

Phylogenetic non-independence in rates of trait evolution

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1 Phylogenetic non-independence in rates of

2 trait evolution

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4	RUNNING HEAD: Phylogenetic non-independence in rates
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14 ABSTRACT

15 Statistical non-independence of species' biological traits is recognized in most traits under 16 selection. Yet, whether or not the evolutionary rates of such biological traits are statistically 17 non-independent remains to be tested. Here we test the hypothesis that phenotypic 18 evolutionary rates are non-independent, i.e. contain phylogenetic signal, using empirical rates 19 of evolution in three separate traits: body mass in mammals; beak shape in birds; and bite force 20 in amniotes. Specifically, we test whether rates are non-independent throughout the 21 evolutionary history of each tree. We find evidence for phylogenetic signal in evolutionary rates 22 in all three case studies. While phylogenetic signal diminishes deeper in time, this is reflective 23 of statistical power owing to small sample and effect sizes. When effect size is large, e.g., owing 24 to the presence of fossil tips, we detect high phylogenetic signals even in deeper time slices. 25 Thus, we recommend that rates be treated as being non-independent throughout the 26 evolutionary history of the group of organisms under study, and any summaries or analyses of 27 rates through time – including associations of rates with traits – need account for the undesired 28 effects of shared ancestry. 29 30 KEY WORDS: evolutionary rates; trait evolution; phylogeny; phylogenetic comparative

31 methods; phylogenetic signal

32

34 Descent with modification [1] is of fundamental importance to evolution and is recognized in 35 most traits under selection. Through evolutionary time, trait values will be more similar in 36 closely related species compared to distantly related species, since the variance of trait values 37 will be proportional to the divergence in evolutionary time [2]. This equates to shared ancestry, 38 i.e. phylogeny. The degree to which shared ancestry affects biological traits can thus be 39 described by the proportion of variance in trait data across a comparative sample of species 40 that can be explained by phylogenetic relations, or phylogenetic signal – e.g., K [3] or λ [4]. This 41 has statistical implications, i.e., phylogenetic non-independence. 42 While acknowledgement of phylogenetic non-independence in phenotypic trait data has 43 become common in comparative studies [5], it is not so for the rates of evolution (how fast 44 organisms' characteristics evolve). As rates are often used as proxies for adaptations [6, 7], it is 45 of immense importance that we understand their statistical properties, in particular, 46 phylogenetic non-independence. However, we have not been able to identify any study in the 47 literature that explicitly tests for phylogenetic signal in phenotypic evolutionary rates aside 48 from rare instances in which this was implied [8]. 49 Here, we test whether evolutionary rates contain phylogenetic signal using three 50 empirical case studies: body mass in mammals [9]; beak shape in birds [10]; and bite force in 51 terrestrial amniotes (ESM). Our basic premise is that if phylogenetic signal is detected in rates, 52 then rates evolve along the branches of a phylogenetic tree in proportion to the passage of 53 time and that closely related species are more similar in rates than distantly related species. 54 Naturally, this necessitates a non-homogenous distribution of rates across the branches of the 55 tree – i.e., variable-rates of phenotypic trait evolution [8, 11-14].

57 MATERIALS AND METHODS

58 We obtained 100 sets of phenotypic evolutionary rates and the associated time-calibrated 59 phylogenetic trees (time-trees) from the authors of the three case studies (ESM). As we would 60 expect rates along all branches of a phylogenetic tree to be affected by shared ancestry, not 61 just the tips, we tested for phylogenetic signal in rates along both terminal and internal 62 branches, by time-slicing the phylogenetic tree. We time-sliced the three time-trees at 1-Myr 63 intervals for the mammals and birds (167 and 109 time slices respectively), and at 5-Myr 64 intervals for amniotes (65 time slices) (see ESM for details). The latter interval was chosen for 65 amniotes owing to their much longer evolutionary history (approx. 350 Myr) compared to 66 mammals and birds. For each time-sliced tree, we matched the terminal branches to the 67 corresponding branches in the complete time-tree (Fig S1). We then assigned the 68 corresponding rates to those terminal branches on the time-sliced tree as tip trait values. We 69 fitted a maximum likelihood (ML) phylogenetic generalized least squares (GLS) model in 70 BayesTraits V3 to estimate phylogenetic signal λ in rates at the tips for each time-slice (GLS_{λ}). 71 We tested GLS_{λ} against the null model in which λ is fixed to 0 (GLS_{$\lambda=0$}) as the likelihood ratio 72 (LR_{λ}) between GLS_{λ} and $GLS_{\lambda=0}$ and determined significance using the χ^2 distribution (df=1). 73 When λ was significant in >95% of the sample in any given time slice, we determined that 74 phylogenetic signal was present in that time slice. We also compared the significance of an 75 alternative model in which λ is fixed to 1 (GLS_{$\lambda=1$}). The root estimate α of a GLS_{$\lambda=0$} model is the 76 equivalent of estimating the non-phylogenetic mean rate, while GLS_{λ} and $GLS_{\lambda=1}$ estimate the 77 phylogenetically corrected mean rates.

- 79
- 80 **RESULTS**

81 Overall, phylogenetic signal at the tips of the complete time trees are high (body mass in 82 mammals, median λ = 0.926; beaks in birds, median λ = 0.729; and bite force in amniotes, λ = 83 1), providing evidence for strong effects of shared ancestry in rates of phenotypic trait 84 evolution along the terminal branches. Phylogenetic signal in rates are generally high and 85 significant in at least 95% of the sample in younger time slices – younger than: 48 Myr ago 86 (mammals); 45 Myr ago (birds); and 30 Myr ago (amniotes) (Fig. 1; ESM). Phylogenetic signal 87 depreciates (drops in strength and significance) rapidly in deeper time slices (Fig. 1; ESM). Fixing 88 λ to 1 (GLS_{λ =1}) result in gualitatively similar patterns across time slices compared to when λ is 89 estimated (GLS_{λ}) (Fig. 1), but depreciation of λ start at younger time slices compared to GLS_{λ} 90 (Fig. 1).

91

92 **DISCUSSION**

Through our time-sliced GLS models on three datasets, we demonstrate that evolutionary rates of phenotypic traits are indeed phylogenetically non-independent – λ is significant and high, both along the terminal and internal branches (Fig. 1). Crucially, although λ ceases to be significant in deeper time slices in all trees tested (Fig. 1), this reduction in phylogenetic signal most likely depends on two aspects of the rates in the focal time slice: 1) number of tips [15] and 2) rate heterogeneity (ESM). Both reflect issues of statistical power with the former concerning sample size (as determined through simulations; ESM) and the latter effect size (as 100 evident from the effects of fossil tips; Fig 1C; ESM). Incidentally, un-sampled tips of any sort 101 (not just fossils) will likely increase rate heterogeneity should they be sampled. Additionally, 102 information contained at the tips of an ultrametric tree (e.g., trait values) is expected to be lost 103 progressively deeper in the tree (proportional to the phylogenetic variance-covariance 104 structure) as subsequent evolution towards the tips overprints ancestral information – this is an 105 issue plaguing phylogenetic comparative methods in general. Furthermore, since rates are 106 estimated from the phylogeny using models with constant rate evolution as the underlying 107 process of evolution, the resulting rates would inevitably contain phylogenetic signal. Whether 108 this is true or not, this does not alter (rather it reinforces) our argument that inferred rates 109 contain phylogenetic signal (regardless of the reason) and crucially that all downstream 110 summaries and analyses of rates thus must account for phylogenetic non-independence. Thus, 111 we argue that it is safest to assume that phylogenetic signal will be present and strong in 112 deeper time slices [8]. 113 An important implication here is that as rates will be statistically non-independent at 114 various time intervals throughout the history of the clade of interest, patterns gleaned from 115 simple summaries (e.g., interval means) of rates-through-time (RTT) can potentially be 116 misleading. Simple RTT plots are prevalent in recent literature [e.g., 10, 16, 17-21], the profiles

117 of which routinely interpreted at face value, with peaks and troughs representing periods of

118 bursts and declines in rates [16, 17, 19]. However, accounting for phylogenetic non-

independence by assuming strong phylogenetic signal uniformly across all time slices [8] – i.e.

120 phylogenetic mean α from our GLS_{$\lambda=1$} models across time slices – results in phylo-RTT profiles

121 that are often different from those of non-phylogenetic RTT (Fig. 2). Thus, non-phylogenetic

RTT profiles cannot be taken at face value without knowledge of phylogenetic signal through
time. More crucially, this implies that statistical analyses of rates need also account for
phylogenetic non-independence. Testing hypotheses of external influences (ecological or
environmental) on rates of evolution would require the application of appropriate phylogenetic
statistical methods – e.g. phylogenetic regression models [14, 22]. Not doing so will run the risk
of resulting in misleading statistical results [2, 5].

128 As phenotypic evolutionary rates have been interpreted as reflecting the intensity of 129 natural selection [6, 14], that they contain phylogenetic signal implies that ancestors and 130 descendants as well as closely related species either: 1) share intrinsic mechanisms for selection 131 responsiveness (e.g., genetic predisposition); 2) share similar levels of extrinsic selection 132 pressures (e.g., similar ecological niches, environments, etc); or 3) both. Two (or more) species 133 descended from a parent species would be expected to start their respective independent 134 evolution with the same level of intrinsic responsiveness as well as extrinsic selection pressures, 135 and thus at the phenotypic evolutionary rate, of the parent species. The daughter species then 136 would be subject to independent genetic mutations and selection pressures depending on their 137 respective environments.

However, this is not to say that descendent rates are rigidly constrained by ancestry; exceptional rate shifts along individual branches are widely observed in many traits across various groups of organisms [8, 9, 11, 14, 23]. Such exceptional rate shifts can often be orders of magnitude greater than the background rate and occur instantaneously (with respect to geological time) such that the effects of ancestry may be marginal.

In conclusion, our analyses demonstrate that rates of phenotypic evolution estimated from phylogenetic trees using models of trait evolution are statistically non-independent (most likely owing to shared ancestry), across the tips and through time – we posit that our results are conservative with phylogenetic signal actually being more prevalent. Thus, we recommend that phylogenetic non-independence be accounted for in summaries and analyses of evolutionary rates through time, using appropriate phylogenetic comparative methods.

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230	Supporting data are made available through the Open Science Framework
231	(https://osf.io/pn4ma/?view_only=2de49c7ad61944ed97c373783a7d1956)[24], and described
232	in ESM.
233	
234	COMPETING INTERESTS:
235	The authors declare no competing interests.

236 FIGURE LEGENDS

237 Figure 1. Phylogenetic signal in rates of trait evolution through time.

238 Phylogenetic signal (λ) was estimated across time sliced phylogenetic trees (top row) in three

- 239 independent data sets: A, mammalian body mass; B, avian beak shape; and C, aminote bite
- 240 force. Faint lines represent each of the 100 samples with the bold line representing the median
- 241 λ . The percentage of the sample in which LR_{λ} (likelihood ratio between GLS_{λ} and GLS_{$\lambda=0$}) was
- significant is shown for each time slice (second row). Further, the fit of $GLS_{\lambda=1}$ is shown as the
- 243 percentage of the sample in which $LR_{\lambda=1}$ (likelihood ratio between $GLS_{\lambda=1}$ and $GLS_{\lambda=0}$) was
- significant for each time slice (third row). Red dashed line represents the 95% threshold. Blue

245 dashed line (top) represents the time slice for the 95% threshold as determined through

simulations (Fig. S3). The relationship between the percentage of significant λ and N_{Tips} (bottom)

row) shows a clear drop off in the percentage from 95% of the sample (red box).

248

Figure 2. Mean evolutionary rates through time compared to phylogenetically corrected

250 mean rates. Simple mean values of evolutionary rates at each time slice across the three

datasets (A, mammalian body mass; B, avian beak shape; C, amniote bite force) show distinctive

- 252 patterns of rates through time. However, these patterns are far less prominent in
- 253 phylogenetically corrected mean rates (α) through time. α are the phylogenetic root estimates

of the $GLS_{\lambda=1}$ model. Faint lines represent each MCMC run while the bold line shows the median value for each time slice.

256