂ drawdown from air to leaf, indexed by χ , is predicted by most vegetation models by simultaneous solution of the FvCB equations to predict assimilation rate as a function of leaf-internal CO₂ (c_i) and the diffusion equation to predict c_i as a function of ambient CO₂ (c_a), stomatal conductance and assimilation rate (Farquhar *et al.*, 1980). Theoretically and empirically well-founded relationships between χ and

environmental variables (Wang *et al.*, 2017b) provide an alternative way to model χ directly as a function of environment, and thus to predict assimilation rates more straightforwardly than in many current models.

Challenges and future directions

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This analysis illustrates the power of large trait data sets spanning a large range of climates, and including measurements from multiple co-existing species at each field site, to reveal general patterns. It also shows the utility of multivariate analysis to summarize patterns, and variance partitioning to attribute trait variability to different (and sometimes intersecting) causes. But despite the availability of large plant-trait data compilations (e.g. Kattge et al., 2011), the number of sites that include all of any specified set of plant traits is often disappointingly small – because different research groups typically collect data on different sets of traits. There remains a need for more extensive trait data collection including photosynthetic traits and isotopic measurements in addition to conventional leaf traits, and for such data collection to extend to the full range of the world's climates. There has been a limited amount of comparative work, for example, on photosynthetic traits, which are essential for all process-based vegetation modelling. Moreover, compared to leaf traits, there is a paucity of data on other field-measurable traits (notably stem hydraulic properties) that may be equally important for plant functional ecology. As is well illustrated by the global data sets that we used to test the predictive capacity of trait-climate relationships, the site- and/or species-metadata available are often limited. There remains a need for extensive, targeted collection and analysis of plant trait data, including co-located morphological, gas-exchange and isotopic measurements, and spanning the world's major environmental and floristic gradients.

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Author contributions

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YY, HW, SPH and ICP collectively devised the analysis strategy and interpreted the results. YY carried out all of the statistical analyses and wrote the first draft of the manuscript. IJW provided additional advice on the analysis and interpretation of trait variation patterns. All authors provided input to the final draft.

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Figure legends

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Fig. 1 Geographical and climatic coverage of the trait dataset. The individual sites are shown as red dots superimposed on a simplified vegetation map of China in (a); these sites have been grouped into eight named regions. The distribution of sites in climate space is shown in (b), where MI is the moisture index defined as the ratio of mean annual precipitation to annual equilibrium evapotranspiration, PAR₀ is the accumulated photosynthetically active radiation during the thermal growing season, and the daily mean temperature during the thermal growing season (mGDD₀) is shown by the colour of the dots. The grey shading indicates the frequency of different climates, as defined by MI and PAR₀, in eastern China as a whole. Fig. 2 Trait dimensions from principal component analysis: grey circles are species-site combinations. The traits are LA: leaf area, SLA: specific leaf area, LDMC: leaf dry matter content, N_{area} : leaf nitrogen per unit area, $V_{\text{cmax}25}$: maximum carboxylation rate standardized to 25°C, $J_{\text{max}25}$: maximum electron transport rate standardized to 25°C, and χ : the ratio of intercellular to ambient CO₂ concentration. The four axes of variability related to LA, χ , the leaf economic spectrum and the photosynthetic traits are shown by coloured ellipses on each plot. Fig. 3 Climate-related trait dimensions from redundancy analysis: grey circles are species-site combinations and coloured dots signify named regions as defined in Fig. 1. The traits are: LA: leaf area, SLA: specific leaf area, LDMC: leaf dry matter content, N_{area} : leaf nitrogen per unit area, $V_{\text{cmax}25}$: maximum carboxylation rate standardized to 25°C, $J_{\text{max}25}$: maximum electron transport rate standardized to 25°C, and χ : the ratio of intercellular to ambient CO₂ concentration. The climate variables are the ratio of mean annual precipitation to annual equilibrium evapotranspiration (MI), the accumulated photosynthetically active radiation during the thermal growing season (PAR₀) and the daily mean temperature during the thermal growing season (mGDD₀).

- Fig. 4 Residual (climate-independent) dimensions of trait variation: grey circles are
- species-site combinations. The traits are: LA: leaf area, SLA: specific leaf area, LDMC:
- leaf dry matter content, N_{area} : leaf nitrogen per unit area, $V_{\text{cmax}25}$: maximum
- carboxylation rate standardized to 25°C, $J_{\text{max}25}$: maximum electron transport rate
- standardized to 25°C, and γ : the ratio of intercellular to ambient CO₂ concentration.
- 780 Fig. 5 Partial correlations between traits, after removal of climate effects. The traits are:
- 781 LA: leaf area, SLA: specific leaf area, LDMC: leaf dry matter content, N_{area} : leaf
- 782 nitrogen per unit area, $V_{\text{cmax}25:}$ maximum carboxylation rate standardized to 25°C,
- 783 $J_{\text{max}25}$: maximum electron transport rate standardized to 25°C, and χ : the ratio of
- 784 intercellular to ambient CO₂ concentration. Colours indicate the strength of the
- 785 correlation, where dark blue indicates perfect correlation.
- 786 Fig. 6 Variance partitioning (%) for all traits considered together, and each trait
- separately. The traits are: LA: leaf area, SLA: specific leaf area, LDMC: leaf dry matter
- 788 content, N_{area} : leaf nitrogen per unit area, $V_{\text{cmax}25}$: maximum carboxylation rate
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- and χ : the ratio of intercellular to ambient CO₂ concentration.
- 791 Fig. 7 Predicting traits globally at site level, from the trait-climate relationships derived
- 792 from data in China. The traits are: LA: leaf area, SLA: specific leaf area, LDMC: leaf
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- 795 25°C, and χ : the ratio of intercellular to ambient CO₂ concentration. (a) Predicted
- 796 ln√LA versus observed ln√LA (Wright et al., 2017). (b) Predicted ln SLA versus
- observed ln SLA (Wright et al., 2004). (c) Predicted ln LDMC versus observed ln
- 798 LDMC (Wright et al., 2004). (d) Predicted $\ln N_{\text{area}}$ versus observed $\ln N_{\text{area}}$ (Wright et al.,
- 799 2004). (e) Predicted ln $V_{\text{cmax}25}$ versus observed ln $V_{\text{cmax}25}$ (De Kauwe et al., 2016). (f)
- Predicted $\ln J_{\text{max}25}$ versus observed $\ln J_{\text{max}25}$ (De Kauwe et al., 2016). (g) Predicted logit
- 801 χ versus observed logit χ (Cornwell et al., 2017). Red squares are site means.

Figures

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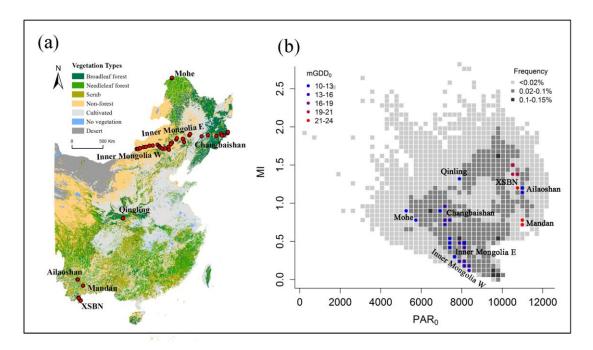


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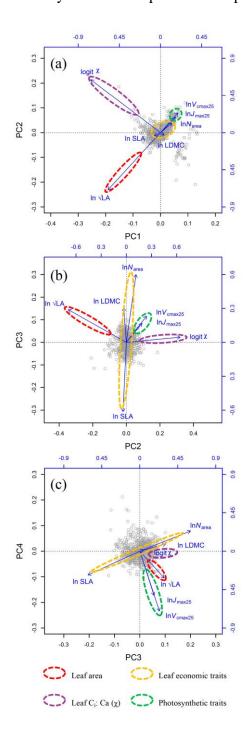


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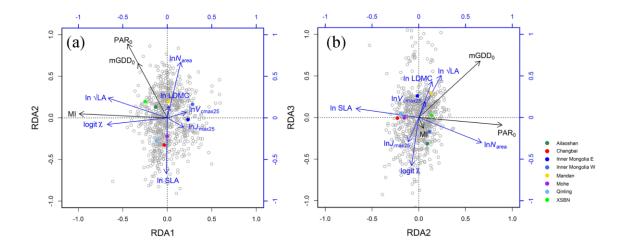


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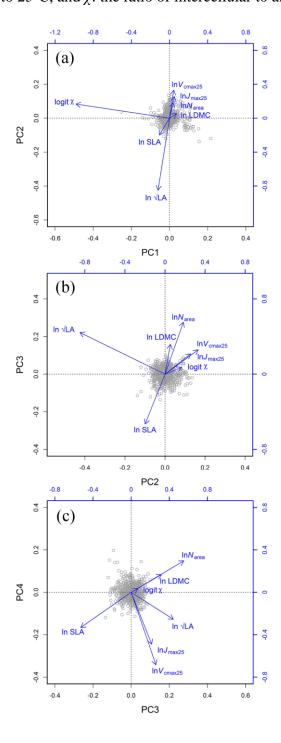


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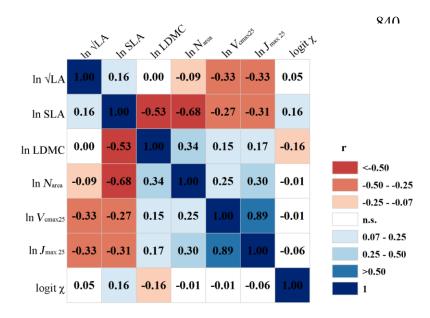


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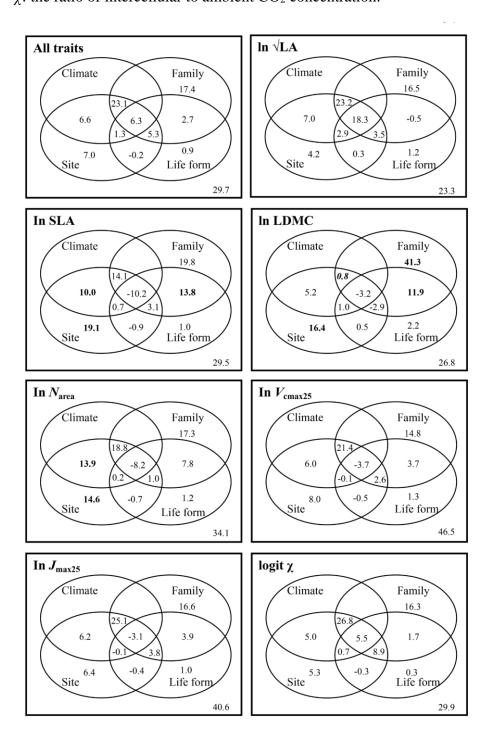


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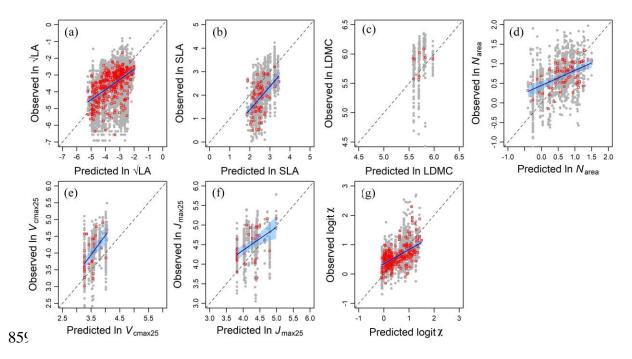


Table 1 Trait loadings, eigenvalues, and the percentage of trait variation explained by successive principal components in the trait PCA. Loadings > 0.3 in magnitude are shown in **bold**.

	PC1	PC2	PC3	PC4
ln √LA	-0.57	-0.69	0.29	-0.31
ln SLA	-0.07	-0.04	-0.61	-0.28
ln LDMC	0.04	-0.03	0.31	0.09
$\ln N_{ m area}$	0.12	0.11	0.60	0.24
$\ln V_{\rm cmax,25}$	0.19	0.24	0.23	-0.70
$\ln J_{ m max,25}$	0.16	0.19	0.17	-0.52
$\operatorname{logit} \chi$	-0.76	0.64	0.05	0.02
Eigenvalue	2.57	0.90	0.50	0.25
Explained (%)	58.0 20.4 11.3 5.		5.6	
Cumulative (%)	58.0	78.5	89.8	95.4

Table 2 Trait loadings, eigenvalues, and the percentage of trait variation explained by successive RDA axes (constrained by climate) and residual principal components, with axes 1 and 2 mirrored to facilitate comparison with the PCA. Loadings > 0.3 in magnitude are shown in **bold**.

	RDA1	RDA2	RDA3	PC1	PC2	PC3	PC4
ln √LA	-0.66	0.24	0.51	0.12	-0.85	-0.44	0.25
ln SLA	-0.01	-0.67	0.11	0.11	-0.20	0.53	0.33
ln LDMC	0.02	0.14	0.43	-0.08	0.05	-0.32	-0.17
ln N _{area}	0.15	0.67	-0.30	-0.04	0.18	-0.55	-0.30
ln V _{cmax,25}	0.22	0.07	0.19	-0.04	0.33	-0.26	0.68
$\ln J_{ m max,25}$	0.18	-0.11	-0.29	-0.05	0.26	-0.22	0.49
\log it χ	-0.67	-0.08	-0.58	0.98	0.17	-0.07	-0.04
Eigenvalue	1.55	0.08	0.02	1.19	0.75	0.42	0.24
Explained (%)	34.9	1.8	0.4	26.8	17.0	9.6	5.3
Cumulative (%)	34.9	36.7	37.1	63.9	80.9	90.5	95.9

Table 3 Total contributions (%) of climate, family, site and life form to trait variation.

Standard deviations (weights) of the transformed variables are also given.

	All traits	ln √LA	ln SLA	ln LDMC	ln N _{area}	$\ln V_{\rm cmax25}$	$\ln J_{ m max25}$	logit χ
Weights		1.17	0.50	0.38	0.59	0.58	0.48	1.37
Climate	37.3	51.4	14.6	3.7	24.7	23.6	28.1	38.0
Family	54.8	61.0	40.5	48.0	36.7	38.8	46.3	59.0
Site	49.4	59.4	35.9	17.8	39.6	33.7	37.9	51.8
Life form	16.3	25.8	7.5	9.4	1.3	3.4	5.1	16.7

Table 4 Prediction accuracy of the trait-climate RDA model for independent global data sets at site level. * indicates that the slope is significantly different from 1 (P < 0.01), # indicates that the intercept is significantly different from 0 (P < 0.01). ** indicates that the regression is significant (P < 0.01).

Traits	Slope	Intercept	R_{adj}^2	n	RMSE	Source of data
ln √LA	0.60*	-1.45#	0.34**	388	0.70	Wright et al. (2017)
	(0.52, 0.70)	(-1.72, -1.10)				
ln SLA	0.99	-0.61	0.31**	87	0.53	Wright et al. (2004)
	(0.68, 1.31)	(-1.41, 0.19)				
ln LDMC	n.s.	n.s.	0.01	9	0.20	Wright et al. (2004)
ln N _{area}	0.38^{*}	0.45#	0.28**	77	0.26	Wright et al. (2004)
	(0.24, 0.52)	(0.34, 0.56)				
$\lnV_{\mathrm{cmax}25}$	1.16	-0.11	0.33**	38	0.40	De Kauwe et al.
	(0.62, 1.69)	(-1.97, 1.76)				(2016)
$\ln J_{ m max25}$	0.59*	1.99#	0.25**	38	0.33	De Kauwe et al.
	(0.27, 0.92)	(0.62, 3.36)				(2016)
logit χ	0.48^{*}	0.35#	0.33**	281	0.29	Cornwell et al. (2017)
	(0.40, 0.57)	(0.30, 0.40)				