Remittances and protests against crime in Mexico

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Remittances and Protests against Crime in Mexico

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Abstract

The resource mobilization theory has long emphasized the role of resources in facilitating collective mobilization. In turn, recent research on crime and insecurity in Mexico has drawn attention to the role of local networks of solidarity in facilitating mobilization against crime. We rely on these two literatures to propose that remittances – that is, the resources that emigrants send to their relatives left behind – deserve attention as international determinants of this type of non-violent anti-crime mobilization. Further, relying on recent research on remittances' impact on political behavior, we hypothesize that the relationship between remittances and contentious action is non-linear, exhibiting a positive effect at low to moderate levels of inflows and declining at higher levels of remittances. We contend that at low to moderate levels, international remittances provide the necessary resources for collective activation. At greater levels of inflows, however, lessened economic and security grievances imply a decline in the probability of protesting. Overall, we show that emigrant remittances matter for organizing protests against criminality at the subnational level but that they produce both an engagement and disengagement effect, depending on the size of the inflows.

Introduction

In this article, we explore whether migrant remittances – the money emigrants send to their relatives and friends – help families left behind organize against crime in Mexico (Pansters 2018; Guerrero 2018; Trejo and Ley 2019).¹ In particular, we study whether remittances help recipients invest time and resources in peaceful protest against the persistent increase in insecurity, as well as against the Mexican state's inefficacy to protect its citizens amid this surge in violence (Cárdenas 2016; Martínez 2017).

The resource mobilization theory is one approach that research on the determinants of protests has identified as a facilitator of mobilization (Gurr 1970; Brady et al. 1995; White et al. 2015). Without resources, this theory proposes, aggrieved audiences may be left without options for manifesting their discontent (McCarthy and Zald 2002). Resource availability, on

¹ In this article, we study the impact of individual remittances, rather than collective remittances, on protest. Collective remittances are raised by migrant organizations in destination countries and frequently finance public investments in migrants' origin communities (e.g., Burgess (2005) and Duquette-Rury (2014)). See also Pérez-Armendáriz and Duquette-Rury (2019) for a study of collective remittances and vigilantism in Mexico.

the other hand, can facilitate collective organization by freeing time to devote to civilian protests, enhancing the sense of citizens' perceived efficiency and facilitating coordination and organization efforts (Gurr 1970; Brady et al. 1995; White et al. 2015). In the particular case of protests against crime and insecurity in violent democracies, recent research has shown that besides financial commitment, the existence of local networks of solidarity in the form of a vibrant civil society also facilitates protests (Ley 2014; Dorff 2017). As we explain below, these networks help give visibility to victims' grief and demands, creating solidarity between victims and non-victims (Ley 2014; Rojo-Mendoza 2014; Durán-Martínez 2016; Dorff 2017). While the literature on mobilization against crime has acknowledged the crucial role of local networks of solidarity in animating victims and non-victims to protest (Ley 2014; Rojo-Mendoza 2014; Durán-Martínez 2016; Dorff 2017), it has overlooked the possibility that international migrant networks may also play a role in those mobilizations.

In this article, we argue that migrants' remittances should be considered determinants of protests against crime. Emigrants keep a host of relationships with their relatives left behind and frequently send financial support (World Bank 2006). We argue that these flows provide extra resources for the collective mobilization of those left behind. In hypothesizing about the relationship between remittances and protest, we rely on the thriving literature on the consequences of remittances for political behavior (Goodman and Hiskey 2008; Bravo 2009; Córdova and Hiskey 2019; Germano 2018). Often, this literature reports contradictory effects of remittances on political participation, with remittances being associated with both political engagement and disengagement (Goodman and Hiskey 2008; Bravo 2009; Pfutze 2014; Córdova and Hiskey 2015; Escribà-Folch et al. 2018; Duquette-Rury and Chen 2018). Some authors report a decrease in electoral participation among remittance recipients (Bravo 2009; Germano 2013, 2018; Pfutze 2014; Duquette-Rury and Chen 2018) while others show that remittances stimulate non-electoral political activism, such as belonging to civil organizations

or engaging in political discussions (Goodman and Hiskey 2008; Córdova and Hiskey 2019). Among the panoply of non-electoral political activities that remittances may affect, protest has barely been researched (Pérez-Armendáriz and Crow 2010; Dionne et al 2014; Escribà-Folch et al. 2018; Germano 2018).

We present a theory to fill this gap. In the particular case of protests against crime, we argue that remittances can cause both an increase and a decrease in the likelihood of protesting and that these effects are evident at different levels of remittance penetration at the local level. In other words, we argue that remittances have a non-linear effect on protests, by which higher volumes of remittances, while still increasing the probability of protesting, do so at a declining rate. We attribute this slowdown to lessened economic and security grievances, as well as to recipients' greater economic autonomy in settings where remittances flow in abundance (Adams and Page 2005; Doyle 2015; Escribà-Folch et al. 2018; López and Doyle 2019).

In developing our theory, we build bridges between the literature on local mobilization against crime and on transnational migrant involvement, with a focus on the vibrant research agenda examining remittances' political impacts in origin countries (Goodman and Hiskey 2008; Bravo 2009; Pfutze 2014; Córdova and Hiskey 2019; Escribà-Folch et al. 2018; Duquette-Rury and Chen 2018; Germano 2018). In so doing, our article contributes to a better understanding of the multifaceted consequences that remittances have in out-migration settings.

The article proceeds as follows. First, we give some background on the evolution of crime in Mexico, an example of a "violent democracy" "in which competitive elections, civil freedoms, and inclusive participation have taken root yet and the state does not control sub-state violence" (Pérez-Armendáriz 2019, 2). We also discuss the correlates of this violence. Second, we discuss the literature on remittances and the mechanisms by which remittances may shape protest against crime. Third, we present our data and empirical strategy. Our

analysis uses an original dataset coding over one-thousand instances of protest against crime in Mexico in the period, 2006–2012 (Ley 2014). We account for the endogenous nature of remittances and use an instrumental variable approach to test a *non-linear* relationship between remittances and protest against crime. We conclude with some reflections on the relevance of our findings and with suggestions for the research agenda ahead.

1. Crime and Protest Against Crime in Mexico

Rising insecurity in Mexico over the past decade has drawn international attention, but violence has been persistently present across the country throughout the twentieth and twenty-first centuries (Pansters 2018). This violence has been sustained by deep connections between the Mexican state and a diverse set of violent actors that proliferated over the course of decades – from *caciques* (local political bosses) to drug lords and vigilante groups (Pansters 2018). As noted by Pérez-Armendáriz (2019), Mexico stands as a “quintessential violent democracy,” having ongoing functional electoral institutions and corresponding civic engagement yet also being a state historically unable to hold the monopoly on violence.

It is against this backdrop that one must understand the “War on Drugs” initiated by Mexican president Felipe Calderón in 2006. A militarized strategy to fight organized crime amid the continued overlap between the Mexican state and drug trafficking organizations resulted in the multiplication of armed actors, and competition among them exacerbated violence and insecurity across the country (Guerrero 2018). Mexico is today the eighth most violent country in Latin America (WHO 2015). Between 2006 and 2016, more than 100,000 people died as a result of armed confrontations among Mexican cartels, their private armies, and the Mexican military and police forces (Justice in Mexico Organized Crime and Violence Reports).² In addition, more than 25,000 people have disappeared, not a few of them at the

² <https://justiceinmexico.org/publications/reports/> (Accessed 24 March, 2021).

hands of Mexican government forces (Human Rights Watch, 2013), while more than 300 local authorities, political candidates, and party activists have been direct victims of organized crime (Trejo and Ley 2019). During the Calderón and subsequent administrations, organized crime groups multiplied in Mexico, no longer consisting of a handful of cartels, but dozens of them, along with hundreds of street gangs (Guerrero 2018).

The deep connections between the Mexican state and organized crime, and the subsequent failure of security and judicial institutions to contain and punish crime (Cárdenas 2016), have discouraged Mexican citizens from reporting criminal activity, despite growing insecurity and their increasing personal encounters with violence (Smulovitz and Peruzzotti 2009). When the traditional institutional channels for reporting crime and achieving justice are weak, civil society can take action to hold governments accountable, expose governmental wrongdoing, or activate horizontal checks (Smulovitz and Peruzzotti 2009). Such has been the case in Mexico. In the face of rising violence, Mexican citizens have made an effort to keep government authorities accountable for the issue of insecurity through diverse non-electoral mechanisms, among which protest has been a recurring tool (Knox 2018). According to Ley (2014), between 2006 and 2012 – the peak period of criminal violence during the Calderón administration – more than a thousand protest events against crime and insecurity were voluntarily organized by civilians across Mexico. Through these actions, victims and their relatives have told of the violence they have experienced, revealed information on the collusion between public authorities and criminal groups, and denounced the many obstacles they face when attempting to report and prosecute their cases through judicial institutions.

As protest scholars have argued (McCarthy and Zald 2002; Brancati 2014), grievances are not enough for mobilization to take place. Specifically, as the resource mobilization theory has put forward, participation in social movements involves spending time, energy, and money

(Brady et al. 1995; White et al. 2015). Consequently, those groups with “few resources are less able to act on grievances or perceived injustices” (McCarthy and Zald 2002, 535). The availability of community resources – broadly defined as actual wealth coming from contributions, transfer payments, and per capita income, together with organizational skills and local associations and groups – has been found to be crucial for the likelihood that social movement organizations can mobilize (McCarthy et al. 1988; Khawaja 1994).

We argue that protests occurring amid criminal violence and in reaction to crime also require a diverse set of resources: in addition to monetary funds that enable the mobilization of relatives searching for their missing loved ones (Ahmed 2017), local networks of support have been crucial to the development of organized responses to crime (Ley 2014; Durán-Martínez 2016; Dorff 2017). The findings of these works are consistent with those that have examined collective action under conditions of high risk and that also emphasize the importance of networks of solidarity for protesting against crime (McAdam 1986; Loveman 1998).

According to Ley (2014), protest against crime in Mexico has been enabled by mobilizing formal (legally constituted) associations and informal arrangements of victims and non-victims that contribute in instrumental and non-instrumental ways to take action in response to crime.³ From a non-instrumental viewpoint, embeddedness in networks generates a sense of solidarity, as participants share their experiences, become aware of their commonalities, overcome fear, and foster a sense of collective indignation, as proposed by Loveman (1998) and Wood (2003). From an instrumental perspective – as widely noted by classic works on social networks and protests (DellaPorta 1988; Gould 1993; Klandermans 1997; McAdam 1986) – socialization within networks opens opportunities to participate, in addition to reducing associated risks (Keck and Sikkink 1998). Both instrumental and non-

³ The distinction between formal and informal social networks is well established in the literature on social capital. See, for instance, Putnam (2000).

instrumental functions within social networks are particularly relevant for protest against crime. As noted by Ley (2014), participants face relatively higher risks in such mobilizations, as they are more likely to be threatened or co-opted by *both* criminals *and* colluded government authorities or security forces that sustain protection networks and that have been historically intertwined with organized crime in Mexico (Pansters 2018). This double threat is unique to protest against crime because when participants in these mobilizations raise their voices to denounce violence, they are likely to attract the unwanted attention of criminals, as well as to put protection networks at risk and reveal the identity of colluding state officials (Ley 2014, 42).

Durán-Martínez (2016) further examines the conditions under which victims and non-victims may come together to develop organized responses to crime. The author finds that when criminals claim responsibility for acts or when attacks by criminals are publicly exposed, this shared knowledge on the responsible actor behind the violence helps mobilize non-victims by making them more sympathetic to victims and creating a sense of victimization among wider segments of the population. Finally, Dorff (2017) has shown the relevance of kinship ties in transforming victimization into political activation.

In view of the relevance of grievances to protest, of resource availability to mobilization, and of civilian networks in fostering protests, it is surprising that the role of emigrant connections and the money flows linked to these connections have remained unexplored in efforts to understand protest against crime in Mexican states. In what follows, we propose a theory based on how remittances shape the perception of grievances, potentially provide resources to the aggrieved, and may facilitate collective action against crime.

2. Remittances, Networks, and Grievances

The role of relatives abroad in supporting protest against crime at home should be given systematic attention: anecdotal and systematic evidence reveals that different types of local mobilization against crime, notably vigilantism, have relied on the individual and collective resources that migrants send back to their families (Pérez-Armendáriz and Duquette-Rury 2019; Ley et al. 2019). Moreover, instances of collective mobilization have often been led by individuals with migrant backgrounds.⁴ Here, we study the role of remittances as one of the international factors that may have facilitated domestic protest against crime in Mexico.

Through which mechanisms, however, may remittances affect the likelihood of recipients and their communities engaging in protest? As explained in the previous section, grievances are a necessary pre-existing condition, but insufficient if resources are lacking. Further, the existence of networks of solidarity helps victims make those grievances visible, catalyzing collective action. We posit that remittances' impact on grievances and on the motivations to collectively express and denounce those grievances are multifaceted and vary with the largess of the inflows of remittances.

On one hand, according to the resource mobilization theory (Brady et al. 1995; White et al. 2015), remittances provide extra income for households left behind, increasing the resources available for gathering information, coordinating, and making more time available to be politically active. As Ley (2014) explains, over the 2006–2012 period, mobilizations against crime in Mexico were initially led by human rights organizations, along with other formal networks such as churches and professional associations of journalists, teachers, and healthcare workers. All of these groups constitute *official* civil society organizations. However, around half the collective action processes were organized by *informal* networks linked through

⁴ “Quiénes son los líderes de los autodefensas,” *Milenio* 12 February, 2014 <http://www.milenio.com/policia/quienes-son-los-lideres-de-las-autodefensas> (accessed 23 March, 2021).

personal ties, not through institutionalized organizations - mainly networks of new victims and their relatives, who had never been part of protests in the past. Two additional informal networks contributed to protests against crime: neighborhood groups and informal merchants, whose interpersonal ties helped them react to shared concerns about rising insecurity and organize accordingly (Ley 2014).

Research on the political consequences of remittances has shown that remittance recipients are more likely to participate in informal cooperative networks, as remittances allow them to enter into “mutual help” and inter-household “risk-sharing arrangements” (Gallego and Mendola 2013, 722; Mendola 2017). In comparison to non-recipients, remittance recipients exhibit pro-social behavior, such as volunteering work to the community, donating, or helping strangers (Fransen 2015; Gerber and Torosyan 2013; Nikolova et al. 2017). Through these activities, remittances strengthen norms of reciprocity and trust, with effects that transcend the recipient household and generate organizational spillover effects at the community level (Pérez-Armendáriz and Crow 2010, 125; Mendola 2017, 289; Fransen 2015, 1295; Gerber and Torosyan 2013, 1283).⁵ Thus, remittances’ income effect can facilitate collective action and shape mobilization repertoires and technologies in remittance-receiving communities, from the types of protest activity a movement can organize and carry out to participants’ level of engagement, possibly transforming them from simple sympathizers to avid adherents. The expectation under the resource mobilization mechanism is straightforward: the income effect of migrants’ remittances should *increase* the likelihood of non-violent mobilization against crime in remittance-receiving communities.

On the other hand, *grievances* are regarded as a major cause of protests (Gurr 1970; Brancati 2014). Recent research shows that remittances shape economic and security

⁵ Some authors have found that remittances also enable the formation of rebel groups (Miller and Ritter 2014).

grievances. First, many scholars have found that the rise in income caused by remittances improves the living conditions of stayers, reduces poverty, and provides insurance against different types of risks (Chami et al. 2008; Adams and Page 2005; Yang and Choi 2007). Particularly in times of economic hardship, remittances reduce economic grievances. According to the predictions of political behavior models that emphasize how the economy shapes voting decisions – also known as “pocketbook” models of voting (Pop-Eleches and Pop-Eleches 2012) – remittance recipients may wrongly attribute the betterment of their economic situation to incumbents rather than to their relatives (Tertychnaya et al. 2018). This wrong attribution, in turn, may boost the approval of the government (Germano 2018; Tertychnaya et al. 2018). Also, remittances’ positive economic effects transcend the receiving household and have multiplier effects that improve the state of the economy at the community level (Durand et al. 1996; Zárate-Hoyos 2005). Overall, remittances reduce economic grievances.

Interestingly, new research on Mexico surmises that besides impacting economic grievances, receiving remittances is positively associated with recipients’ improved perceptions of their security situation in their neighborhoods or, in other words, with fewer *security* grievances. For instance, according to Doyle and López (2019), remittance recipients are also more likely to make investments to improve their safety, such as moving to safer neighborhoods or avoiding public transportation. For these reasons, remittance recipients evaluate their personal and family exposure to crime more positively than do those who do not receive remittances (Doyle and López 2019, 5). In other words, receiving remittances makes recipients feel safer in comparison to non-recipient peers. Fewer economic and security grievance leads to the expectation that remittances should be associated with *less* mobilization against crime.

Further, and related to the above, the extra income of remittances and the subsequent reduction in poverty provide recipient households with access to goods often provided by states, making them less dependent on publicly provided services (Yang and Choi 2007; Adams and Page 2005; Adida and Girod 2011; Ambrosius 2019). The bulk of research on this substitution effect has focused on policies related to education and health, as well as housing, public infrastructure, water, and sanitation; but preliminary evidence suggests that remittances could allow recipients to have greater and better access to public and private security and to afford legal assistance (López and Doyle 2019). As Brito et al. (2014, 8) state, “the family abroad can send extra money to pay for private security.” As a result of this substitution effect, if recipients can afford to become their own providers of security and, in general, feel safer than non-remittance recipients, remittance inflows could reduce recipients’ incentives to organize collectively with others in reaction to crime. With rising remittances and fewer grievances, recipients may feel less motivation to join efforts with others to protest against crime and instead “bond” with their closest family network (Fransen 2015). In sum, then, remittances’ substitution effect also anticipates *less* likelihood of engaging in protests against insecurity. High levels of remittance penetration would be associated with *no effect or a decreasing impact* of remittances on the likelihood of protesting.

Note that different mechanisms lead to different predictions regarding the relationship between remittance inflows and the likelihood of engaging in collective action against crime. Remittances provide the aggrieved with the resources to protest; but at the same time, remittances lower economic and security grievances and, consequently, the motives to protest. Rather than adjudicating between these two alternative observable implications, we postulate that these effects are prevalent at different degrees of local penetration of remittances. We contend that the resource mechanism prevails at a low to medium level of remittances, increasing the likelihood of protesting, but that the reduction in grievances and substitution

effects are evident at high levels of remittances. We posit that because remittances are primarily spent on basic needs (World Bank 2006; Chami et al. 2008; Adams and Page 2005), it may take sizable transfers for recipients to perceive a clear reduction in economic hardship and observe multiplier effects at the local level and an even larger inflow to be able to afford private means of protection and have better access to justice. For poor households, however, the best hope of attracting public attention and securing their own protection may well be devoting resources and time to organizing with other victims and non-victims to voice their grievances (Ley 2014; Phillips 2017; Ley et al. 2019). In other words, the option of disengaging from networks of mutual help or experiencing a tangible reduction in economic and security threats may be visible only in settings of high remittance presence. At those levels, recipients may be less dependent on risk-sharing arrangements, as remittances may ‘crowd out’ “their incentives to participate in activities that cross[ed] social divides” (Fransen 2015, 1297). Therefore, our working hypothesis is:

There is a non-linear relationship between the inflow of remittances and the probability of engagement in protests against crime: Remittances increase the likelihood of protesting at low to moderate levels, but as grievances decrease and substitution effects kick in at high levels of remittance inflows, remittances decrease or have no effect on the likelihood of protesting.

3. Data and Empirical Strategy

3.1. Data

For reasons that we explain next, the state level of analysis is the most suitable, given our data on protest events. But also, as explained below, because remittances have social and economic effects that transcend the individual and the household, we contend that our theory has observable implications at this level of analysis. Our dataset includes a balanced panel of

31 Mexican states (excluding Mexico City) and 22 time periods, from the first quarter of 2006 to the second quarter of 2011, yielding a total sample size of 682 state–quarter observations. Socioeconomic and political data were collected from various sources: National Statistical Institute (INEGI), Mexico’s Interior Ministry, the Central Bank of Mexico (Banxico), the Mexican Protest against Crime Dataset (MPC, Ley 2014), and Nyblade and O’Mahony (2014). The summary statistics of all variables, as well as their periodicity, are reported in Table 1.

*** Table 1 about here ***

Protest against Crime

The dependent variable, protest against crime, is a count variable recording the number of protests in state i and quarter t . These original data were collected at the state level, taken from the Mexican Protest Against Crime (MPC) Dataset (Ley, 2014). This unique database provides detailed information on 1,014 protest events against crime and insecurity that occurred during the 2006–2012 period across 31 Mexican states. It focuses exclusively on *non-violent* mobilization events organized by *citizens* as a means of *freely expressing their opposition* to a particularly violent event or general insecurity and demanding specific changes to security policies. As such, the MPC Dataset excludes protest events organized by criminal organizations against a particular branch of government or security force and public protests by police forces demanding greater security for their working conditions.⁶ The acts of citizen protest in this dataset include marches, demonstrations, road blockages, community meetings with

⁶ The MPC Dataset excludes protests that explicitly supported specific criminal organizations as part of their goals or such public expressions during the protest event, as in the case of demonstrations organized in 2010 and 2011 in Michoacán to show support for La Familia – a criminal organization that largely controlled the Mexican state of Michoacán between 2009 and 2011.

authorities, labor strikes, collective public prayers, sit-ins, collective press conferences and press releases, hunger strikes, flyer distribution, signature collection, and the occupation of government buildings. The dataset includes collective acts of protest in which a minimum of five individuals participated. Additionally, it is important to note that marches and demonstrations make up 86 percent of the total number of recorded events.

The information included in the MPC Dataset is mainly derived from a systematic review of fifty local newspapers and one national newspaper listed in Online Appendix Table A1.⁷ When newspaper information was insufficient, activity reports from human rights NGOs were used. While the MPC Dataset cannot provide an official count of protests against crime occurring in Mexico during the Calderón administration, the use of multiple sources significantly reduces sources of geographical or temporal bias in the dataset. Finally, the MPC Dataset collected the municipality of occurrence, along with other protest characteristics.⁸ However, municipal disaggregation would naturally generate a major urban bias, as victims and their relatives tended to organize protests in capital cities to have more impact and generate

⁷ The national daily newspaper was *Reforma*, which has extensive coverage of northern Mexico, one of the regions most affected by violence (Shirk and Wallman 2015) and has covered news on marches for peace since as early as 1999. The sample of fifty local dailies includes two newspapers for twelve states, one newspaper in nine states, and three newspapers in five states. Ley (2014) did not have direct access to news sources in the states of Campeche, Chiapas, Nayarit, Oaxaca, or Quintana Roo, but this limitation was partially overcome through the use of multiple regional newspaper sources. For instance, *Diario de Yucatán* covers information for the entire Yucatán Peninsula, including the states of Campeche and Quintana Roo. *El Mural*, based in the state of Jalisco, has a wide coverage of Pacific coast states, including Nayarit. The inclusion of local newspapers from *Organización Editorial Mexicana* (OEM) and *Milenio*, which reprint news from neighboring states, also allowed the author to obtain information on protest events in Chiapas and Oaxaca.

⁸ The MPC Dataset also gathered information on the type of protest, the organizers' identity, and specific claims made by participants.

more pressure,⁹ and assign protest events incorrectly, since news reports generally do not provide detailed information on different participants' places of origin. Thus, it is impossible to correctly disaggregate data at the municipal level, making the state level the ideal setting for data collection.

Our objective in this analysis is to explain the variation in the mobilization of civil society against crime as a function of received remittances across Mexican states, using quarterly data. As shown in Figure 1, there is significant variation in the spatial distribution of protests across Mexico, reinforcing the importance of explaining such state-level variance. Two northern states, Chihuahua and Nuevo León, along with the southern state of Guerrero, saw the highest number of protests during the almost six-year study period. Baja California, Sinaloa, and Veracruz followed closely and experienced between 49 and 60 protests during the same period.

*** Figure 1 about here ***

Remittances

Our main explanatory variable is *remittances* for each state–quarter in its logarithmic form, measured in constant 2003 Mexican pesos (Nyblade and O'Mahony 2014). Throughout the period, the average state level of remittances was 1.6 billion Mexican pesos. Though with some seasonal variation, the average quarterly change in national remittances was 0.44 percent (see Figure A1 in the Online Appendix). The spatial distribution of average remittances across states during our study period is shown in Figure 2. Southern and central Mexican states had the highest volume of remittances, as they are historically also the areas of higher emigration (Burgess 2005; Bada 2016)). Recall that our goal is to explore whether remittances' posited

⁹ Municipal authorities have no jurisdiction over organized crime-related activity.

contradictory effect on protests varies in a non-linear way with the size of remittance inflows. Thus, in our estimations, we include the quadratic term of remittances and investigate whether remittances increase the probability of engaging in protests against crime, but at a declining rate after a certain amount of remittance inflows.

*** Figure 2 about here ***

Controls

As in previous work (Ley et al. 2019), the control variables are mostly collected from census data, especially *Encuesta Intercensal* in 2005 and *Censo de Población y Vivienda* in 2010. Therefore, these variables are relatively constant in the resulting dataset. We control for the *homicide rate*, given that we are explaining protests against crime. The average number of homicides per 100,000 inhabitants is 4.1.

In Mexico, discontent with the security situation has motivated individuals to displace internally and internationally (Atuesta and Paredes 2016; Ríos 2014; Basu and Pearlman 2017), and security-motivated emigration likely deprives sending communities of those more critical of the government and, therefore, more likely to protest (Pfaff and Kim 2003; Kapur 2014). Emigration may, thus, reduce the incidence of protest if potential protesters are more likely to leave (Pfaff and Kim 2003; Kapur 2014). To account for this possibility, we control for *emigration rates*.¹⁰ This variable measures the percentage of households with emigrants in the five years previous to the survey collection period (INEGI).

¹⁰ The correlation between the two variables is high but lower than might be expected (0.54).

As explained in Section 1, the literature on contentious mobilization has widely considered social networks to be important determinants of participation in social protests (Diani 1995; Friedman and McAdam 1992; Fujii 2008; Gould 1993; McAdam 1982, 1988; McAdam and Paulsen 1993; Passy 2001, 2002). In Section 1, we argued that the resource effect of remittances increases participation in informal risk-sharing networks of mutual help aimed at providing insurance against risks (Gallego and Mendola 2013; Mendola 2017). To our knowledge, there are no available data on informal social networks for which we can control; however, if we find that remittances have explanatory power after controlling for a variety of official organizations and other competing explanations, this finding will be suggestive evidence that remittances shape the likelihood of protesting through changes in households' incentives to organize informally against crime.

We control for the *number of formal civil-society associations* per 100,000 inhabitants. Following Ley (2014), we use the figures reported by the Mexican Interior Ministry's Register of Civil Society Organizations, which includes non-partisan, non-religious, and non-profit groups that are *legally constituted* and that may pursue a wide range of activities, from human rights defense to welfare provision. Because of the centrality of *church associations* in building tight local networks and stimulating political involvement (Smith 1996; Fransen 2015), we control for the number of these groups. Finally, *trade unions* have played important roles in social and political movements (Chenoweth and Ulfelder 2017, 305). Thus, we include a control for the number of *labor unions* and professional associations per 100,000 inhabitants. These three variables are taken from Mexico's economic census data (Ley 2014) and are expected to correlate positively with the likelihood of protesting.

As in Ley et al. (2019, 8-9), we control for the *incumbent party* in state government. In Mexican states governed by the national incumbent party during our study period (*Partido de Acción Nacional*, PAN), voters may have been better able to assign responsibility for growing

insecurity (Ley 2017). However, criminal violence was higher in states governed by opposition parties, particularly the *Partido de la Revolución Democrática*, PRD, during the Calderón administration (Trejo and Ley 2016). We control for partisanship by including a dummy variable that equals 1 if the ruling party was from the opposition and 0 if the state was governed by the federal incumbent party (PAN).

Since we are using the level of remittances (logged), we control for the *population* (logged) of the state (INEGI). We also control for *gross domestic product* in constant 2008 pesos (logged) to distinguish the income effect of remittances from the income effect of initial wealth. Controlling for education (years of schooling) allows us to proxy the stock of human capital in the state at a given time. Additionally, we include the *percentage of indigenous population* in the state: these communities have had long traditions of strong formal and informal networks, as well as know-how for social collaboration, shared identities, solidarity, and resolution of collective action problems (Trejo 2009). Finally, we include a one-quarter lag of the dependent variable to control for possible inertia in protest activities.

3.2. Empirical Strategy¹¹

Our empirical strategy at the state level exploits the quarter-to-quarter variation in remittances by state to estimate their association with protests against crime. Our specification takes the following form:

$$Protest_{it} = \beta_0 + \alpha_i + \beta_1 \ln(remitt_{it}) + \beta_2 [\ln(remitt_{it})]^2 + \gamma_1 X_{it} + \varepsilon_{it} \quad \text{Eq (1)}$$

¹¹ We closely follow the empirical strategy used in Ley et al.'s (2019) study on remittances and vigilantism in Mexico.

for state i and quarter t . The dependent variable, $protest_{it}$, is a count variable measuring the number of protests that took place. In line with our argument, our main independent variable, $remitt_{it}$, captures remittances in the recipient state (logged) in both linear and quadratic forms. X_{it} is a vector of socioeconomic, demographic, and political determinants of protests, α_i is a state-specific effect, and ε_{it} is the error term.¹²

Since the dependent variable is a count and the protest events are not independent, the most appropriate estimation technique is a negative binomial (NB) regression. Ignoring the data's temporal variation would yield less precise estimates. Given that some explanatory variables change slowly over time and that the study period is relatively short, the use of fixed effects is less viable (Allison and Waterman 2002; Ley 2014). Thus, we test for a negative binomial regression with random effects on a panel of 31 states.

It is possible that states with a larger number of protests and greater social unrest could experience a decline in remittance flows coming from abroad. Migrants may be reluctant to send remittances in contexts of rampant crime and insecurity due to uncertainty about financial security (Meseguer et al. 2017). In other words, reverse causality between protests and remittances must be accounted for. To address this endogeneity concern, we exploit an instrumental variable approach. The instrumental variable is constructed using two distinct data sources: i) quarterly US state unemployment rates (seasonally adjusted, end of period – U.S. Bureau of Labor Statistics 2006–2014) and ii) shares of the diaspora in the top 3 US states from each Mexican state, as of 2008 – based on *matrículas emitidas* – a document issued by consulates to registered Mexicans (Institute for Mexicans Abroad, IME).

This instrumental variable $Z_{i,t}$ for Mexican state i in quarter t multiplies the seasonally adjusted quarterly unemployment rate in US destination state j at time t , $j=[1,\dots,4]$ by the share

¹² McKenzie and Rapoport (2007) also model an endogenous variable (migration) in quadratic form.

of Mexican state i 's diaspora residing in US state j , where j identifies the top three US destination states and the United States as a whole. The intuition behind this variable is that Mexican states that have larger shares of migrants in US states and that experienced strong increases in unemployment saw remittance inflows decline, as is, in fact, the case: increases in unemployment in the destination have a *negative* impact on remittances sent back home. As far as the exogeneity condition goes, it is unlikely that the unemployment rate at the migrants' US destination will have a relationship with protests against crime in the Mexican sending state other than through its effect on remittances. While there could be some concern about *matrículas* being impacted by crime-induced emigration, by using data *up to* 2008, early in our research period, we can be reasonably confident that *matrículas* were not affected by crime-induced emigration. Also, the IME 2008 data on consular *matrículas* are left without variation so that all time variation in the instrumental variable is due to fluctuations in employment at destination and not to changes in migration patterns (see Ambrosius and Meseguer 2020, 5, for a similar strategy). Our potentially endogenous variable of interest appears in the equation in linear and quadratic terms. We employ the so-called *nonlinear in endogenous variables system of equations* (Wooldridge, 2002) to estimate this specification.

4. Results

Table 2 reports the panel of negative binomial estimates of the independent variables' expected effects on the protest count. To control for more populated states being at higher risk of experiencing more protest activity, we use (log) population as the exposure variable in the model. In Column 1, the model includes remittances as a linear function only, and the coefficient is negative and significant. However, when we add the quadratic term (Column 2), the linear term turns positive and equally significant, while the quadratic term is negative and significant. We conducted a likelihood ratio (LR) test between Models 1 and 2 and are able to

reject the null hypothesis that the additional term equals zero at the 1-percent level of significance.¹³ In other words, a non-linear relationship between protests and remittances provides a better fit.

*** Table 2 about here ***

As expected, remittances have a positive effect on the expected protest count up to a certain level of remittance penetration, past which the impact decreases. Next, we calculated different specifications and introduced controls progressively in Columns 3–8. In the full specification in Column 8, increases in quarterly remittances, holding other predictors constant, are associated with increases in the difference in the log of expected counts of mobilizations against crime, with a decreasing effect after an inflection point at about 1.7 billion pesos. This figure happens to be slightly above the mean of state–quarterly remittances. Thus, remittances sent from abroad produce both an engagement and a disengagement effect when it comes to mobilizing citizens to collective action against crime. This finding shows that there is a false dichotomy when we theorize about remittances’ effect on protest. Remittances can have either effect, depending on the largesse of the inflows. Only moderate to high levels of remittance inflows are “demobilizing.” Figure 3 shows remittances’ predicted effect on the expected protest count (based on Model 2). As expected, the relationship follows an inverted U-shape.

*** Figure 3 about here ***

¹³ The LR statistic for Models 1 and 2 is 15.73. The likelihood ratio test is also conducted for the full specification (Column 8), with and without the squared term of remittances. With an LR statistic of 6.71, we are still able to reject the null hypothesis at the 1-percent level of significance.

Therefore, our results indicate that for the average Mexican state, remittances may enable protests below the estimated threshold, while above 1.7 billion pesos, they discourage them. Considering the distribution of state-remittances by quarter (Figure 4) helps clarify the substantive meaning of these estimates: a low percentage of states receive high volumes of remittances, which means that in a majority of Mexican states, remittances could have an activation effect on the probability of protesting against crime.

*** Figure 4 about here ***

Regarding control variables, we posited that emigration was more likely to deprive Mexican states of those most unsatisfied with the state of affairs, thus depressing protest. Our first finding backs this significant negative effect (Column 3), but it becomes insignificant as we include further controls. Not surprisingly, higher levels of crime proxied by crime rates increase the likelihood of protests at the state level consistently throughout all specifications.

Secondly, civil networks have explanatory power after we consider international financing of protests via remittances. This finding is very robust in the case of civil associations, confirming previous findings (Ley 2014). Indirectly, this result suggests that remittances' income effect on protest is likely to occur through the positive effect that remittances have on individual incentives to support informal social arrangements, on top of pre-existing formal networks of civil associations. Somewhat surprisingly, the number of church associations is negatively associated with protest counts, albeit with a very small effect. Note, however, that what our measure of religious density may be capturing is a proxy of religiosity, potentially associated with a more conservative reaction to crime and insecurity. It

is likely that a more accurate measure to address the role that religious networks have on protest is religious competition, whereby different churches and religious leaders organize *and* compete to serve distinct but overlapping groups in the population (Trejo 2009). However, as Trejo (2009) shows, the dynamics of religious competition occur at the diocese or municipal level, and we lack corresponding data to test such a potential mechanism.

Thirdly, having an opposition party ruling the state does not increase the expected incidence of protests against crime. Fourth, states with higher average years of education exhibit higher expected numbers of protests, and, as expected, large shares of indigenous population are also associated with more protests. State wealth, while not statistically significant, does modify the size of the estimated effect of remittances on protests, suggesting that we need to control for subnational income levels. Finally, as Column 8 shows, the number of previous protests is positively, but not significantly, associated with contemporaneous protests after the rest of covariates are controlled for.

*Robustness*¹⁴

Table 3 contains the instrumental variable approach results that provide evidence on the robustness of the estimated relationship. To address possible endogeneity issues stemming from our potentially endogenous variable of interest in both linear and quadratic terms, we adopt an instrumental variable approach (Wooldridge 2002) that is similar to a three-stage least squares estimation. The instrumental variable estimation is restricted to an OLS with random effects because the instrumental variable procedure implemented only holds when the system

¹⁴ The procedure to deal with endogeneity closely follows the procedure in our study on remittances and vigilantism in Mexico (Ley et al. 2019, 13).

of equations is linear in the parameters (Wooldridge 2002). The estimates for the three-stepwise instrumental variable approach are reported in Table 3.

*** Table 3 about here ***

In the first stage, Column 1, we regress the exogenous instrument of diaspora-weighted unemployment in the United States and the other exogenous covariates on remittances.¹⁵ The instrument is significantly associated with remittances: in line with theoretical expectations, rates of unemployment in migrants' destinations are negatively associated with remittances received. The linear prediction (fitted value) of remittances and its squared term from the previous stage are used as the excluded instruments in a 'two-stage least squares' estimation with two endogenous variables. We then have two additional first-stage regressions, one for each endogenous variable, and two instruments.¹⁶ Columns 2 and 3 in Table 3 have been labeled as 'second stage.' Note that these columns include the first-stage regressions of the two endogenous variables – namely, remittances and remittances squared. The linear predictions obtained in Column (1) are the excluded instruments ($\text{Pr}[\text{Remittances}]$ and $\text{Pr}[\text{Remittances}^2]$). We confirm their relevance separately in the equations for each endogenous regressor and jointly for the last stage.

The procedure's third step is shown in the last column (4) of Table 3, which gives the estimates of the second-stage regression (labeled as third stage) for the linear and quadratic effects of remittances on protests. The inverted U-shaped relationship between remittances and

¹⁵ The Wald chi squared statistic indicates that we can reject the null hypothesis that the coefficients of the joint instruments in the reduced form equal zero.

¹⁶ The equation is exactly identified; thus, we are unable to test for exogeneity.

the protest count is robust to instrumenting the endogenous terms of remittances and remittances squared. The interpretation of these estimates is straightforward: a 10-percent increase in remittances is associated with an increase of 0.83 in the expected number of protests, holding everything else constant. We illustrate the inverted U-shape by calculating the predictive margins for specific values of the remittance distribution, keeping the rest of the covariates constant (Table 4).

*** Table 4 about here ***

The predicted numbers of protests for different levels of remittances show a positive but decreasing effect. The expected number of protests increases from 0.448 protests at the 5th percentile to 1.073 at the median value of remittances. After this level of quarterly state remittances, the predicted number of protests decreases to 0.919 (for the 75th percentile) and drops further to 0.646 for states in the 95th percentile of remittances. Thus, remittances increase the probability of protesting, but at a declining rate, which provides robust evidence that remittances sent from abroad produce both an engagement and a disengagement effect on protest against crime.¹⁷

Finally, we carried out a number of other robustness tests and controlled for multiple alternative explanations. First, the absolute volume of remittances may be hiding a state-size effect. To rule out this possibility, we estimated our main specification, using two alternative measures that capture the effect of remittance density on protests – namely, *remittances per capita* and *the percentage of households receiving remittances at the state level* (Online

¹⁷ Another approach to instrument selection is to include higher-order terms of exogenous variables appearing in the system. We used this approach, and the main finding holds (Online Appendix Table A6).

Appendix Table A2).¹⁸ Our result holds. We also included a *quarterly indicator of economic activity* (INEGI), instead of the yearly GDP, to match the periodicity of remittances (see Column 5 in Online Appendix Table A4). While the level of economic activity positively affects the likelihood of protesting, the inclusion of this control does not alter our main finding.

Second, we acknowledge that other types of crime, besides homicide rates, may influence protests. However, adding several crime variables in the same specification would only introduce multicollinearity, due to a high correlation among crime variables. Nonetheless, we provide an analysis of sensitivity to different crime measures, including the rate of disappearances and the rate of drug-related homicides separately (Online Appendix Table A3). All have the expected positive effect on protest, without affecting our main finding. Third, we controlled for *inequality*, which other authors have demonstrated to be an important determinant of violent protests in the form of *vigilante* organizations (Phillips 2017). Inequality is marginally relevant to explaining the incidence of protest (significant at 10 percent), but not consistent, and it does not alter our main finding (Online Appendix Table A4).

Fourth, remittances could potentially increase state capacity if, for instance, recipients demand more accountability and help reduce corruption (Burgess 2005; Tyburski 2014). This increase in state capacity, in turn, may reduce the motivation to protest. To address this possibility, we controlled for the *number of libraries* at the state level, following Phillips (2017), who notes that while the number of libraries is not a measure of state capacity in terms of security or the judicial apparatus, it is a good *proxy* for overall government capacity. Additionally, libraries are exogenous to security issues, which allows for a more accurate assessment. This variable turned out to be statistically insignificant and its inclusion does not affect our main finding (Online Appendix Table A4). Fifth, the very existence of self-defense

¹⁸ We prefer our measure to one of normalizing by state wealth, which is not exogenous to remittances.

organizations may correlate with the type of protests we explore here or even make protests against crime unnecessary to some extent. The presence of *vigilante* organizations does increase the probability of protest against crime, but including this variable does not alter our main result (Online Appendix Table A5).

Finally, while we controlled for various types of local formal networks, we also controlled for transnational migrant organizations, known as *Hometown Associations* (HTAs, Institute for Mexicans Abroad, IME). While the resources these organizations raise and invest in their communities (also known as collective remittances) are very small in comparison to the volume of individual remittances, these organizations bring organizational skills and experience to their communities (Burgess 2005; Duquette-Rury 2014; Bada 2016; Pérez-Armendáriz and Duquette-Rury 2019) that could well facilitate the organization of protests. HTAs do not seem to predict the protest count, and our main finding is robust to controlling for the number of existing HTAs in a given state (Online Appendix Table A5).¹⁹

5. Discussion

Workers' remittances help those left behind organize and protest against crime. Because they also improve recipients' living conditions, reduce economic risks, and improve perceptions of the security situation, however, remittances finance protests against crime at a declining rate. In other words, remittances provide resources to protest, but they also reduce the reasons to do so. This finding is relevant to several literatures. First, research on the international determinants of protests is only starting to pay attention to emigration and financial remittances (Barry et al. 2014; Miller and Ritter 2014; Escribà-Folch et al., 2018;

¹⁹ Since information on these variables exist for only one year, the models including *Vigilantes* and HTAs are cross-sections of 31 states. As such, this evidence should be taken as preliminary.

Germano 2018). In the literature on protests, resources are deemed essential for grievances to be transformed into collective action. Remittances provide resources to the senders' families and communities, which can then engage in collective action to protest against crime and insecurity. As such, this source of international flows should be systematically taken into consideration in future research on the international determinants of violent and non-violent protest.

Second, we contribute to the thriving literature on the political consequences of remittances. This literature has so far treated remittances as either causing political engagement or causing disengagement (Goodman and Hiskey 2008; Bravo 2009; Germano 2013, 2018; Pfutze 2014). We have added an important nuance to this finding by showing that remittances can cause *both*. As we demonstrated, remittances' impact on protests is not linear, and their positive effect on protest against crime declines in settings where remittances are slightly above average. Relying on recent scholarship (Adida and Girod 2011; Doyle 2015; Pfutze 2014; Fransen 2015), we have argued that this finding has to do with the income and substitution effects of remittances, which grant recipients more autonomy from the state, reduce grievances, and facilitate the formation of informal networks of mutual help. Informal networks and formal organizations have been behind the protest movement against crime studied in this article (Ley 2014; Durán-Martínez 2016; Dorff 2017). Interestingly, we find a similar non-linear effect of remittances on the probability of financing *vigilante* organizations at the municipal level in Mexico (Ley et al. 2019); but in the case of vigilantism, the tipping point after which remittances impacted the formation of self-defense organizations at a declining rate was located at a higher level of penetration at the municipal level. In other words, it takes large inflows of remittances to slow the formation of grass-root self-defense organizations. We interpret this finding as indicative that in comparison to non-violent protests, vigilantism is a more resource-intensive form of collective mobilization that requires greater and sustained financial support

(Phillips 2017; Moncada 2017).

Future research should systematically explore remittances' impact on formal and informal civilian engagement. While some research on Mexico has investigated remittances' role in facilitating some forms of social networks (Duquete-Rury and Chen 2018), scholars have not yet made a consistent distinction between formal and informal networks. However, the fact that remittances retain explanatory power at the state level after controlling for a number of formal civil associations suggests that these inflows have been relevant in the spontaneous organization of mutual-help, risk-sharing networks, which have played a sizeable role in the protests we study here (Ley 2014).

Second, researchers should explore whether other types of migrant connectivity in the form of social remittances (Levitt 1998) also help those left behind mobilize to demand better protection and access to justice and to protest against crime and impunity. Research has shown that migrants often get familiar with, if not directly participate in, collective mobilization in their destinations (Pérez-Armendáriz and Crow 2010; Nikolova et al. 2017; Petrova 2019). This experience helps them become acquainted with different repertoires of collective action that can then be shared with relatives through transborder conversations with relatives back home (Pérez-Armendáriz and Crow 2010; Nikolova et al. 2017; Petrova 2019). Likely, the international transfer of norms also plays a role in the phenomenon we explore in this piece.

Finally, our findings should be tested in other instances of violent democracies to gain further external validity. Unfortunately, contexts of high crime and high remittance dependence in which to test our findings are abundant (Pérez-Armendáriz 2019), and these tests should figure at the top of this research agenda. All in all, our research calls attention to remittances as determinants of protest against crime and helps advance understandings of how remittances impact this particular example of non-electoral political activity. International networks of migrant solidarity with those left behind matter, as do the financial resources migrants send

back home.

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FIGURES

Figure 1. Protest Against Crime, Count by State, 2006/1–2011/2

Source: Authors, with data from Mexican Protest Against Crime (MPC) Dataset (Ley, 2014).

Figure 2. Average Remittances by State, 2006/1–2011/2

Figure 3. Predicted non-linear effect of remittances on expected count of protests

Figure 4. Distribution of quarterly remittances

TABLES

Table 1. Protests against Crime and Remittances: Summary Statistics

VARIABLES	(1) N	(2) mean	(3) SD	(4) Min	(5) Max
<i>Protests</i>					
Number of protests ^q	682	1.078	2.604	0	29
<i>Remittances</i>					
Remittances ^q	682	1,618	1,486	61.97	6,277
Log Remittances ^q	682	6.923	1.051	4.127	8.745
<i>Controls</i>					
Emigration ^{5y}	682	3.834	2.908	0.446	12.18
Log Population ^{5y}	682	14.68	0.744	13.15	16.54
Average years of schooling ^{5y}	682	6.900	0.962	4.992	9.020
Share of indigenous population ^{5y}	682	0.104	0.136	0.002	0.542
Opposition party in state government ^q	682	0.756	0.428	0	1
Homicide rate ^q	682	4.148	5.912	0.0466	58.54
Unions ^y	682	112.5	82.27	33	530.7
Civil associations ^y	682	224.6	161.6	40	1,025
Churches ^y	682	1,736	1,171	272.0	5,873
Log GDP per capita ^y	682	12.40	0.732	11.05	13.91
<i>Instrument</i>					
Weighted US unemployment rates ^q	682	7.420	2.512	4.082	11.72

Notes: The superscripts on each variable indicate the level of variation: quarterly^q, yearly^y, or 5-years^{5y}

Table 2. Protest and Remittances: Negative Binomial Random Effects by Quarter.

Dep. Var. Protests	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Log(remittances)	-0.3205*** (0.119)	4.5368*** (1.285)	4.5247*** (1.366)	4.2227*** (1.294)	3.4262*** (1.323)	3.4735*** (1.333)	4.3534*** (1.533)	4.2970*** (1.518)
Log(remittances) ²		-0.3493*** (0.092)	-0.3344*** (0.098)	-0.3198*** (0.092)	-0.2444** (0.096)	-0.2462** (0.097)	-0.2881*** (0.110)	-0.2880*** (0.108)
Emigration			-0.0991*** (0.037)	-0.0393 (0.036)	-0.0899** (0.041)	-0.0926** (0.041)	-0.0189 (0.052)	-0.0147 (0.051)
Homicide rate				0.0450*** (0.006)	0.0408*** (0.007)	0.0405*** (0.007)	0.0413*** (0.007)	0.0397*** (0.008)
Civil associations					0.0026*** (0.001)	0.0027*** (0.001)	0.0022*** (0.001)	0.0018** (0.001)
Unions					0.0004 (0.002)	0.0003 (0.002)	0.0012 (0.002)	0.0014 (0.002)
Churches					-0.0006*** (0.000)	-0.0007*** (0.000)	-0.0007*** (0.000)	-0.0007*** (0.000)
Opposition party						0.1614 (0.190)	0.3876* (0.219)	0.2837 (0.218)
Average schooling							0.6697** (0.266)	0.5940** (0.256)
Indigenous population							0.0320* (0.017)	0.0295* (0.016)
Log (GDP)							0.1332 (0.255)	0.1252 (0.250)
Lagged protest								0.0120 (0.014)
Observations	682	682	682	682	682	682	682	651
Number of states	31	31	31	31	31	31	31	31
Log likelihood	-851.3	-843.4	-839.7	-823.4	-813.7	-813.3	-809.6	-787.3

Notes: Robust standard errors (observed information matrix, OIM) in parentheses. *** p<0.01, ** p<0.05, * p<0.1

Table 3. Protest and Remittances: Instrumental Variables

Dep. Var. Protests	<i>first stage</i> (1)	second stage Remittances (2)	second stage Remittances ² (3)	third stage Protest (4)
<i>Log(remittances)</i>				9.0438** (4.547)
<i>Log(remittances)²</i>				-0.2802** (0.138)
Emigration	-0.0190*** (0.007)	0.1042*** (0.007)	1.4502*** (0.101)	-0.5097* (0.298)
Homicide rate	0.0031* (0.002)	-0.0000 (0.003)	-0.0174 (0.041)	0.2248*** (0.060)
Civil associations	0.0001 (0.000)	0.0007*** (0.000)	0.0099*** (0.002)	-0.0030 (0.002)
Unions	-0.0013 (0.001)	0.0015*** (0.000)	0.0213*** (0.004)	-0.0039 (0.003)
Churches	-0.0000 (0.000)	-0.0002*** (0.000)	-0.0023*** (0.000)	0.0009 (0.001)
Opposition party	-0.0729** (0.028)	-0.0607* (0.031)	-0.9063** (0.429)	0.3518 (0.341)
Average schooling	-0.3181*** (0.042)	-0.1244** (0.054)	-1.8428*** (0.700)	1.3951* (0.805)
Indigenous population	-0.0222*** (0.008)	-0.0020 (0.004)	-0.0310 (0.051)	0.0679* (0.039)
Log (population)	1.3244*** (0.157)	0.7236*** (0.230)	10.9862*** (2.877)	-7.1600* (4.198)
Log (GDP)	-0.3449 (0.222)	-0.3258*** (0.073)	-4.8733*** (0.912)	2.7278* (1.536)
Lagged protest	-0.0020 (0.002)	-0.0120** (0.006)	-0.1456* (0.080)	0.2824** (0.110)
<i>IV Unemployment</i>	-0.0416*** (0.006)			
Pr[<i>Log(remittances)</i>]		-0.1026 (0.246)	-17.0580*** (2.893)	
Pr[<i>Log(remittances)²</i>]		0.0415*** (0.015)	1.6569*** (0.161)	
R-squared	0.798	0.917	0.921	0.050
Wald Chi2	700.3	-	-	-
F	-	7.38	55.15	-
Sanderson-Windmeijer F	-	9.70	17.24	-
Sanderson-Windmeijer Chi2	-	9.92	17.62	-
Kleibergen-Paap rk LM	-	-	-	8.429
Kleibergen-Paap rk Wald F	-	-	-	5.360
Observations	651	651	651	651

Robust standard errors in parentheses *** p<0.01, ** p<0.05, * p<0. The F tests of excluded instruments are reasonably high, reassuring us of the relevance of our instruments. The Sanderson-Windmeijer F statistic confirms that the parameters are not weakly identified. The Sanderson-Windmeijer chi-squared Wald statistics rejects the null hypothesis that the endogenous parameter is under-identified. We also test the null hypothesis that our two excluded instruments are redundant, and we reject the null hypothesis at the 1 percent level of significance. Kleibergen-Paap LM statistic rejects the null that the model is under-identified. Kleibergen-Paap rk Wald F tells us that the equation is not weakly identified.

Table 4. Predicted Number of Protests	
Percentile of remittances	Fitted value of protests
5 th	0.448
25 th	1.031
Median	1.073
75 th	0.919
95 th	0.646

Online Appendix

Figure A1. Quarterly variation of remittances to Mexico

Source: Authors' own elaboration on data from Nyblade and O'Mahony (2014)

Table A1. List of newspapers used for the MSD dataset

State	Newspaper	Years
Aguascalientes	El Sol de Aguascalientes	2008-2012
Baja California	Frontera	2006-2008
	Crónica	2009
	El Vigía	2010-2012
Baja California Sur	El Sudcaliforniano	2006-2012
Coahuila	El Siglo de Torreón	2006-2012
	El Zócalo de Saltillo	2006-2012
Colima	Diario de Colima	2006-2012
Chihuahua	The Chihuahua News Database, provided by Información Procesada(INPRO)	2006-2012
Durango	El Siglo de Durango	2006-2012
	El Sol de Durango	2008-2012
Guanajuato	Periódico AM	2006-2012
	Milenio León	2006-2012
Guerrero	El Sur	2006-2012
	El Sol de Acapulco	2008-2012
Hidalgo	El Sol de Hidalgo	2006-2012
	Milenio Pachuca	2006-2012
Jalisco	El Mural	2006-2012
	El Informador	2006-2012
México	Milenio Estado de México	2006-2012
Michoacán	El Sol de Morelia	2008-2012
	Cambio	2009-2012
Morelos	El Sol de Cuernavaca	2008-2012
	La Unión	2007-2012
Nuevo León	El Norte	2006-2012
	El Porvenir	2006-2012
Puebla	El Sol de Puebla	2006-2012
	Milenio Puebla	2006-2012
Querétaro	Diario de Querétaro	2006-2012
San Luis Potosí	El Sol de San Luis	2006-2012
	La Jornada de San Luis	2006-2012
Sinaloa	El Sol de Sinaloa	2008-2012
	Noroeste	2008-2012
Sonora	El Imparcial	2006-2012
Tabasco	Milenio Villahermosa	2006-2012
Tamaulipas	El Sol de Tamaulipas	2008-2012
	El Mañana	2009-2012
	Milenio Tampico	2006-2012
Tlaxcala	El Sol de Tlaxcala	2008-2012
Veracruz	El Sol de Orizaba	2008-2012
	El Sol de Córdoba	2008-2012
	Milenio Xalapa	2006-2012
	Liberal	2008-2012
	La Jornada Veracruz	2011-2012
Yucatán	Diario de Yucatán	2006-2012
Zacatecas	El Sol de Zacatecas	2008-2012
	Imagen	2006-2007
	NTR	2008-2012
National newspaper	Reforma	2006-2012

Table A2. Protest and Remittances: Negative Binomial with density of remittance penetration measured in per capita terms and as the % of households receiving remittances at the state level.

Dep. Var. Protests	<i>Remittances per capita</i> (1)	<i>% of households with remittances</i> (2)
<i>Density of remittances</i>	0.0036** (0.001)	0.3169*** (0.091)
<i>Density of remittances</i> ²	-0.000002* (0.000)	-0.0107** (0.005)
Emigration	-0.0629 (0.059)	-0.0559 (0.051)
Homicide rate	0.0399*** (0.008)	0.0364*** (0.008)
Civil associations	0.0024*** (0.001)	0.0011 (0.001)
Unions	0.0002 (0.002)	0.0006 (0.002)
Churches	-0.0006** (0.000)	-0.0000 (0.000)
Opposition party	0.3195 (0.230)	0.1290 (0.210)
Average schooling	0.6290** (0.259)	0.4818** (0.223)
Log (GDP)	0.4309 (0.330)	0.5925** (0.238)
Indigenous population	0.0249 (0.017)	0.0258* (0.015)
Lagged protest	0.0122 (0.015)	0.0095 (0.014)
Observations	651	651
Number of states	31	31
Log likelihood	-788.7	-787.9

Notes: Robust standard errors (OIM) in parentheses. *** p<0.01, ** p<0.05, * p<0.1. The variable percentage of households receiving remittances is taken from *Censos de Población y Vivienda*, INEGI.

Table A3. Protest and Remittances: Negative Binomial sensitivity to different crime measures

Dep. Var. Protests	(1) <i>Homicides</i>	(2) <i>Drug-related Homicides</i>	(3) <i>Disappearances</i>	(4) <i>Num. Cartels</i>
Log(remittances)	4.2970*** (1.518)	4.2969*** (1.552)	4.2549*** (1.583)	4.4165*** (1.510)
Log(remittances) ²	-0.2880*** (0.108)	-0.2857*** (0.111)	-0.2748** (0.113)	-0.3033*** (0.109)
Emigration	-0.0147 (0.051)	-0.0146 (0.051)	-0.0281 (0.052)	-0.0053 (0.053)
<i>Crime variable</i>	<i>0.0397*** (0.008)</i>	<i>0.0099*** (0.003)</i>	<i>0.0607*** (0.015)</i>	<i>0.0076** (0.004)</i>
Civil associations	0.0018** (0.001)	0.0024*** (0.001)	0.0027*** (0.001)	0.0013 (0.001)
Unions	0.0014 (0.002)	0.0011 (0.002)	0.0012 (0.002)	0.0008 (0.002)
Churches	-0.0007*** (0.000)	-0.0008*** (0.000)	-0.0009*** (0.000)	-0.0006*** (0.000)
Opposition party	0.2837 (0.218)	0.3496 (0.222)	0.3423 (0.224)	0.1474 (0.221)
Average schooling	0.5940** (0.256)	0.5563** (0.252)	0.6374** (0.260)	0.5553** (0.245)
Log (GDP)	0.1252 (0.250)	0.0933 (0.257)	0.0746 (0.264)	-0.0192 (0.244)
Indigenous population	0.0295* (0.016)	0.0314* (0.016)	0.0366** (0.017)	0.0282* (0.015)
Lagged protest	0.0120 (0.014)	0.0216 (0.016)	0.0423*** (0.014)	0.0410*** (0.015)
Observations	651	651	651	651
Number of states	31	31	31	31
Log likelihood	-787.3	-790.1	-788.5	-793.2

Notes: Robust standard errors (OIM) in parentheses.*** p<0.01, ** p<0.05, * p<0.1. Additional variables come from different sources: Drug-related homicides (*Presidencia de la República Mexicana, 2011*), Disappearances (*Registro Nacional de Datos de Personas Extraviadas o Desaparecidas*), Number of cartels (Coscia and Rios, 2012).

Table A4. Protest and Remittances: NB Additional controls; GINI, Libraries, Economic Activity

VARIABLES	(1)	(2)	(3)	(4)	(5)
Log(remittances)	4.2970*** (1.518)	4.0287*** (1.479)	4.3447*** (1.544)	4.0767*** (1.502)	4.2219*** (1.443)
Log(remittances) ²	-0.2880*** (0.108)	-0.2674** (0.106)	-0.2849*** (0.110)	-0.2660** (0.107)	-0.2870*** (0.103)
Emigration	-0.0147 (0.051)	-0.0132 (0.051)	-0.0165 (0.051)	-0.0152 (0.051)	0.0025 (0.048)
Homicide rate	0.0397*** (0.008)	0.0393*** (0.008)	0.0413*** (0.008)	0.0407*** (0.008)	0.0408*** (0.008)
Civil associations	0.0018** (0.001)	0.0018** (0.001)	0.0020** (0.001)	0.0019** (0.001)	0.0018** (0.001)
Unions	0.0014 (0.002)	0.0005 (0.002)	0.0017 (0.002)	0.0009 (0.002)	0.0013 (0.002)
Churches	-0.0007*** (0.000)	-0.0006*** (0.000)	-0.0007*** (0.000)	-0.0007*** (0.000)	-0.0006*** (0.000)
Opposition party	0.2837 (0.218)	0.2597 (0.218)	0.2237 (0.224)	0.2123 (0.223)	0.1790 (0.213)
Average schooling	0.5940** (0.256)	0.6395** (0.261)	0.6302** (0.256)	0.6570** (0.258)	0.5121** (0.244)
Log (GDP)	0.1252 (0.250)	0.0871 (0.253)	0.1238 (0.253)	0.0894 (0.254)	
Indigenous population	0.0295* (0.016)	0.0259 (0.016)	0.0300* (0.016)	0.0265 (0.016)	0.0218 (0.014)
Lagged protest	0.0120 (0.014)	0.0096 (0.015)	0.0135 (0.015)	0.0113 (0.015)	0.0134 (0.014)
GINI		0.0618* (0.037)		0.0535 (0.037)	
Libraries per capita			2.8212 (1.949)	2.3149 (1.992)	
Economic Activity					0.0419*** (0.013)
Observations	651	651	651	651	651
Number of states	31	31	31	31	31
Log likelihood	-787.3	-785.9	-786.2	-785.2	-782.2

Notes: Robust standard errors (OIM) in parentheses. *** p<0.01, ** p<0.05, * p<0. Additional variables used come from different sources: GINI (Phillips, 2017), libraries per capita (*Sistema Estatal y Municipal de Bases de Datos*, INEGI), indicator of economic activity (*Banco de Información Económica*, INEGI)

Table A5. Protest and Remittances: Cross-sectional evidence on Vigilantes and HTAs

VARIABLES	(1)	(2)	(3)
Log(remittances)	2.8746** (1.275)	2.6358** (1.196)	3.4979*** (1.315)
Log(remittances) ²	-0.2511*** (0.095)	-0.2387*** (0.089)	-0.3115*** (0.100)
Emigration	0.2246** (0.091)	0.2680*** (0.092)	0.2668*** (0.088)
Homicide rate	0.1377*** (0.020)	0.1300*** (0.020)	0.1294*** (0.019)
Civil associations	-0.0026** (0.001)	-0.0023* (0.001)	-0.0021* (0.001)
Unions	0.0040** (0.002)	0.0034** (0.002)	0.0040** (0.002)
Churches	0.0002 (0.000)	0.0001 (0.000)	0.0001 (0.000)
Opposition party	-0.1823 (0.229)	-0.2699 (0.220)	-0.1847 (0.218)
Average schooling	0.3089 (0.218)	0.2946 (0.212)	0.3455* (0.206)
Log (GDP)	0.0285 (0.231)	0.1470 (0.222)	0.0791 (0.222)
Indigenous population	0.0110 (0.013)	0.0131 (0.013)	0.0146 (0.013)
Hometown migrant associations	0.0016 (0.001)		0.0023 (0.001)
Vigilante group		0.2879 (0.217)	0.4035* (0.221)
Observations	31	31	31
Log likelihood	-103.6	-103.4	-102.1

Notes: Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Additional variables come from different sources: Hometown migrant associations (*Instituto de los Mexicanos en el Exterior*, 2015) and vigilante groups (Phillips, 2017)

Table A6. Protest and Remittances: Alternative Instrumental Variables

Dep. Var. Protests	First stage Remittances (1)	First stage Remittances ² (2)	Second stage Protest (3)
<i>Log(remittances)</i>			15.0200** (7.291)
<i>Log(remittances)²</i>			-0.9011** (0.429)
Emigration	0.0989*** (0.007)	1.4476*** (0.099)	-0.1944 (0.162)
Homicide rate	0.0007 (0.001)	-0.0027 (0.012)	0.0415*** (0.014)
Civil associations	0.0008*** (0.000)	0.0109*** (0.002)	-0.0001 (0.001)
Unions	0.0008*** (0.000)	0.0118*** (0.003)	-0.0021 (0.003)
Churches	-0.0002*** (0.000)	-0.0026*** (0.000)	0.0004 (0.000)
Opposition party	-0.0988*** (0.033)	-0.9868** (0.464)	0.5566 (0.380)
Average schooling	-0.2701*** (0.021)	-3.4576*** (0.302)	0.9092 (0.574)
Indigenous population	-0.0091*** (0.002)	-0.0890*** (0.024)	0.0610* (0.034)
Log (population)	1.3448*** (0.051)	17.8102*** (0.670)	-4.3802 (2.832)
Log (GDP)	-38.4995*** (12.933)	-527.2928*** (172.166)	2.1861* (1.246)
Lagged protest	-0.0139*** (0.005)	-0.1904** (0.074)	0.3528*** (0.105)
<i>IV Unemployment</i>	-0.0178*** (0.007)	-0.1777** (0.088)	
<i>IV Log (GDP)²</i>	2.9787*** (1.041)	40.4865*** (13.883)	
<i>IV Log (GDP)³</i>	-0.0775*** (0.028)	-1.0421*** (0.372)	
R-squared	0.918	0.914	0.080
F	10.11	14.43	-
Sanderson-Windmeijer F	9.37	12.22	-
Sanderson-Windmeijer Chi2	19.18	25.02	-
Kleibergen-Paap rk LM	-	-	12.482
Kleibergen-Paap rk Wald F	-	-	4.367
Observations	651	651	651

Robust standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1. In this specification we use the exogenous unemployment IV, plus the quadratic and cubic terms of GDP per capita. These two additional instruments are in fact correlated with remittances, and they can arguably be excluded from the main equation. The end result is a system of equations with three excluded instruments. The intuition is that nonlinear functions of the endogenous variable have a linear projection that depends on new functions of the exogenous variables.